

Article **Frequency of Italian Record-Breaking Floods over the Last Century (1911–2020)**

Attilio Castellarin ¹ [,](https://orcid.org/0000-0002-6111-0612) Andrea Magnini 1,*, Kay Khaing Kyaw ¹ [,](https://orcid.org/0009-0009-8940-1161) Filippo Ciavaglia ¹ , Miriam Bertola ² , Gunter Blöschl ² , Elena Volpi ³ [,](https://orcid.org/0000-0002-9511-1496) Pierluigi Claps ⁴ [,](https://orcid.org/0000-0002-9624-7408) Alberto Viglione ⁴ , Alberto Marinelli ¹ and Richard M. Vogel [5](https://orcid.org/0000-0001-9759-0024)

- ¹ Department of Civil, Environmental, Chemical and Materials Engineering (DICAM), University of Bologna, 40126 Bologna, Italy
- 2 Institute of Hydraulic Engineering and Water Resources Management, Vienna University of Technology, 1040 Vienna, Austria
- ³ Engineering Department, University of Roma Tre, 00154 Rome, Italy
- ⁴ Department of Environment, Land and Infrastructure Engineering, Polytechnic University of Turin, 10129 Turin, Italy
- ⁵ Department of Civil and Environmental Engineering, Tufts University, Medford, MA 02155, USA
- ***** Correspondence: andrea.magnini@unibo.it

Abstract: This study provides an in-depth analysis of the frequency of extreme streamflow in Italy, adopting the innovative perspective of the theory of records, and focusing on record-breaking floods. (i.e., annual maximum series, AMS) observed in Italy between 1911 and 2020. Our research employs an extensive dataset of 522 annual maximum series (AMS) of streamflow observed across Italy between 1911 and 2020. We consider three time intervals (1911–2020, 1911–1970, and 1971–2020), and we define pooling-groups of AMSs based on (a) hydrological (e.g., catchment size, mean annual precipitation, etc.) and (b) spatial proximities of the gauged sites. First, within each group and for each time period, we compute the regional average number of record-breaking events (NRbins). Second, with a series of resampling experiments that preserve the spatial correlation among the AMSs, we test the hypothesis that NRbins result from a group of stationary sequences. Our results show spatially coherent patterns of an increasing number of record-breaking floods in central and in northeastern Italy over the last 50 years. In the same time interval, significant deviations in the regional number of record-breaking events from what would be expected for stationary flood sequences seem to be more common in drier climates or at higher altitudes, while the catchment size does not seem to be a meaningful descriptor.

Keywords: floods; record-breaking; climate change; bootstrap sampling

1. Introduction

The economic losses and social consequences caused by floods have been steadily increasing over the last five decades worldwide [\[1](#page-12-0)[,2\]](#page-12-1). Detecting changes in flood frequency is a topical research issue, and the scientific community calls for a common effort to better understand recent flood dynamics and their spatiotemporal evolution.

An increasing number of studies investigate the presence of non-stationarities in the frequency and magnitude of observed and projected peak flows at both the global [\[3](#page-12-2)[,4\]](#page-12-3) and continental scale $[5-7]$ $[5-7]$. With specific reference to Europe, recent studies show a complex system of changes in frequency and magnitude of observed floods over the last six decades [\[8](#page-12-6)[–10\]](#page-12-7). In particular, Ref. [\[11\]](#page-12-8) observes that in the last five decades, floods dynamics have changed dramatically across Europe. Regional patterns are particularly evident in some areas, e.g., Southern Europe, yet sub-regional patterns (e.g., specific areas of specific regions) may be masked by the filtering criteria (e.g., minimum sample length) adopted in these studies. Several studies investigate the presence of non-stationary patterns at a

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smaller scale (i.e., national [\[12–](#page-12-9)[14\]](#page-12-10) or regional [\[15,](#page-12-11)[16\]](#page-13-0)), considering local morpho-climatic effects and uncertainty due to data scarcity, inaccuracy, or fragmentation ([\[17](#page-13-1)[,18\]](#page-13-2)).

A considerable number of publications adopt a non-stationary framework for modeling flood statistics [\[19](#page-13-3)[,20\]](#page-13-4). Some of these focus on particular elements of flood dynamics in single river catchments [\[21\]](#page-13-5), while others compare different models for frequency analysis of annual maxima in small-to-medium regions [\[22\]](#page-13-6).

Generally, the theoretical frameworks adopted for this type of analysis are the extreme value [\[23\]](#page-13-7) and time series analysis [\[24\]](#page-13-8) theories. These aim to define the best-fit probability distribution that describes the occurrence of extreme values (e.g., streamflow and precipitation [\[19,](#page-13-3)[25\]](#page-13-9)) and to provide non-parametric tests to detect non-stationarity in the mean or other statistics of the variables of interest (e.g., the Mann–Kendall test, see [\[26\]](#page-13-10)). However, the extreme value theory is very sensitive to the length of the time series and the possible presence of outliers, which may negatively affect the reliability of the fitted probability distribution [\[27](#page-13-11)[,28\]](#page-13-12).

The theory of records, or of record-breaking events, is still underutilized in this research area, despite its relevance and potential [\[29\]](#page-13-13). This theory investigates the statistical properties of the occurrence and magnitude of those events that exceed (upper records) or are exceeded (lower records) by any other formerly observed events (herein referred to as +records and −records, respectively). This theory offers a framework for analyzing the frequency of extreme events which is nearly independent of the probability distribution of the original variable—in our case, annual maximum discharge [\[30,](#page-13-14)[31\]](#page-13-15)—and which does not require the probability distribution to be fit to the observed data, thus being less sensitive to the length of the time series than standard flood frequency analyses. For these reasons, it has been widely and successfully applied for assessing the possible presence of nonstationarity in time sequences of several variables, such as temperature [\[32\]](#page-13-16), monthly and daily precipitation [\[33](#page-13-17)[,34\]](#page-13-18), and earthquake magnitudes [\[35\]](#page-13-19), and in a variety of research fields (e.g., sports [\[36\]](#page-13-20), biology [\[37\]](#page-13-21), etc., see e.g., [\[38\]](#page-13-22)). It is worth mentioning that the analysis of non-stationarity in the extremes, based on record-breaking events, recently inspired a software tool for statistical testing [\[39\]](#page-13-23). Thus, the application of the theory of records to annual maximum flood sequences can be very useful to improve our knowledge about flood frequency for events of high and very high magnitude [\[40\]](#page-13-24).

Our study contributes to this research field with an in-depth analysis of the frequency regime of record-breaking floods in Italy. In particular, we focus on an extensive dataset of flood sequences (i.e., annual maximum series, AMS) observed at 522 stream gauges in Italy between 1911 and 2020 and we consider the entire observation period, as well as two consecutive time intervals: 1911–1970 and 1971–2020. We look at two different kinds of record-breaking events in a time series, that is the events with the maximum (+record) or minimum (−record) intensity up to that year, and we focus on the average number of record-breaking events in a region (i.e., in a pooling-group of AMSs), and in the three time-intervals of interest. By performing a series of resampling experiments that preserve the spatial correlation among the AMSs of the pooling group, we test the hypothesis that the observed regional average number of record-breaking events results from a group of stationary sequences. We identify the pooling group of sites by referring to (a) hydrological and (b) spatial proximity. In the first case, we group catchments that are similar in terms of size, mean annual precipitation, mean annual snow depth, elevation, or latitude; in the second case, we use a moving square pooling window of 90 \times 90 km².

Our research questions are as follows: (1) Does the number of record-breaking floods in Italy deviate from what is expected under the *iid* hypothesis? If so, (2) are recent deviations larger or more frequent than those in the past? (3) Are larger deviations associated with specific sub-regions, or morphological and climatic characteristics?

The main aim of the study is to assess and understand the presence of non-stationarities in the frequency of occurrence of record-breaking events in Italy. Our study is complementary (for the Italian territory) to the work carried out in Ref. [\[8\]](#page-12-6), as it uses a different perspective and theoretical framework (i.e., analysis of the frequency of record-breaking

events instead of the analysis of the magnitude of extreme values) and a much finer spatial resolution (i.e., more sequences of annual maximum floods). These latter features are elements of novelty relative to those employed in previously published studies, combined with the original and general framework adopted for the analyses, which we present in the following paragraph. *Atmosphere* **2024**, *15*, x FOR PEER REVIEW 3 of 15

2. Data and Methods

2.1. Study Area

Our analysis focuses on 522 annual maximum series (AMS) of flood flows collected across Italy $[41]$ between 1911 and 2020. Beside the total time period (i.e., 1911–2020), the 1911-1970 and 1971-2020 sub-intervals are also considered and studied in the present research[. Th](#page-13-26)ese have been selected based on the findings of Refs. [11,42], who clearly demonstrate that the dynamics of flood events have experienced a dramatic change across Europe over the last five decades. Table 1 illust[ra](#page-2-0)tes the variability of the AMS record length over the three different time intervals considered in the study. Figure [1a](#page-2-1) depicts the geographical distribution of the study stream gauges, their drainage areas, and the availability of the observed annual maxima in the three time intervals considered.

Table 1. Record length statistics for the annual maximum series (AMS) of the study for the flood flood flows collected in the flood floo flows in the three different time intervals considered the analyses.

Figure 1. Study area and basins (a); location of study area (b); discretization of the study area into 90 × 90 km and 30 × 30 km tiles (**c**). 90 × 90 km and 30 × 30 km tiles (**c**).

A number of catchment descriptors are retrieved. These include the basin's size (A), mean annual precipitation (MAP), mean annual snow water equivalent (SWE), minimum elevation (ME), and latitude of catchment centroid (LC). All morphological descriptors (i.e., A, ME, LC) were computed from the 90 m Multi Error Removed Improved Terrain Data Elevation Model (MERIT DEM, [43,44]), while catchment scale climatic indices (i.e., MAP and SWE) were computed from the 1 km gridded datasets available from the BIGBANG project (Italian GIS-b[ased](#page-14-1) national water balance, see [45]). These static (i.e., time-invariant) catchment descriptors are selected based on the methods of previous relevant studies [\[46–](#page-14-2)[48\]](#page-14-3).

Figure 2 reports the elevation, mean annual precipitation, and mean annual snow Figure [2](#page-3-0) reports the elevation, mean annual precipitation, and mean annual snow water equivalent for the Italian territory for a better characterization of the study area. water equivalent for the Italian territory for a better characterization of the study area.
The main mountain chains are the Alps in the north, with the highest elevation, and the Apennines, which cross all the peninsula (Figure 2a[\).](#page-3-0) The MAP decreases from north to south, and generally increases where the elevation is higher (Figure 2b). [Th](#page-3-0)e SWE is basically different from zero only in mountain areas and in the north (Figure [2c](#page-3-0)).

Figure 2. Elevation (DEM, panel (a)), mean annual total precipitation $(MAP, (b))$, and mean annual snow water equivalent (SWE, (**c**)) in the study area. In red, the catchments considered in the study. snow water equivalent (SWE, (**c**)) in the study area. In red, the catchments considered in the study.

2.2. Number of Record-Breaking Events in Independent and Identically Distributed (iid) Series 2.2. Number of Record-Breaking Events in Independent and Identically Distributed (iid) Series

A record-breaking event, or record event, is defined as an event whose magnitude A record-breaking event, or record event, is defined as an event whose magnitude either exceeds (+record) or is exceeded (−record) by all previously observed events [\[29](#page-13-13)[,49\]](#page-14-4); the first observation in a sequence is generally counted as a (trivial) record. Let us consider the first observation in a sequence is generally counted as a (trivial) record. Let us consider a sequence of independent and identically distributed (*iid*) random variables of length *n n*; then, the number of records in the series *R* is a random variable, whose probability distribution of the number of records in the series *R* is a random variable, whose probability distribution is totally independent of the probability distribution of the original variable (e.g., $\frac{1}{2}$) (e.g., annual maximum floods in a given cross-section) with mean, μ_R , and variance, σ_R^2 , defined as: defined as:

$$
\mu_R = \sum_{i=1}^n 1/i
$$
 (1)

$$
\sigma_R^2 = \sum_{i=1}^n 1/i - \sum_{i=1}^n 1/i^2
$$
 (2)

Figure [3](#page-4-0) illustrates the progression of μ_R (black line) with *n*, together with the 2.5th and 97.5th percentiles of *R*. It is interesting to note the rather limited expected number of records for an *iid* sequence with a length of 100 elements. This characteristic renders the
rumber of record breaking exerts on interesting variable for detecting pessible shapese and 97.5th percentiles of *R. It is interesting to and the rather limited in the rather* limited number of α *R* in the frequency distribution of the variable of interest (e.g., annual maximum flooding).
In the frequency distributions in the variable of magnetic resulting surprising an absenced time. number of μ interest α records and interesting variable for detecting possible cases and interesting μ series relative to what is expected under the *iid* hypothesis indicate non-stationary time
cories [20] number of record-breaking events an interesting variable for detecting possible changes Indeed, significant deviations in the number of record-breaking events in an observed time series [\[30\]](#page-13-14).

time series [30].

2.3. Testing the Statistical Significance of the Deviation in the Average Number of Records in Pooling-Groups of AMS of Flood Flows Relative to What Is Expected under the iid Hypothesis Pooling-Groups of AMS of Flood Flows Relative to What Is Expected under the iid Hypothesis 2.3. Testing the Statistical Significance of the Deviation in the Average Number of Records in

groups of AMS of flood flows. We refer to three different historical time periods, namely groups of Flood in periods, namely periods, namely periods, namely periods, namely periods, namely periods, namely periods, na floods according to two criteria, i.e., hydrological similarity and spatial proximity of the catchments. Both criteria will be better described in Section [2.4.](#page-5-0) In this study, we focus on the average number of record-breaking events of pooling-

The main objective of the study is to assess whether the number of record-breaking events observed in a given pooling-group of AMS is compatible with the *iid* hypothesis (i.e., the AMS of flood flows in the pooling-group may be regarded as a stationary series relative to the frequency of record-breaking events).

First, we compute the number of record-breaking events (for both +record and −record events) in each AMS of the pooling-group. Suppose we have a total of *m* series in the group; it is worth noting here that each AMS i , with $i = 1, 2, \ldots, m$, has its own number of records n_i . Then, we compute the mean number of records in the pooling-group. Finally, we assess the deviation of this mean value from what would be expected for the same group of AMS (i.e., the group counts m AMS in total) under the *iid* hypothesis (i.e., each AMS consists of n_i realizations of independent and identically distributed random variables, with $i = 1, 2, \ldots, m$). This is achieved by comparing the average number of records observed in the pooling-group with the distribution of the average number of records in synthetic pooling-groups of *iid* series having the same sample length of the original series, which we obtained through a bootstrap resampling experiment [50]. In particular, we design the resampling experiments to preserve the existing spatial correlation between annual maximum floods observed in the same year at different stream-gauges, which cannot be neglected when testing the statistical significance of a null hypothesis at a regional scale (on the possible impact of spatial correlation of time series on statistical testing at regional scale see also [\[30,](#page-13-14)[51](#page-14-6)[,52\]](#page-14-7)). We randomly shuffle 10,000 times the group of AMS year-wise, so that the sequence of data in each time series differs from the original series, yet the distribution of observations among the various AMS in any given year is preserved, together with the AMS length n_i , with $i = 1, 2, \ldots, m$. For each synthetic (i.e., reshuffled) dataset, we compute the regional average number of records as described above for the original pooling-group of AMS. Then, we compare the empirical regional average number of record events with the distribution of 10,000 synthetic regional means obtained through the bootstrap resampling experiments, to test the statistical significance of the observed deviations.

2.4. Criteria for Pooling-Groups Definition 2.4. Criteria for Pooling-Groups Definition

deviations.

First, we analyzed hydrologically homogeneous pooling-groups of catchments, group-First, we analyzed hydrologically homogeneous pooling-groups of catchments, ing the dataset according to the five catchment descriptors mentioned in Section [2.1:](#page-2-2) size (A), mean annual precipitation (MAP), mean annual snow water equivalent (SWE), minimum catchment elevation (ME), and latitude of the catchment centroid (LC). The objective of these experiments is to group hydrologically similar catchments and to verify whether the frequency of record-breaking events at the national scale shows non-stationarity; if so, the non-stationary behavior depends on the catchment descriptors.

We grouped the study catchments into a total of 36 classes by combining the size of We grouped the study catchments into a total of 36 classes by combining the size of the catchments (three classes) with one of the four remaining catchment descriptors (three classes each). For each of the 36 classes, we calculated the regional mean number of record classes each). For each of the 36 classes, we calculated the regional mean number of record events (+records and −records) by referring to the three time periods, namely 1911–2020, events (+records and −records) by referring to the three time periods, namely 1911–2020, 1911–1970, and 1971–2020. The regional mean number of records (NRbin) was then com-1911–1970, and 1971–2020. The regional mean number of records (NRbin) was then compared with 10,000 synthetic mean regional records obtained from the bootstrap experiment described above, and the deviation of the empirical value from what is expected under the *iid* hypothesis assessed at a 10% significance level. The 36 classes of this assessment are presented in Fi[gu](#page-5-1)re 4 for the four combinations (i.e., A-MAP, A-SWE, A-ME, A-LC) and the three different time intervals considered in our study (i.e., 1911–2020, 1911–1970, 1971–2020). 1970, 1971–2020).

Figure 4. Size (y axis) and other descriptors (x axis) of the study area: mean annual precipitation (a), mean annual snow precipitation (b) , outlet elevation (c) , latitude of catchment centroid (d) . Colors of the dots follows the legend of Figu[re](#page-2-1) 1.

The second part of the study aims to detect the spatial pattern of changes in the fre-The second part of the study aims to detect the spatial pattern of changes in the frequency of record-breaking events. This is performed by grouping catchments based on quency of record-breaking events. This is performed by grouping catchments based on spatial proximity. We consider a discretization of Italy into square grid cells of 30 km \times 30 km in size, and we pool AMS observed at stream gauges located in any given cell, as well as in the eight surrounding ones (i.e., overall size of the pooling-group: $90 \text{ km} \times 90 \text{ km}$). This spatial subdivision can be visualized in Figure [1c](#page-2-1). Specifically, only 90 km \times 90 km tiles containing at least five stream gauges are considered when plotting the results so that the outcomes for any tile are never associated with at-site, or local, conditions and are therefore, more robust.

This spatial subdivision can be visualized in Figure 1c. Specifically, one 1 and 90%

Analogously to what has been described for the hydrological grouping criterion, for each pooling-group of AMS, we compute the regional average number of records (both for +record and −record events) and compare it with the empirical values with the 10,000 synthetic bootstrap regional number of records, focusing on the previously mentioned three time periods.

In summary, in the present study, the possible presence of non-stationarity in annual In summary, in the present study, the possible presence of non-stationarity in annual maximum flood sequences is assessed through the theory of records, and statistical hy-maximum flood sequences is assessed through the theory of records, and statistical hypothesis testing is performed on pooling-groups of the catchments. Thus, the proposed pothesis testing is performed on pooling-groups of the catchments. Thus, the proposed framework assesses the statistical significance of the results at a regional level, instead of framework assesses the statistical significance of the results at a regional level, instead of locally, by looking at single time series. According to several recent studies [\[17,](#page-13-1)[18,](#page-13-2)[53,](#page-14-8)[54\]](#page-14-9), locally, by looking at single time series. According to several recent studies [17,18,53,54], most statistical approaches used for detecting trends or non-stationarity would require very most statistical approaches used for detecting trends or non-stationarity would require protracted observed time series to perform optimally; these are seldom available, and are very sparse in regards to space. Moreover, time series are often affected by measurement errors and fragmentation. The aggregation of time series into pooling-groups (i.e., regional analysis) helps in reducing the uncertainty and increasing the robustness of these types of analyses.

3. Results 3. Results

3.1. Pooling-Groups of Hydrologically Similar Catchments 3.1. Pooling-Groups of Hydrologically Similar Catchments

For each single pooling-group, the distribution of NRbin is computed using a bootstrap method, as described in Section [2.3.](#page-4-1) This distribution essentially reproduces the natural variability of the NRbin in the *iid* sequences, and is represented via boxplots in Figures 5[–8.](#page-8-0) The NRbin computed using the original time series is represented as a red circle for each pooling-group and each time interval. In cases where NRbin falls outside the confidence interval for *iid* sequences, a red triangle is adopted.

Figure 5. Average number of +(panels (**a**–**i**)) and −(panels (**j**–**l**,**p**–**u**)) records for hydrologically similar catchments. The pooling descriptors include catchment area (y axis) and mean annual precipitation (MAP, x axis). Values for the boxplots are obtained through bootstrap experiments; a red circle represents the average number of records for each pooling-group. Average numbers of records falling outside the confidence interval are marked with red triangles.

falling outside the confidence interval are marked with red triangles.

falling outside the confidence interval are marked with red triangles.

Figure 6. Average number of $+($ panels $(a-i)$) and $-($ panels $(j-l,p-u)$) records for hydrologically similar catchments. The pooling descriptors include catchment area (y axis) and mean annual snow water equivalent (SWE, x axis). Values for the boxplots are obtained through bootstrap experiments; red circles represent the average number of records for each pooling-group. Average numbers of records f_{ellip} outside the confidence interval are marked with red triangles. Also see the legend in Figure 5. falling outside the confidence interval are marked with red triangles. Also see the legend in Figure [5.](#page-6-0) charge represent the average number of records for each pooling-group. Average numbers of records

Figure 7. Average number of $+($ panels $(a-i)$) and $-($ panels $(j-l,p-u)$) records for hydrologically similar catchments. The pooling descriptors include catchment area (y axis) and minimum elevation (ME, or outlet elevation, x axis). Values for the boxplots are obtained through bootstrap experiments; red circles represent the average number of records for each pooling-group. Average numbers of records falling outside the confidence interval are marked with red triangles. Also see the legend [in](#page-6-0) Figure 5. falling outside the confidence interval are marked with red triangles. Also see the legend in Figure 5.

Figure 8. Average number of +(panels (a–i)) and -(panels (j–l,p–u)) records for hydrologically similar catchments. The pooling descriptors include catchment area (y axis) and latitude of the catchment's centroid (LC, x axis). Values for the boxplots are obtained through bootstrap experiments; red dots represent the average number of records for each pooling-group. Average numbers of records falling outside the confidence interval are marked with red triangles. Also see the legend in Figure [5.](#page-6-0)

dots outside the confidence interval of 5–95% of the boxplot) in several cases. Regarding the first hydrological criterion (i.e., catchment area and mean annual precipitation), six cases are evident. The number of higher records shows an increase during the 1971–2020 time interval for intermediate- and large-sized catchments with low (Fig[ur](#page-6-0)e 5d,g) and intermediate MAP (Figure 5e,h). The lower records show an increase for the whole 1911–2020 time interval in drier intermediate catchments (Figure 5p), and a decrease f[or](#page-6-0) the 1971–2020 time interval in smaller catchments with intermediate MAP (Figure 5k). The number of records is significantly different from that in the *iid* hypothesis (i.e., red

Regarding the second criterion (i.e., catchment area and mean annual snow precipitation), three cases of significant non-stationarity are observed. The higher records exhibit an increase during the 1971–2020 time interval for intermediate-sized catchments with average snow precipitation (Figure 6e) and larger-sized catchment[s w](#page-7-0)ith higher SWE (Figure 6i). In the same time interval, larger-sized catchments with higher SWE also show a significant decrease in the occurrence of lower records (Figure 6u).

Regarding the third hydrological criterion (i.e., catchment area and outlet elevation), four cases of significant non-stationarity are observed. The +records show an increase [du](#page-7-1)ring the 1971–2020 time interval in intermediate sized catchments with lower (Figure 7d) and higher (Fig[ur](#page-7-1)e 7f) ME, as well as in large catchments with average ME (Figure 7h).

Regarding the fourth hydrological criterion (i.e., catchment area and centroid's latitude), higher records exhibit an increase in non-stationary occurrence for all catchments with central latitude (Figure 8b,e,h).

with central latitude (Figure 8b,e,h). *3.2. Pooling-Groups of Spatially Close Catchments*

Figures 9 and 10 show the average number [of](#page-9-1) records and the statistical significance of the deviation from what would be expected under the *iid* hypothesis for +records $\frac{1}{1}$ and -records, respectively, with a 90×90 km spatial discretization of the study area see Section 2.9). They felter to the empirical p value, i.e., empirical non-exectuance p for ability of the average number of record events in each tile relative to the distribution of the synthetic average number of record events in each the relative to the distribution of the synthetic average number of records resulting from the 10,000 bootstrap realizations. are synthetic average number of records resulting from the 10,000 sociolity reductions.
Specifically, only tiles containing at least five stations are showed. Figures [9](#page-9-0) and [10](#page-9-1) refer to $\epsilon_{\rm p}$ records records records records records records realizations of representations of $\epsilon_{\rm p}$ records $\epsilon_{\rm p}$, $\epsilon_{\rm p}$ the last 50 years, while the results we obtained relative to the time periods 1911 –2020 and (see Section [2.3\)](#page-4-1). They refer to the empirical *p*-value, i.e., empirical non-exceedance prob1911-1970 (not shown here for the sake of brevity) show statistically significant deviations from the *iid* hypothesis regarding the regional number of +record or $-$ record events only in small areas and isolated pixels, if any.

Figure 9. Spatial analysis of the average number of +record events and its statistical significance **Figure 9.** Spatial analysis of the average number of +record events and its statistical significance $(10\%$ significance level) relative to pooling-groups of *iid* series having the same length and spatial correlation as the original sample: (a) average number of +records (see color scale); (b) tiles for which the average number of records variable to the state of the state (1) , (b) the state which the average number of record events is associated with p -values > 0.95 or <0.05 under the *iid* hypothesis. which the average number of record events is associated with p -values > 0.95 or ≤ 0.05 under the *li*

Figure 10. Spatial analysis of the regional average number of -record events and its statistical nificance (10% significance level) relative to pooling-groups of *iidade series of iidade series having the same*
Same length and same length significance (10% significance level) relative to pooling-groups of *iid* series having the same length and spatial correlation as the original sample: (a) average number of $-$ records (see color scale); (b) tiles for which the average number of record events is associated with p -values > 0.95 or <0.05 *iid* hypothesis. under the *iid* hypothesis.

Clear and spatially coherent geographical patterns are visible. For example, we detect Clear and spatially coherent geographical patterns are visible. For example, we detect an intensification in the frequency of record-breaking flood events in central and northeastern (i.e., Triveneto) Italy (Figur[e](#page-9-0) 9). In these areas, +record events are significantly more frequent than what is expected under the *iid* hypothesis, while $-$ records are less frequent than expected for *iid* sequences (F[igu](#page-9-1)re 10). Also, -record events show a higher frequency than for *iid* sequences in southern Italy, particularly for stream basins with river mouths originating in the Adriatic Sea.

4. Discussion

The presented results show that if we consider the entire available period (1911–2020) or the period of 1911–1970, the empirical average numbers of record events (both +record and −record) do not show significant deviations from the *iid* hypothesis, with a few exceptions. In particular, regional numbers of −record events present few, but consistent, significant deviations in the time period of 1911–2020 in smaller and intermediate catchments (Figures [5p](#page-6-0), [7k](#page-7-1) and [8j](#page-8-0)), which may be linked to a decrease in the magnitude of flood events over time.

Differently, in the last 50 years, statistically significant deviations from what would be expected under the *iid* hypothesis have been found for pooling-groups of catchments. Specifically, when considering pooling-groups that are identified based on hydrological sim-ilarity, +records show a larger number of deviations (Figures [5d](#page-6-0),e,g,h, [6e](#page-7-0),i, [7d](#page-7-1),f,h and [8b](#page-8-0),e,h) relative to those of the −records (Figures [5k](#page-6-0) and [6u](#page-7-0)). It is interesting to note that the behavior observed for +record events is in agreement with that of -record events (e.g., an increase in time of the frequency of +record events is coupled with a decrease in the frequency of -record events; the overall magnitude of annual maximum floods is increasing in time) in two cases. This happens for larger catchments with high SWE (see Figure [6i](#page-7-0),u, for +records and −records, respectively) and for pooling-groups of sites (i.e., tiles) in northeastern Italy (see Figures [9b](#page-9-0) and [10b](#page-9-1)).

Regarding the hydrological subdivision of the available catchments, the alterations we detected vary among homogeneous groups, as in some classes, we observe an increase in the number of record-breaking floods, while a statistically significant decrease can be observed in others.

Interestingly, the catchment area does not show a direct and interpretable link with the alterations. Other descriptors appear to have a more structured influence on the number of records. As an example, during the last 50 years, catchments with lower MAP (mostly located in the central southern Italy, see Figure [2b](#page-3-0)) tend to show more statistically significant deviations in the frequency of record-breaking events relative to what would be expected for a stationary annual maximum series (Figure [5\)](#page-6-0). This may be due to the fact that basins with higher MAP, located in the north and at high elevations in the central Italy, can exhibit opposite non-stationarities (see Figures [9](#page-9-0) and [10\)](#page-9-1); thus, aggregating by means of MAP only can hide local signals. Differently, larger catchments with high SWE, which could describe the northeast basins, exhibit significant non-stationarity (Figures [6,](#page-7-0) [9](#page-9-0) and [10\)](#page-9-1).

Notably, these observations are in line with the work of Ref. [\[9\]](#page-12-12); this study detected a flood-rich period in northern Italy over the last decades. Positive trends of flood quantiles showed in Ref. [\[55\]](#page-14-10) are in agreement with the increase in +record occurrence in the northeast (Figure [10b](#page-9-1)). The increase in −records in the south (Figure [10b](#page-9-1)) seems to be related to the negative trends observed over all of southern Europe [\[8](#page-12-6)[,11](#page-12-8)[,55\]](#page-14-10). However, the findings in Figures [9](#page-9-0) and [10](#page-9-1) suggest that, when considering regional/national scales, more complex non-stationary patterns may arise with respect to continental scale studies that ignore smaller basins or shorter series. An interesting example is the non-stationarity signals in central Italy. This area was not marked as a flood change hot-spot in recent continental scale studies, most likely because of the lack of relevant information in such studies. Instead, the analysis framework we propose allows us to consider short time series, therefore increasing the spatial resolution of data coverage in the study area. Also, still concerning differences between our outcomes and those of existing studies, it is very important to underline once more the specific and novel perspective of our study, referring to the frequency of record-breaking events, which is markedly different from the approaches of most of the existing literature. Anomalies in the frequency of +records/−records are related to changes in very large/very small floods rather than to changes in the mean of flood frequency curves, which is usually considered with regression-based trend analyses.

It should be mentioned that a limitation of the present approach consists in the quality and quantity of the input data adopted within the study. In fact, the length and completeness of the time series employed for these analyses can strongly affect the recognition of spatial patterns and non-stationarity [\[53\]](#page-14-8). This is evident when looking at the availability of flood measurements for the 1971–2020 time interval in our study area (Figures [1a](#page-2-1), [9a](#page-9-0) and [10a](#page-9-1)), which is concentrated in specific areas, mostly in central and northern Italy. A possible solution to this problem could be time series reconstruction [\[54](#page-14-9)[,56](#page-14-11)[,57\]](#page-14-12). However, the statistical properties of a synthetic time series would introduce strong and unreliable assumptions.

Nevertheless, studying the behavior of the number of records within the pooling-groups of catchments allows us to simultaneously consider a higher number of measured time series with a variable length, which increases the robustness of the results [\[17\]](#page-13-1). It is positive that significant alterations from the stationary case are not present wherever observations for 1971–2020 time interval are available, but only in specific areas (e.g., compare Figure [9a](#page-9-0) with Figures [9b](#page-9-0) and [10a](#page-9-1) with Figure [10b](#page-9-1)). In fact, while all of northern and part of central Italy express a rather uniform data coverage over the last 50 years, statistically significant deviations are detected only for some sub-regions. Indeed, where data for the 1971–2020 time interval are sparse, the absence of statistically significant non-stationarity signals could be an artifact of data scarcity (see southern Italy, Figures [9](#page-9-0) and [10\)](#page-9-1).

Generally, our findings are consistent across different pooling strategies (see the discussion above), and their agreement with those of previous studies can be taken as proof of their reliability. Future studies could investigate the drivers of floods and their temporal evolution, such as the presence of non-stationary behaviors in time sequences of extreme rainfall and frequencies of record-breaking rainstorms.

5. Conclusions

A large number of studies investigate the presence of non-stationary behaviors in the frequency and magnitude of peak flows at global [\[3](#page-12-2)[,4\]](#page-12-3) and continental scale [\[5–](#page-12-4)[7\]](#page-12-5). Usually, the theory of extreme values and trend analysis is adopted. Differently, our work relates to the theory of records [\[29\]](#page-13-13), and in particular, to the frequency of record-breaking events, which is a complementary perspective that might be very useful to improve our understanding regarding the frequency of extraordinarily intense events [\[30,](#page-13-14)[40\]](#page-13-24). In particular, our study aims to assess the temporal and spatial evolution of the frequency of record-breaking flood events observed in Italy over the last century.

In our study, we define record-breaking events as the annual maximum floods that exceed, or are exceeded, by all other maxima that were previously observed for the same river cross-section. Our analysis considers 522 annual maximum series (AMS) of flood flows and based on the outcomes of some recent continental analysis of flood changes, refers to three distinct time intervals: 1911–2020, 1911–1971, and 1971–2020. For these three time periods, we compute the mean value of the number of record-breaking events in the pooling-groups of AMS identify based on hydrological affinity (expressed in terms of geomorphological and climatic descriptors) or geographical proximity of the corresponding catchments. Then, we compare the empirical values of the regional mean number of records with the average number of record events expected for groupings of *iid* AMS exhibiting the same characteristics, in terms of the number of available observations.

Our results show evident and statistically significant deviations in the frequency of record-breaking floods relative to what is expected under *iid* hypothesis. Stronger signals are associated with the last 50 years of observations, as observed in Ref. [\[11\]](#page-12-8). In particular, dryer catchments show an intensification in the magnitude of annual maximum floods in the last five decades, and hot spots of such an intensification can be found in Central and northeastern Italy. These outcomes are partially in agreement with the findings in recent studies at larger scales [\[8,](#page-12-6)[9,](#page-12-12)[11\]](#page-12-8), but they also highlight the presence of previously unobserved patterns and hotspots of non-stationarity in flood frequency due to the finer scale and higher data coverage of our analyses. We believe that these concerning outcomes call for further investigations, particularly under the most sensitive climatic conditions and in the geographical areas identified in the study, as well as under similar morphological

and climatic conditions found in the Alps, as well as in the Mediterranean Basin, to better assess the degree and extent of flood hazard intensification.

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