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ON THE CHVÁTAL-JANSON CONJECTURE

LUCIO BARABESI, LUCA PRATELLI, AND PIETRO RIGO

ABSTRACT. Let $q_m = P(X \le m)$, where *m* is a positive integer and *X* a binomial random variable with parameters *n* and m/n. Vašek Chvátal conjectured that, for fixed $n \ge 2$, q_m attains its minimum when *m* is the integer closest to 2n/3. As shown by Svante Janson, this conjecture is true for large *n*. Here, we prove that the conjecture is actually true for every $n \ge 2$.

1. INTRODUCTION

Denoting by B(n, p) a binomial random variable with parameters n and p, Janson [4] investigates the following conjecture suggested by Chvátal in a personal communication.

Conjecture 1 (Chvátal). For any fixed $n \ge 2$, as m ranges over $\{0, \ldots, n\}$,

 $q_m := P(B(n, m/n) \le m)$

is smallest when m = [2n/3] where $[\cdot]$ represents the nearest integer function.

In addition to be intriguing, Conjecture 1 may have useful applications, since the probability that a binomial random variable exceeds its expected value plays a role in the machine learning framework; see e.g. [1], [3], [9] and references therein. Such a probability is also connected to an equation given by Ramanujan, as emphasized by [5]. See also [6] and [8] for further results on this topic.

For large n, Conjecture 1 is actually true and the q_m have a unique minimum.

Theorem 1 (Janson, [4]). There exists an integer n_0 such that, for each $n \ge n_0$: i) q_m is minimum for m = [2n/3] and ii) $q_m > q_{m+1}$ or $q_m < q_{m+1}$ according to whether $m + \frac{1}{2} < 2n/3$ or $m + \frac{1}{2} > 2n/3$.

As noted in [4, Remark 1.5], in principle, the value of n_0 could be computed and then Conjecture 1 could be proved (or disproved) by considering all $n < n_0$. Even if potentially possible, however, this strategy looks not practically feasible and Janson wishes for a general proof of Conjecture 1.

The only purpose of this note is to prove that Conjecture 1 is actually true.

Theorem 2. For each $n \ge 2$, one obtains $q_m > q_{m+1}$ or $q_m < q_{m+1}$ according to whether $m + \frac{1}{2} < 2n/3$ or $m + \frac{1}{2} > 2n/3$. Hence, if $m_0 = [[2n/3]]$, then $q_{m_0} < q_m$ for each $m \ne m_0$.

Key words and phrases. Binomial distribution, Binomial tail probability, Bernoulli inequality.

Our proof of Theorem 2 is quite plain and relies on completely different arguments with respect to [4]. In fact, [4] exploits the version for integer-valued random variables of the asymptotic Edgeworth expansion for probabilities in the central limit theorem - as proposed by Esseen [2]. Instead, our proof is closer to the approach introduced by [7, Appendix B] for showing that $q_m \ge q_{m+1}$ for $0 \le m < n/2$ and $n \ge 2$.

2. Two preliminary lemmas

Let U_1, \ldots, U_n be *n* independent copies of a uniform random variable on [0, 1] and $U_{(1)} \leq \ldots \leq U_{(n)}$ the corresponding order statistics. For m < n, since $U_{(m+1)}$ has a beta distribution with parameters m + 1 and n - m, one obtains

$$q_m = P\left(\sum_{i=1}^n I_{\{U_i \le m/n\}} \le m\right) = P(U_{(m+1)} > m/n)$$
$$= (m+1)\binom{n}{m+1} \int_{m/n}^1 x^m (1-x)^{n-m-1} dx.$$

Lemma 1. Let $n \ge 2$ and $m \le n-2$. Then, $q_m \ge q_{m+1}$ if and only if

$$\int_0^1 (1 - \frac{v}{m+1})^m (1 + \frac{v}{n-m-1})^{n-m-1} \, dv \ge 1.$$
(2)

Proof. First note that $q_m \ge q_{m+1}$ is equivalent to

$$\frac{m+1}{n-m-1}\int_{m/n}^{1} x^m (1-x)^{n-m-1} \, dx \ge \int_{(m+1)/n}^{1} x^{m+1} (1-x)^{n-m-2} \, dx.$$

Integrating the left-hand side by parts, this inequality becomes

$$\int_{m/n}^{(m+1)/n} x^{m+1} (1-x)^{n-m-2} \, dx \ge \frac{(m/n)^{m+1} (1-m/n)^{n-m-1}}{n-m-1}.$$

Letting x = (m + t)/n in the integral, one obtains

$$\int_0^1 (1+t/m)^{m+1} (1-t/(n-m))^{n-m-2} dt \ge \frac{n-m}{n-m-1}.$$

Integrating again the left-hand side by parts, such inequality turns into

$$\int_0^1 (1+\frac{t}{m})^m (1-\frac{t}{n-m})^{n-m-1} dt \ge (1-\frac{1}{n-m})^{n-m-1} (1+\frac{1}{m})^m$$

or equivalently

$$\int_0^1 (\frac{t+m}{1+m})^m (\frac{n-m-t}{n-m-1})^{n-m-1} dt \ge 1.$$

Now, inequality (2) follows from the transformation t = 1 - v.

Lemma 2. Fix $n \ge 3$ and define

$$g_v(x) = \left(1 - \frac{v}{x+1}\right)^x \left(1 + \frac{v}{n-x-1}\right)^{n-x-1}$$

for all $v \in (0,1]$ and $x \in [1, n-2]$. Then, $x \mapsto g_v(x)$ is strictly decreasing for each fixed v. In particular, if

$$h(m) = \int_0^1 g_v(m) \, dv = \int_0^1 \left(1 - \frac{v}{m+1}\right)^m \left(1 + \frac{v}{n-m-1}\right)^{n-m-1} dv$$

for $m \in \{1, ..., n-2\}$, the function h is strictly decreasing.

Proof. Fix $(v, x) \in (0, 1] \times [1, n - 2]$, and note that

$$g'_{v}(x) = g_{v}(x) \left[\log(1 - \frac{v}{x+1}) + \frac{\frac{vx}{(x+1)^{2}}}{1 - \frac{v}{x+1}} - \log(1 + \frac{v}{n-x-1}) + \frac{\frac{v}{n-x-1}}{1 + \frac{v}{n-x-1}} \right].$$

Therefore,

$$g'_v(x) < 0 \quad \Longleftrightarrow \quad \frac{\frac{vx}{(x+1)^2}}{1 - \frac{v}{x+1}} + \frac{\frac{v}{n-x-1}}{1 + \frac{v}{n-x-1}} < \log\left(\frac{1 + \frac{v}{n-x-1}}{1 - \frac{v}{x+1}}\right).$$

In addition,

$$\log\left(\frac{1+\frac{v}{n-x-1}}{1-\frac{v}{x+1}}\right) = \log\left[\left(1+\frac{v}{n-x-1}\right)\left(1+\frac{\frac{v}{x+1}}{1-\frac{v}{x+1}}\right)\right]$$
$$= \log\left(1+\frac{v}{n-x-1}\right) + \log\left(1+\frac{\frac{v}{x+1}}{1-\frac{v}{x+1}}\right).$$

Hence, in order to prove $g'_v(x) < 0$, it suffices to show that

$$\frac{\frac{v}{n-x-1}}{1+\frac{v}{n-x-1}} < \log\left(1+\frac{v}{n-x-1}\right)$$
(3)

and

$$\frac{\frac{vx}{(x+1)^2}}{1-\frac{v}{x+1}} < \log\left(1+\frac{\frac{v}{x+1}}{1-\frac{v}{x+1}}\right).$$
(4)

To prove (3)-(4), first note that $\log(1 + c) > c/(1 + c)$ for each c > 0. Therefore, (3) holds with c = v/(n - x - 1). Similarly, letting c = v/(x + 1 - v), inequality (4) reduces to

$$\log(1+c) - c + \frac{c^2}{v(c+1)} > 0.$$

Finally, the above inequality is true, since

$$\log(1+c) - c + \frac{c^2}{v(c+1)} \ge \log(1+c) - c + \frac{c^2}{c+1} = \log(1+c) - \frac{c}{c+1} > 0.$$

3. A proof of the Chvátal-Janson conjecture

We are now ready to attack Theorem 2. By a direct computation, Theorem 2 holds true for $n \leq 5$. Hence, it can be assumed n = 3s + r where $s \geq 2$ and $r \in \{0, 1, 2\}$. In this case, because of Lemmas 1-2, it suffices to prove that

$$\int_0^1 (1 - \frac{v}{2s})^{2s-1} (1 + \frac{v}{s})^s \, dv > 1 \tag{5}$$

and

$$\int_0^1 (1 - \frac{v}{2s+1})^{2s} (1 + \frac{v}{s-1})^{s-1} \, dv < 1 \tag{6}$$

if r = 0, while

$$\int_{0}^{1} (1 - \frac{v}{2s+1})^{2s} (1 + \frac{v}{s+r-1})^{s+r-1} \, dv > 1 \tag{7}$$

and

$$\int_0^1 (1 - \frac{v}{2s+2})^{2s+1} (1 + \frac{v}{s+r-2})^{s+r-2} \, dv < 1 \tag{8}$$

if $r \in \{1, 2\}$.

We point out that, since all the previous inequalities are strict, one obtains $q_m \neq q_{m+1}$ for all m provided such inequalities are true.

Inequalities (5) and (6). Let r = 0. Recalling the Bernoulli inequality

$$(1+c)^s \ge 1+sc \qquad \text{for all } c > -1,$$

one obtains

$$\begin{split} \int_{0}^{1} (1 - \frac{v}{2s})^{2s-1} (1 + \frac{v}{s})^{s} \, dv &= \int_{0}^{1} [(1 - \frac{v}{2s})^{2} (1 + \frac{v}{s})]^{s} (1 - \frac{v}{2s})^{-1} \, dv \\ &= \int_{0}^{1} [1 - \frac{3v^{2}}{4s^{2}} + \frac{v^{3}}{4s^{3}}]^{s} (1 - \frac{v}{2s})^{-1} \, dv \\ &\geq \int_{0}^{1} (1 - \frac{3v^{2}}{4s} + \frac{v^{3}}{4s^{2}}) (1 - \frac{v}{2s})^{-1} \, dv \\ &> \int_{0}^{1} (1 - \frac{3v^{2}}{4s} + \frac{v^{3}}{4s^{2}}) (1 + \frac{v}{2s} + \frac{v^{2}}{4s^{2}}) \, dv \\ &= 1 + \frac{5}{96s^{2}} - \frac{1}{80s^{3}} + \frac{1}{96s^{4}} > 1. \end{split}$$

Here, the first inequality is because of the Bernoulli's one while the second depends on $(1-c)^{-1} > 1 + c + c^2$ for all $c \in (0, 1)$. Hence, inequality (5) is actually true. Let us turn to inequality (6). We have to show that $I_s < 1$, where

$$I_s = \int_0^1 (1 - \frac{v}{2s+1})^{2s} (1 + \frac{v}{s-1})^{s-1} dv.$$

First note that

$$\begin{split} I_s &= \int_0^1 \frac{2s+1}{2s+1-v} \exp\left((2s+1)\log(1-\frac{v}{2s+1}) + (s-1)\log(1+\frac{v}{s-1})\right) dv \\ &= \int_0^1 \frac{2s+1}{2s+1-v} \exp\left(\sum_{k=2}^\infty \frac{v^k}{k} \left(\frac{(-1)^{k-1}}{(s-1)^{k-1}} - \frac{1}{(2s+1)^{k-1}}\right)\right) dv \\ &< \int_0^1 \frac{2s+1}{2s+1-v} \exp\left(\sum_{k=2}^3 \frac{v^k}{k} \left(\frac{(-1)^{k-1}}{(s-1)^{k-1}} - \frac{1}{(2s+1)^{k-1}}\right)\right) dv \end{split}$$

where the last inequality depends on

$$\sum_{k=4}^{\infty} \frac{v^k}{k} \left(\frac{(-1)^{k-1}}{(s-1)^{k-1}} - \frac{1}{(2s+1)^{k-1}} \right) < 0.$$

Since $\exp(c) < 1 + c + \frac{c^2}{2}$ for c < 0 and

$$\gamma(s,v) := \sum_{k=2}^{3} \frac{v^k}{k} \left(\frac{(-1)^{k-1}}{(s-1)^{k-1}} - \frac{1}{(2s+1)^{k-1}} \right) = \frac{-3v^2s}{2(s-1)(2s+1)} + \frac{v^3s(s+2)}{(s-1)^2(2s+1)^2} < 0,$$

one also obtains

$$I_s < \int_0^1 \frac{2s+1}{2s+1-v} \exp\left(\gamma(s,v)\right) dv < \int_0^1 \frac{2s+1}{2s+1-v} \left(1+\gamma(s,v)+\frac{\gamma(s,v)^2}{2}\right) dv.$$

Moreover, since

$$\begin{aligned} \frac{2s+1}{2s+1-v} &= \frac{1}{1-\frac{v}{2s+1}} = 1 + \frac{v}{2s+1} + \frac{v^2}{(2s+1)^2} \frac{1}{1-\frac{v}{2s+1}} \\ &\leq 1 + \frac{v}{2s+1} + \frac{5v^2}{4(2s+1)^2}, \end{aligned}$$

it follows that

$$I_s < \int_0^1 \left(1 + \frac{v}{2s+1} + \frac{5v^2}{4(2s+1)^2}\right) \left(1 + \gamma(s,v) + \frac{\gamma(s,v)^2}{2}\right) dv.$$

After some (tedious) algebra, the above integral can be evaluated and the previous inequality can be written as

$$I_s < 1 + \frac{1 + 3s(-11s^3 + (s+4)^2)}{(s-1)^4(2s+1)^6} + \frac{s^6(29 + 7s - 15s^2/2)}{(s-1)^4(2s+1)^6}.$$

Both fractions in the previous expression are negative for $s \ge 3$. Hence, $I_s < 1$ for each $s \ge 3$. Finally, $I_2 < 1$ follows from a direct calculation.

This concludes the proof of inequality (6).

Inequalities (7) and (8). Let $r \in \{1, 2\}$ and

$$J_s^{(r)} = \int_0^1 (1 - \frac{v}{2s+1})^{2s} (1 + \frac{v}{s+r-1})^{s+r-1} dv.$$

Since $(1 + \frac{v}{s})^s \leq (1 + \frac{v}{s+1})^{s+1}$, one obtains $J_s^{(1)} \leq J_s^{(2)}$. Hence, to prove (7), it suffices to show $J_s^{(1)} > 1$. To this end, we first write

$$J_s^{(1)} = \int_0^1 (1 - \frac{v}{2s+1})^{2s} (1 + \frac{v}{s})^s \, dv$$

= $\int_0^1 [(1 - \frac{v}{2s+1})^2 (1 + \frac{v}{s})]^s \, dv$
= $\int_0^1 [1 + \frac{v(1-2v)}{s(2s+1)} + \frac{v^2}{(2s+1)^2} (1 + \frac{v}{s})]^s \, dv.$

Hence, the Bernoulli inequality yields

$$\begin{split} J_s^{(1)} &\geq \int_0^1 [1 + \frac{v(1-2v)}{2s+1} + \frac{sv^2}{(2s+1)^2}(1+\frac{v}{s})] \, dv \\ &= 1 - \frac{1}{6(2s+1)} + \frac{s}{3(2s+1)^2} + \frac{1}{4(2s+1)^2} = 1 + \frac{1}{12(2s+1)^2}. \end{split}$$

This proves inequality (7).

Finally, we turn to (8). Let

$$H_s^{(r)} = \int_0^1 \left(1 - \frac{v}{2s+2}\right)^{2s+1} \left(1 + \frac{v}{s+r-2}\right)^{s+r-2} dv.$$

Once again, $H_s^{(1)} \leq H_s^{(2)}$. Thus, to prove (8), it suffices to show that $H_s^{(2)} < 1$. To this end, we argue as in the proof of $I_s < 1$. Precisely, we first note that

$$\begin{split} H_s^{(2)} &= \int_0^1 (1 - \frac{v}{2s+2})^{2s+1} (1 + \frac{v}{s})^s \, dv \\ &= \int_0^1 \frac{2s+2}{2s+2-v} \exp\left((2s+2)\log(1 - \frac{v}{2s+2}) + s\log(1 + \frac{v}{s})\right) dv \\ &= \int_0^1 \frac{2s+2}{2s+2-v} \exp\left(\sum_{k=2}^\infty \frac{v^k}{k} (\frac{(-1)^{k-1}}{s^{k-1}} - \frac{1}{(2s+2)^{k-1}})\right) \, dv \\ &< \int_0^1 \frac{2s+2}{2s+2-v} \exp\left(\sum_{k=2}^3 \frac{v^k}{k} (\frac{(-1)^{k-1}}{s^{k-1}} - \frac{1}{(2s+2)^{k-1}})\right) \, dv \end{split}$$

where the last inequality is because

$$\sum_{k=4}^{\infty} \frac{v^k}{k} \left(\frac{(-1)^{k-1}}{s^{k-1}} - \frac{1}{(2s+2)^{k-1}} \right) < 0.$$

Moreover,

$$\lambda(s,v) := \sum_{k=2}^{3} \frac{v^k}{k} \left(\frac{(-1)^{k-1}}{s^{k-1}} - \frac{1}{(2s+2)^{k-1}} \right) < 0,$$

and
$$\frac{2s+2}{2s+2-v} \le 1 + \frac{v}{2s+2} + \frac{6v^2}{5(2s+2)^2}.$$

Hence, recalling that $\exp(c) < 1 + c + \frac{c^2}{2}$ for c < 0, one obtains

$$\begin{aligned} H_s^{(2)} &< \int_0^1 \frac{2s+2}{2s+2-v} \exp\left(\lambda(s,v)\right) dv \\ &< \int_0^1 \left(1 + \frac{v}{2s+2} + \frac{6v^2}{5(2s+2)^2}\right) \left(1 + \lambda(s,v) + \frac{\lambda(s,v)^2}{2}\right) dv. \end{aligned}$$

Finally, evaluating the integral, the previous inequality turns into

$$H_s^{(2)} < 1 + \frac{-2s^5 - 9(s^4 + s^3) + 8s^2 + 22s + 18}{64s(s+1)^6} < 1.$$

This proves (8) and concludes the proof of Theorem 2.

Added in proof: After writing this paper, we learned (from an anonymous referee) of the existence of another paper very similar to ours, that is: Ping Sun (2021) Strictly unimodality of the probability that the binomial distribution is more than its expectation, *Discrete Applied Mathematics* 301, 1–5. However, we point out that a preliminary draft of our paper appeared on arXiv previous to Sun's paper; see: arXiv:2104.11971v1 [math.PR]

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Lucio Barabesi, Dipartimento di Economia Politica e Statistica, Università di Siena, via San Francesco 7, 53100 Siena, Italy

 $Email \ address: \verb+lucio.barabesi@unisi.it+$

LUCA PRATELLI, ACCADEMIA NAVALE, VIALE ITALIA 72, 57100 LIVORNO, ITALY *Email address:* pratel@mail.dm.unipi.it

Pietro Rigo (corresponding author), Dipartimento di Scienze Statistiche "P. Fortunati", Universita' di Bologna, via delle Belle Arti 41, 40126 Bologna, Italy

 $Email \ address: \verb"pietro.rigo@unibo.it"$