

CT/MRI and CEUS LI-RADS Major Features Association with Hepatocellular Carcinoma: Individual Patient Data Meta-Analysis

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Conflicts of interest are listed at the end of this article.

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Background: The Liver Imaging Reporting and Data System (LI-RADS) assigns a risk category for hepatocellular carcinoma (HCC) to imaging observations. Establishing the contributions of major features can inform the diagnostic algorithm.

Purpose: To perform a systematic review and individual patient data meta-analysis to establish the probability of HCC for each LI-RADS major feature using CT/MRI and contrast-enhanced US (CEUS) LI-RADS in patients at high risk for HCC.

Materials and Methods: Multiple databases (MEDLINE, Embase, Cochrane Central Register of Controlled Trials, and Scopus) were searched for studies from January 2014 to September 2019 that evaluated the accuracy of CT, MRI, and CEUS for HCC detection using LI-RADS (CT/MRI LI-RADS, versions 2014, 2017, and 2018; CEUS LI-RADS, versions 2016 and 2017). Data were centralized. Clustering was addressed at the study and patient levels using mixed models. Adjusted odds ratios (ORs) with 95% CIs were determined for each major feature using multivariable stepwise logistic regression. Risk of bias was assessed using Quality Assessment of Diagnostic Accuracy Studies 2 (QUADAS-2) (PROSPERO protocol: CRD42020164486).

Results: A total of 32 studies were included, with 1170 CT observations, 3341 MRI observations, and 853 CEUS observations. At multivariable analysis of CT/MRI LI-RADS, all major features were associated with HCC, except threshold growth (OR, 1.6; 95% CI: 0.7, 3.6; *P* = .07). Nonperipheral washout (OR, 13.2; 95% CI: 9.0, 19.2; *P* = .01) and nonrim arterial phase hyperenhancement (APHE) (OR, 10.3; 95% CI: 6.7, 15.6; *P* = .01) had stronger associations with HCC than enhancing capsule (OR, 2.4; 95% CI: 1.7, 3.5; *P* = .03). On CEUS images, APHE (OR, 7.3; 95% CI: 4.6, 11.5; *P* = .01), late and mild washout (OR, 4.1; 95% CI: 2.6, 6.6; *P* = .01), and size of at least 20 mm (OR, 1.6; 95% CI: 1.04, 2.5; *P* = .04) were associated with HCC. Twenty-five studies (78%) had high risk of bias due to reporting ambiguity or study design flaws.

Conclusion: Most Liver Imaging Reporting and Data System major features had different independent associations with hepatocellular carcinoma; for CT/MRI, arterial phase hyperenhancement and washout had the strongest associations, whereas threshold growth had no association.

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Abbreviations

APHE = arterial phase hyperenhancement, CEUS = contrast-enhanced US, HCC = hepatocellular carcinoma, IPD = individual patient data, LI-RADS = Liver Imaging Reporting and Data System, OR = odds ratio, QUADAS = Quality Assessment of Diagnostic Accuracy Studies

Summary

Most CT/MRI and CEUS LI-RADS major features had independent associations with hepatocellular carcinoma; arterial phase hyperenhancement and washout had the strongest associations whereas, for CT/MRI, threshold growth had no association.

Key Results

- \blacksquare In this meta-analysis of 32 studies with 1170 CT observations, 3341 MRI observations, and 853 contrast-enhanced US (CEUS) observations, all CT/MRI Liver Imaging Reporting and Data Systems (LI-RADS) major features except threshold growth (odds ratio [OR], 1.6; *P* = .07) were independently associated with hepatocellular carcinoma (HCC).
- On CEUS images, arterial phase enhancement (OR, 7.3; $P = .01$), late and mild washout (OR, 4.1; *P* = .01), and size of at least 20 mm (OR, 1.6; $P = .04$) were associated with HCC.

The Liver Imaging Reporting and Data System (LI-RADS) is used to assign a risk category of hepatocellular carcinoma is used to assign a risk category of hepatocellular carcinoma (HCC) to liver observations at imaging in patients at high risk for HCC. LI-RADS algorithms have been developed for screening (US), diagnosis (CT/MRI and contrast-enhanced US [CEUS]), and after local-regional treatment assessment. The LI-RADS framework aims to standardize reporting and data collection of imaging for HCC to enhance communication, reduce interobserver variability, and facilitate quality assurance and research (1). LI-RADS is regularly updated to achieve these aims as new evidence emerges.

The CT/MRI diagnostic algorithm uses a combination of major features (size, nonrim arterial phase hyperenhancement [APHE], nonperipheral washout, enhancing capsule, and threshold growth) to assign categories. Similarly, CEUS uses a combination of major features (size, nonrim APHE, and late and mild washout) to assign categories. In both diagnostic algorithms, each category reflects a relative probability of benignity, malignancy in general, or HCC. Recent systematic reviews found that the percentage of HCC (equivalent to positive predictive value) differed for each CT/MRI LI-RADS category and increased as LI-RADS category increased (2–4). Limitations of these reviews include the lack of individual patient data (IPD) to determine the independent impact of each of the major imaging features on the final diagnosis of HCC. Therefore, it remains unclear if some major features increase the likelihood of HCC more than others, whether features should be assigned different weights, or if some major features are unnecessary and could be eliminated. Understanding the relative contributions of imaging features to system performance is important for continued development and improvement of LI-RADS, which could be achieved using IPD meta-analysis (5).

Given the many LI-RADS imaging features, it is challenging for a single study to achieve sufficient statistical power to analyze the impact of each imaging feature. An IPD meta-analysis would improve understanding of which imaging features drive

LI-RADS performance. The IPD meta-analyses use large and detailed data sets at the patient level to perform more complex subgroup analysis than can be achieved in any single study or by using meta-analysis of aggregate study-level data (6). These involve collecting and pooling de-identified primary patient data from authors of prior publications (7). The purpose of this systematic review and IPD meta-analysis was to establish the likelihood of HCC for each LI-RADS major feature using CT/MRI LI-RADS and CEUS LI-RADS in patients at high risk for HCC.

Materials and Methods

The study protocol was approved by the Ottawa Hospital Research Ethics Board, is Health Insurance Portability and Accountability Act compliant, and was registered on PROSPERO (CRD42020164486). Methodologic guidance was per best practice in diagnostic test accuracy systematic reviews (8,9). Reporting is in accordance with the Preferred Reporting Items for a Systematic Review and Meta-analysis of Diagnostic Test Accuracy Studies and Individual Patient Data (10–14).

Eligibility Criteria

All CT, MRI, and CEUS studies reporting the percentage of HCC and overall malignancy for LI-RADS categories 1–5, tumor in vein, and malignancy in patients at high risk of HCC (hepatic cirrhosis, chronic hepatitis B viral infection, current or prior HCC) were eligible for inclusion. The CT, MRI, and CEUS techniques were evaluated for each study to determine concordance with the LI-RADS technical imaging guidelines (15). All liver observations were required to have been categorized using CT/MRI LI-RADS version 2014, 2017, or 2018 or CEUS LI-RADS version 2016 or 2017 (16–20). A preferred reference standard was established to assess bias risk (Appendix E1 [online]).

Database Search and Study Selection

With the assistance of an experienced hospital librarian, we performed a search of the MEDLINE, Embase, Cochrane Central Register of Controlled Trials, and Scopus databases for studies from January 2014 to September 2019 that evaluated the diagnostic accuracy of CT, MRI, or CEUS for HCC using LI-RADS (Appendix E2 [online]). The corresponding authors of each study identified for inclusion were contacted (Appendix E3 [online]).

Data Collection Process and Definitions for Data Extraction

Authors who did not respond to the invitation to collaborate were sent follow-up emails in an effort to maximize data set size. All authors agreeing to participate were sent a formal confidentiality agreement explaining that data would be stored securely and only accessed by authorized coinvestigators with a copy of the data contribution form, data extraction sheet, data dictionary, and a list of frequently asked questions (Appendixes E4–E7 [online]). The request for de-identified data included instructions to transfer data to an encrypted directory. On the basis of institutional policies, when necessary, data sharing agreements were obtained. Efforts were made to keep all collaborators involved and informed of progress. IPD were not distributed elsewhere.

Risk of Bias and Applicability

A previously customized Quality Assessment of Diagnostic Accuracy Studies 2 (QUADAS-2) tool for application to LI-RADS in a prior study-level systematic review was modified to assess risk of bias for each data set (Appendix E8 [online]) (2). QUADAS-2 divides sources of bias into four categories, including patient selection, index test, reference standard, and flow and timing (21,22). Incomplete reporting of major features was flagged using QUADAS-2 under the flow and timing domain. Risk of bias and applicability assessment were performed in duplicate and independently by two authors (C.B.v.d.P., J.P.S.; each with experience conducting risk of bias assessment for diagnostic test accuracy studies), and differences were resolved by discussion with a third author (M.D.F.M.). A pilot of one study with subsequent discussion was performed by these three authors to improve subsequent interobserver agreement.

Figure 1: Flow diagram shows search results, study review, and study inclusion. CENTRAL = Cochrane Central Register of Controlled Trials, IPD = individual patient data.

Diagnostic Accuracy Measures

The main model estimates of interest were odds ratios (ORs) to determine the association of each LI-RADS major feature with a diagnosis of HCC, both independently and in combination.

Synthesis of Results

All data were pooled into a master data set, with each observation assigned a unique identifier. IPD received from primary study investigators were compared against the published reports for each study. When data were unclear or inconsistent, primary study investigators were contacted to resolve the differences (nine studies). When multiple readers were present, the data from one reader was chosen at random (six studies). One study included observations made with an extracellular contrast agent and the same observations made again with a hepatobiliaryspecific contrast agent. Examinations using the contrast agent less represented in our cohort (gadoxetic acid) were included to improve representation of that agent.

Statistical Analysis

We used a one-step IPD meta-analysis approach to pool the IPD across studies and model them simultaneously to compute OR for the association of each LI-RADS major feature with HCC (Appendix E9 [online]). Liver observation clustering was addressed at the study and patient levels through random intercepts. The ORs for all the variables are presented with 95% CIs. Collinearity between variables was assessed by calculating the variance inflation factor, the tolerance statistic, and eigenvalues. The strength of the association of the variables with the outcome of interest was determined based on the statistical significance and the magnitude of the ORs derived in the multivariable model. A sensitivity analysis was performed by limiting the same analyses to studies at overall low risk of bias. Forest plots show individual study results. τ^2 was used to quantify heterogeneity, and funnel plots were generated to demonstrate publication bias. The level of significance was set at $P < .05$. All analyses were performed by study authors (J.P.S., B.L.) using the glmer function in the Lme4 package in R (R Core Team, version 4.0.0; R Foundation for Statistical Computing) (23).

Results

Study Selection and Characteristics

A total of 865 studies were identified during the initial search, with 466 remaining once duplicates were removed. On title and abstract review, 161 studies were identified for possible inclusion (Fig 1). After full-text review, authors of 81 studies were invited to collaborate, 47 of whom agreed and 37 of whom provided data (Appendix E10 [online] lists studies whose authors did not respond). Studies were then excluded at this stage for the following reasons: incomplete and redundant data with other studies (24,25), multisite data not readily available (26), only patient-level and not observation-level data were available (27), and data formatting issues precluding extraction of relevant parameters (28). The final cohort included 32 studies, including 28 articles and four published conference abstracts (Table 1) (29–60).

Risk of Bias and Applicability

Of the 32 studies, seven were considered at low risk of bias, and 23 had low concern regarding applicability (Fig 2). Study flow

Note.— Data were averaged if there was more than one reader. Data in parentheses are percentages. Studies were considered case control studies if groups of patients were selected based on final diagnosis and then imaging findings were compared. Gadobenate dimeglumine was documented as a hepatobiliary contrast agent if hepatobiliary phase imaging was used; otherwise, it was considered an extracellular agent. Can = Canada, Chi = China, CCRS = composite reference standard, CEUS = contrast-enhanced US, ECA = extracellular contrast agent, HPB = hepatobiliary contrast agent, Kor = Korea, NR = not recorded, P = pathology, PC = prospective cohort, RC = retrospective cohort, RCCon = retrospective case control, Ref = reference, USA = United States.

* The symbol indicates the first risk factor was more represented in the cohort than the second risk factor. The ~ symbol indicates both risk factors were represented approximately equally. The \gg symbols indicate that the first risk factor was substantially more represented in the cohort than the second risk factor.

and timing was the domain most often at risk for bias, which was usually due to unclear or inappropriate intervals between the index test and the reference standard or verification bias (from tissue sampling of only a subset of observations) (Table E1 [online]). Patient and observation selection were also frequently at risk for bias due to multiple studies with case-control design and studies limited to only patients with malignant lesions. The index test and reference standard domains were at low risk of bias for most studies. Funnel plots are available in Appendix E11 [online].

Synthesis of Results

All observations were classified using either pathology or the composite reference standard (Table 2).

*CT/MRI.—*A total of 1170 observations obtained with CT in 812 patients from six studies and 3341 observations obtained with MRI in 2639 patients from 17 studies had sufficient data to be incorporated into the model (Table E2 [online]). From these cohorts, 813 observations were obtained with both CT and MRI. All five major features had been assessed for 887 ob-

Overall Risk of Bias and Applicability Assessment

■Low risk ■Concerns

Figure 2: Quality Assessment of Diagnostic Accuracy Studies 2 (QUADAS-2) risk of bias assessment. CEUS = contrast-enhanced US.

servations, while 3547 observations included assessment of all major features except threshold growth. Threshold growth was the major feature that was reported least often; authors reported that this was due to a lack of prior imaging examinations available for comparison in 75% of studies (21 of 28), rather than a feature of study design.

Results of the univariable analyses are presented in Table E3 (online), which found all five major features to be associated with HCC. Based on the clinical and statistical significance in the univariable analyses, the size variable was investigated as a categorical variable in the multivariable analysis. On multivariable analysis of the cohort including only observations with all five major features assessed (*n* = 887 observations), all major features were associated with HCC except threshold growth (OR, 1.6; 95% CI: 0.7, 3.6; *P* = .07) (Fig 3, Table 3). Multivariable analysis was repeated on the subset of observations, with all major features assessed except threshold growth (3547 observations). In this subset, we did not find evidence of association of enhancing capsule with HCC (OR, 1.3; 95% CI: 0.7, 2.5; *P* = .08). However, when multivariable analysis was performed on the largest cohort including observations with APHE, enhancing capsule, and nonperipheral washout consistently reported (4434 observations), each was associated with HCC. The variance inflation factor, tolerance statistic, and eigenvalues were computed and were within guidelines (61). On sensitivity analysis limited to observations from studies at low risk of bias, these associations persisted (Table E4 [online]).

Using the largest cohort of observations with APHE, enhancing capsule, and nonperipheral washout consistently

Table 2: Observation Diagnosis and Reference Standard

criteria were fulfilled on another imaging modality study and there was threshold growth, or if LR-5 criteria were fulfilled and recurred after local-regional treatment on CT or MRI scans based on treatment response criteria. Other malignancies required histopathology for confirmation. LR-3, LR-4, and LR-M observations with recurrence on CT or MRI scans after local treatment were considered malignant but not specifically indicative of HCC.

reported ($n = 4434$), observation size smaller than 10 mm was associated with decreased odds of HCC diagnosis (OR, 0.1; 95% CI: 0.0, 0.2; *P* = .01) compared with a size of 10–19 mm. Observation size of at least 20 mm was not associated with HCC compared with a size of 10–19 mm (OR, 1.6; 95% CI: 0.95, 2.7; *P* = .06).

Of all CT/MRI major features, differences between CT and MRI were found only for observation size. For patients with observations of at least 10 mm on MRI scans, the odds of having HCC were higher than for those imaged with CT, namely 3.6 (95% CI: 1.04, 12.4; *P* = .04) for 10–19-mm observations and 3.1 (95% CI: 1.9, 5.1; *P* = .03) for observation 20 mm or larger.

*CEUS.—*A total of 853 observations were imaged using CEUS in 833 patients from six studies, and assessments of all major features were available. Results of the univariable analysis are presented in Table E5 [online]. On multivariable analysis, the following were associated with HCC: nonrim APHE (OR, 7.3; 95% CI: 4.6, 11.5; *P* = .01), late and mild washout (OR, 4.1; 95% CI: 2.6, 6.6; *P* = .01), and size of at least 20 mm (OR, 1.6; 95% CI: 1.04, 2.5; *P* = .04) compared with 10–19-mm observations (Table 4). Rim or peripheral discontinuous globular enhancement was associated with decreased odds of HCC (OR, 0.3; 95% CI: 0.1, 0.9; *P* = .02), as was early $(<$ 60 seconds) washout (OR, 0.3; 95% CI: 0.1, 0.5; $P = .03$). Marked washout was not associated with diagnosis or nondiagnosis of HCC (OR, 0.7; 95% CI: 0.2, 2.8; $P = .10$). A sensitivity analysis including only the two CEUS studies at overall low risk of bias was not performed due to model instability.

Forest plots show individual study results (Appendix E12 [online]), and τ^2 was used to quantify heterogeneity (Appendix E13 [online]).

Discussion

This individual patient data (IPD) meta-analysis of 1170 CT, 3341 MRI, and 853 contrast-enhanced US (CEUS) observations found that all Liver Imaging Reporting and Data System (LI-RADS) major features were independently associated with hepatocellular carcinoma (HCC) except for threshold growth. For CT/ MRI, nonperipheral washout (odds ratio [OR], 13.2; 95% CI: 9.0, 19.2; *P* = .01) and nonrim arterial phase hyperenhancement (APHE) (OR, 10.3; 95% CI: 6.7, 15.6; *P* = .01) had the strongest association with HCC, followed by enhancing capsule (OR, 2.4; 95% CI: 1.7, 3.5; *P* = .03). Threshold growth

Figure 3: Multivariable analysis odds ratios with 95% CIs (error bars) for the association of each CT/MRI and contrastenhanced US (CEUS) Liver Imaging Reporting and Data System major feature with a diagnosis of hepatocellular carcinoma. APHE = arterial phase hyperenhancement.

Note.—Data in parentheses are 95% CIs. Reference categories are as follows: for size, 10–19 mm; for nonrim arterial phase hyperenhancement (APHE), absent; for enhancing capsule, absent; for nonperipheral washout, absent.

*References 29, 32, 37, 43, 44, 54, 55, 57, and 59.

† References 29, 30, 32, 33, 35–37, 40, 41, 47, 48, 53, 54, 56, 57, 60.

‡ References 29, 30, 32, 33, 35–37, 40, 41, 43, 44, 47, 48, 53–57, 59, and 60.

 $\,^{\circ}$ MRI and CT were compared for all major features and were only found to differ for size. The bottom rows list odds ratios for size cutoffs comparing MRI and CT, with CT as the reference standard.

was infrequently reported (mostly due to a lack of available prior examinations followed by study design) and was not independently associated with HCC (OR, 1.6; 95% CI: 0.7, 3.6; *P* = .07). For CEUS, nonrim and nonperipheral discontinuous globular APHE had the strongest association with HCC (OR, 7.3; 95% CI: 4.6, 11.5; *P* = .01), whereas late and mild washout

nonperipheral washout, absent.

* References 34, 38, 42, 51, 52, and 58.

was also associated with HCC (OR, 4.1; 95% CI: 2.6, 6.6; *P* = .01). Early $(<$ 60 seconds) washout was associated with a non-HCC diagnosis (OR, 0.3; 95% CI: 0.1, 0.5; *P* = .03), whereas marked washout was not useful for differentiating between HCC and non-HCC (OR, 0.7; 95% CI: 0.2, 2.8; *P* = .10).

Prior studies exploring the diagnostic performance of LI-RADS major features for establishing HCC have mostly focused on the sensitivity, specificity, and predictive values of individual imaging features and the LI-RADS categories (62). Multivariable modeling that includes all LI-RADS major features to establish the relative strength of association of each feature with HCC may not have been possible in single centers or with study-level meta-analyses (32,63,64).

Observation size did not have an association with HCC when treated as a continuous variable at univariable analysis. Associations were observed for larger observations using cutoffs of 10, 15, and 20 mm. However, at multivariable analysis, size of at least 20 mm was not significantly associated with HCC compared with size of 10–19 mm. Size of at least 20 mm was associated with HCC for the smaller cohort with threshold growth reported, likely due to decreased interstudy variability for these observations. Interstudy variability likely also explains the difference between cohorts when comparing CT and MRI.

Our findings suggest that threshold growth is not a significant predictor of HCC relative to the other LI-RADS major features. Of note, CT/MRI LI-RADS version 2014, 2017, and 2018 were included. The criteria for threshold growth in LI-RADS version 2018 was limited to a size increase of at least 50% of a mass in no more than 6 months, whereas for LI-RADS versions 2014 and 2017, threshold growth also included a size increase of at least 100% on imaging examinations more than 6 months apart and new observations of 10 mm or larger in 24 months or less. The impact of this change could not be further explored, as the specific criterion for threshold growth using LI-RADS versions 2014 and 2017 could not be retrospectively identified.

Prior CEUS LI-RADS studies mostly explore the performance of the LI-RADS categories for HCC diagnosis (26,65–67). Our finding that both nonrim and nonperipheral discontinuous globular APHE and late and mild washout each had strong association with HCC corroborate prior works on the CEUS imaging characteristics of HCC and support the application of these as major features in the LI-RADS framework (68).

Most studies (25 of 32) had a high risk of bias. A prolonged interval between the index test and the reference standard was frequent. However, an optimal interval remains uncertain. Verification bias from tissue sampling of only a subset of observations is another frequent potential source of bias, as many lowerrisk observations are less likely to have histopathologic proof. Despite the high risk of bias, a sensitivity analysis including only low-risk-of-bias studies confirmed the findings, increasing confidence in the results for the larger cohort.

Our study had several limitations. First, less than half of the eligible data sets were made available, which precluded a more detailed analysis. The findings of this study must be interpreted in the broader context of the liver imaging literature, and we recommend clinicians refer to the current version of LI-RADS until a future iteration is released. Second, threshold growth was not reported for most observations, and it remains unclear if its predictive value may differ when applied at centers that routinely perform CT/MRI within 6-month intervals. Third, ancillary features were not included, because they were incompletely reported in most studies. Fourth, comparison of extracellular and hepatobiliary contrast agents was not performed, because only one study included direct comparison of each contrast agent using the same observations, and our data set was skewed toward extracellular agent. Finally, we required only 12 months of stability to establish benignity. A longer interval might have been more specific but at the expense of sensitivity.

In conclusion, the CT/MRI and contrast-enhanced US (CEUS) Liver Imaging Reporting and Data System (LI-RADS) major features have different independent associations with hepatocellular carcinoma (HCC). Arterial phase hyperenhancement and washout pattern have strong independent associations with HCC using CT/MRI and CEUS LI-RADS. Threshold growth was infrequently reported and was not a significant independent predictor of HCC. The utility of ancillary features, which were not included in our study, would benefit from more comprehensive reporting in future research.

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