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CONTINUOUS-TIME RANDOM WALK BETWEEN LÉVY-SPACED TARGETS IN THE REAL LINE

ALESSANDRA BIANCHI, MARCO LENCI, AND FRANÇOISE PÈNE

ABSTRACT. We consider a continuous-time random walk which is defined as an interpolation of a random walk on a point process on the real line. The distances between neighboring points of the point process are i.i.d. random variables in the normal domain of attraction of an α -stable distribution with $0 < \alpha < 1$. This is therefore an example of a *random walk in a Lévy random medium*. Specifically, it is a generalization of a process known in the physical literature as *Lévy-Lorentz gas*. We prove that the annealed version of the process is superdiffusive with scaling exponent $1/(\alpha + 1)$ and identify the limiting process, which is not càdlàg. The proofs are based on the technique of Kesten and Spitzer for random walks in random scenery.

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Keywords: Lévy-Lorentz gas, random walk on point process, anomalous diffusion, Lévy random medium, stable processes, random walk in random scenery.

1. INTRODUCTION

In this paper we study a continuous-time random walk in a random medium on the real line. The random medium is given by a marked point process in which the distances between consecutive points are i.i.d. variables taken in the basin of attraction of an α -stable distribution, with $0 < \alpha < 1$. The marked points are referred to as *targets*. A particle travels in \mathbb{R} with unit speed visiting the sequence of targets selected by an independent random walk $(S_n, n \in \mathbb{N})$. This random walk is called the *underlying random walk*.

So, for example, if the realization of (S_n) is $(0, 2, -3, -1, \dots)$, the particle will travel with velocity 1 from the origin O to the second target to the right of O . Then it will instantaneously change its velocity to -1 to travel toward the third target to the left of O . Then it will travel with velocity 1 toward the first target to the left of O , and so on.

Since the distances between targets are fat-tailed variables (their first moment is infinite), the trajectories of the process occasionally experience extremely long inertial segments, leading one to believe that the process is superdiffusive. By that we mean that it scales like a power of time with exponent bigger than $1/2$. The purpose of the paper is to show this in a very specific sense.

More precisely, if $X(t) \equiv X^\omega(t)$ denotes our process, with the label ω representing the random medium, we prove that the finite-dimensional distributions of $(n^{-1/(\alpha+1)}X(nt), t \geq 0)$, w.r.t. both the random medium and the random dynamics,

converges to those of a certain process which we identify (Theorem 2.1). In other words, we prove an annealed generalized CLT for the finite-dimensional distributions of $(X(t), t \geq 0)$.

Processes of this kind have wide application in the physical sciences, where they are used as models for *anomalous diffusion*; see [SZU, KRS, ZDK, CGLS, VBB] and references therein. Indeed, in many real-world applications (in statistical physics, optics, epidemics, etc.), the long inertial segments, or *ballistic flights*, that cause superdiffusion are not due to an anomalous mechanism that governs the dynamics of the agent (a particle, a photon, an animal, etc.) but rather to the complexity of the medium in which the motion occurs. An example is molecular diffusion in porous media [Le] (see the reference lists of [Le, BFK, BCV] for more examples). As a matter of fact, the model presented here is a generalization of the so-called *Lévy-Lorentz gas*, which was introduced in [BFK] precisely as a one-dimensional toy model for transport in porous media. (The Lévy-Lorentz gas is the case where the underlying random walk is a simple symmetric random walk.)

In the language of probability it makes sense to call the process at hand a continuous-time *random walk in a Lévy random medium*, meaning that it is the medium that causes the long ballistic flights. We occasionally also use the expression ‘random walk in a Lévy random environment’ to lay emphasis on the analogies between our process and the random walks in random environment — even though the two types of processes are technically different.

A rigorous study of our system was initiated in [BCLI] where the authors considered the easier case where the distances between the targets have finite mean and infinite variance. This includes all random variables with distribution in the basin of attraction of an α -stable law with $1 < \alpha < 2$. For this regime, subject to certain hypotheses, they prove a quenched normal CLT, that is, for a.e. realization ω of the medium, $t^{-1/2}X^\omega(t)$ converges, as $t \rightarrow \infty$, to a Gaussian variable independent of ω . This implies in particular the annealed normal CLT, confirming and adding on some of the predictions of [BCV] for the Lévy-Lorentz gas.

Our main result in this paper confirms and extends some other predictions of [BCV] and for the first time, to our knowledge, prove superdiffusion for some processes in Lévy random media. The scaling we obtain is also in agreement with the one that has been identified recently for a related model [ACOR]. Our proofs are based on an adaptation of the technique of Kesten and Spitzer for *random walks in random scenery* [KS]. This explains why our limit processes involve the so-called *Kesten-Spitzer process* Δ .

The paper is organized as follows: In Section 2 we give the precise definitions of all the processes associated with our random walk, and present our most important results (Theorems 2.1 and 2.2). In Section 3 we prove these results by means of what we call our Main Lemma (Lemma 3.1). The Main Lemma is then proved in Section 4, using ideas from the theory of random walks in random scenery.

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2. SETUP

Let $\zeta := (\zeta_j, j \in \mathbb{Z})$ be a sequence of i.i.d. positive random variables, the common distribution of which belongs to the normal basin of attraction of an α -stable distribution, with $0 < \alpha < 1$. This means that, as $n \rightarrow \infty$, $n^{-1/\alpha} \sum_{j=1}^n \zeta_j$ converges in distribution to a variable Z_1 , which must then be positive and non-integrable. The recursive sequence of definitions

$$(2.1) \quad \omega_0 := 0, \quad \omega_k - \omega_{k-1} := \zeta_k, \quad \text{for } k \in \mathbb{Z}.$$

determines a marked *point process* $\omega := (\omega_k, k \in \mathbb{Z})$ on \mathbb{R} , which we call the *random medium* or *random environment*, or simply the environment. Each point ω_k will be called a *target*. The set of all possible environments is denoted Ω_{en} , and the law on it P .

Let $S := (S_n, n \in \mathbb{N})$ be a \mathbb{Z} -valued random walk starting from $S_0 := 0$ with centered i.i.d. increments $\xi_j := S_j - S_{j-1}$, for $j \in \mathbb{Z}^+$. We assume that there exists $\gamma > 2/\alpha$ such that the absolute moment of ξ_j of order γ is finite (this assumption is used in Lemma 4.3). This implies that the variance of ξ_j , which we denote v_ξ , is also finite. In order for the problem to make sense, we also assume $v_\xi > 0$. We refer to S as the *underlying random walk* and call Q its distribution (on $\mathbb{Z}^{\mathbb{N}}$ endowed with the σ -algebra generated by cylinders).

Given $\omega \in \Omega_{\text{en}}$, we define the so-called *random walk on the point process* to be, for all $n \in \mathbb{N}$,

$$(2.2) \quad Y_n \equiv Y_n^\omega := \omega_{S_n}.$$

In simple terms, $Y := (Y_n, n \in \mathbb{N})$ performs the same jumps as S , but on the points of ω .

The process we are interested in here is the continuous-time random walk $X \equiv X^\omega := (X^\omega(t), t \geq 0)$ whose trajectories interpolate those of the walk Y and have unit speed (save at collision times). Formally it can be defined as follows: Given a realization ω of the medium and a realization S of the dynamics, we define the sequence of *collision times* $T_n \equiv T_n(\omega, S)$ via

$$(2.3) \quad T_0 := 0, \quad T_n := \sum_{k=1}^n |\omega_{S_k} - \omega_{S_{k-1}}|, \quad \text{for } n \geq 1.$$

Since the length of the n^{th} jump of the walk is given by $|\omega_{S_n} - \omega_{S_{n-1}}|$, T_n represents the global length of the trajectory up to the n^{th} collision.

Finally, $X(t) \equiv X^\omega(t)$ is defined by the equations

$$(2.4) \quad X(t) := Y_n + \text{sgn}(\xi_{n+1})(t - T_n), \quad \text{for } t \in [T_n, T_{n+1}).$$

Notice that $\xi_{n+1} = 0 \Leftrightarrow S_{n+1} = S_n \Leftrightarrow T_{n+1} = T_n$. Therefore the definition (2.4) is never used in the case $\xi_{n+1} = 0$, making the value of $\text{sgn}(0)$ irrelevant there. More importantly, the self-jumps of the underlying RW, i.e., $S_{n+1} = S_n$, are not seen by the continuous-time process X . This implies that one can remove any lazy

component of S by replacing the common distribution of the ξ_j by the conditional distribution of ξ_1 given $\xi_1 \neq 0$ (this is well defined since $v_\xi > 0$). We thus assume from now on that $Q(\xi_j = 0) = 0$.

Also without loss of generality we assume that ξ_1 , equivalently S , is not supported on a proper subgroup of \mathbb{Z} . If it were supported, say, on $d\mathbb{Z}$, with $d \in \mathbb{N}$, $d \geq 2$, one could divide S by d and replace ζ_j with $\sum_{k=(j-1)d+1}^{jd} \zeta_k$. The resulting continuous-time walk would be a simple rescaling of the original process X .

The law that embraces all the random aspects of the process is the product probability $\mathbb{P} := P \times Q$ on $\Omega_{\text{en}} \times \mathbb{Z}^{\mathbb{N}}$. The probability \mathbb{P} is sometimes called the *annealed* or *averaged* law.

Our main result is the finite-dimensional generalized CLT for the family of processes

$$(2.5) \quad (\bar{X}^{(q)}(s), s \geq 0) := \left(\frac{X(sq)}{q^{1/(\alpha+1)}}, s \geq 0 \right),$$

where $q \in \mathbb{R}^+$ plays the role of the *scaling parameter*. In order to state it, we need some preliminary notation. Let $(B(t), t \geq 0)$ be a Brownian motion such that $B(t)$ has mean 0 and variance $v_\xi t$, and denote by $(L_t(x), x \in \mathbb{R})$ its corresponding local time. Let $(Z_\pm(x), x \geq 0)$ be two i.i.d. càdlàg α -stable processes with independent and stationary increments such that $Z_\pm(0) \equiv 0$ and $Z_\pm(1)$ is distributed like Z_1 , cf. beginning of the section. Notice that the processes Z_\pm are strictly increasing subordinators.

Assume that the processes B , Z_+ and Z_- are mutually independent and consider the process $\Delta := (\Delta(t), t \geq 0)$ given by

$$(2.6) \quad \Delta(t) := \mu_\xi \left(\int_0^{+\infty} L_t(x) dZ_+(x) + \int_0^{+\infty} L_t(-x) dZ_-(x) \right),$$

with $\mu_\xi := \mathbb{E}(|\xi_1|)$. The above is well-defined because $x \mapsto L_t(x)$ is almost surely continuous and compactly supported. Also, let us define $(Z(x), x \in \mathbb{R})$ to be

$$(2.7) \quad Z(x) := \begin{cases} Z_+(x), & x \geq 0; \\ -Z_-(-x), & x < 0. \end{cases}$$

Observe moreover that Δ is a.s. continuous and strictly increasing.

Theorem 2.1. *Under the law \mathbb{P} , the finite dimensional distributions of $(\bar{X}^{(q)}(s), s \geq 0)$ converge, as $q \rightarrow \infty$, to those of $(Z \circ B \circ \Delta^{-1}(s), s \geq 0)$.*

The process Δ , often called the Kesten-Spitzer process, was introduced in [KS] as the limit, for $n \rightarrow \infty$, of the processes

$$(2.8) \quad \left(\frac{1}{n^{(1+\alpha)/2\alpha}} \sum_{k=1}^{\lfloor nt \rfloor} \zeta_{S_k}, t \in \mathbb{R}^+ \right),$$

where $\lfloor a \rfloor$ denotes the largest integer less than or equal to $a \in \mathbb{R}$. The sum $\sum_k \zeta_{S_k}$ is referred to as a *random walk in the random scenery* ζ ; see also [Bo1, Bo2]. In the present context Δ appears as the limit of the suitably rescaled collision time processes, as clarified by the next result, which also gives the asymptotic behavior of Y ,

the random walk on the point process. (The recent reference [MS] also investigates limit theorems for Y in the case $\alpha \in (1, 2)$, and with some restrictions. This process is of course simpler than the actual Lévy-Lorentz gas.)

Theorem 2.2. *Under \mathbb{P} , the (joint) finite dimensional distributions of*

$$\left(\left(\frac{T_{\lfloor tq \rfloor}}{q^{(1+\alpha)/2\alpha}}, \frac{Y_{\lfloor tq \rfloor}}{q^{1/2\alpha}} \right), t \geq 0 \right)$$

converge, for $q \rightarrow \infty$, to the corresponding distributions of $((\Delta, Z \circ B), t \geq 0)$.

As a last remark for this section, observe that neither Theorem 2.1 nor Theorem 2.2 can be extended to a functional limit theorem w.r.t. *any* Skorokhod topology, since the limit processes do not belong to the Skorokhod space. Indeed $Z \circ B \circ \Delta^{-1}$ and $Z \circ B$ have discontinuities without one-sided limits. This can be seen as follows: If x_0 is a jump discontinuity of Z , which exists almost surely, and $B(t_0) = x_0$, then almost surely $B(t)$ oscillates around x_0 , both as $t \rightarrow t_0^+$ and as $t \rightarrow t_0^-$. Therefore neither limit

$$(2.9) \quad \lim_{t \rightarrow t_0^\pm} Z \circ B(t)$$

exists. The situation is analogous for $Z \circ B \circ \Delta^{-1}$.

Heuristically, the cause of this phenomenon is that the gaps between targets are non-integrable i.i.d. variables. Therefore, from time to time, the walker will encounter a gap of the same order as the sum of all the gaps visited till then. The variation in position achieved upon traveling that gap will thus be of finite order in the limit. Since, as $q \rightarrow \infty$, time shrinks by a factor q^{-1} and space by a factor $q^{-1/(\alpha+1)} \ll q^{-1}$, the trajectory segments corresponding to these exceptionally long gaps become jump discontinuities. On the other hand, some of these gaps will be traveled on, back and forth, a larger and larger number of times, creating in the limit an accumulation of jump discontinuities of the same size.

3. RESCALED PROCESSES

At the core of the proofs lies a functional CLT for the joint distribution of all the processes involved here: the underlying random walk, the point process and the collision time process. In this Section we state this limit theorem, later referred to as the Main Lemma, and use it to derive Theorems 2.1 and 2.2.

Fix a value $q \in \mathbb{R}^+$ of the scaling parameter and let the processes $(\bar{\omega}^{(q)}(x), x \in \mathbb{R})$, $(\bar{S}^{(q)}(t), t \geq 0)$, $(\bar{T}^{(q)}(t), t \geq 0)$ be defined by

$$(3.1) \quad \bar{\omega}^{(q)}(x) := \frac{\omega_{\lfloor x\sqrt{q} \rfloor}}{q^{1/2\alpha}};$$

$$(3.2) \quad \bar{S}^{(q)}(t) := \frac{S_{\lfloor tq \rfloor}}{\sqrt{q}};$$

$$(3.3) \quad \bar{T}^{(q)}(t) := \frac{T_{\lfloor tq \rfloor}}{q^{(\alpha+1)/2\alpha}}.$$

Convention. Throughout the paper, the notation $[a, b]$ denotes the closed interval between the real numbers a and b , irrespective of their order.

This is our Main Lemma:

Lemma 3.1. *Under \mathbb{P} , as $q \rightarrow \infty$, the joint process*

$$\left((\bar{\omega}^{(q)}(x), x \in \mathbb{R}), (\bar{S}^{(q)}(t), t \geq 0), (\bar{T}^{(q)}(t'), t' \geq 0) \right)$$

converges to

$$\left((Z(x), x \in \mathbb{R}), (B(t), t \geq 0), (\Delta(t'), t' \geq 0) \right)$$

w.r.t. the (product) J_1 -metric on $\mathbb{D}(\mathbb{R}) \times \mathbb{D}(\mathbb{R}_0^+) \times \mathbb{D}(\mathbb{R}_0^+)$. Here $\mathbb{D}(\mathbb{R})$ is the space of the trajectories $\bar{\omega}(x) : \mathbb{R} \rightarrow \mathbb{R}$ such that $(\bar{\omega}(x), x \geq 0)$ and $(-\bar{\omega}(-x), x \geq 0)$ are càdlàg; the space is endowed with the topology induced by the J_1 -metric on every interval $[0, y]$, $y \in \mathbb{R}$ (see [B], Secs. 12 and 16).

3.1. Proof of Theorems 2.1 and 2.2. If $f : \mathbb{R} \rightarrow \mathbb{R}$ is a (not necessarily strictly) increasing function, let $f^{-1}(s) := \sup\{u > 0 : f(u) < s\}$ denote our choice for the generalized inverse of f .

The function

$$(3.4) \quad t \mapsto \bar{T}^{(q^{2\alpha/(\alpha+1)})}(t) = \frac{T_{\lfloor tq^{2\alpha/(\alpha+1)} \rfloor}}{q},$$

cf. (3.3), is an increasing càdlàg function which is constant on any interval of the form $[q^{-2\alpha/(\alpha+1)}n, q^{-2\alpha/(\alpha+1)}(n+1))$, $n \in \mathbb{N}$. Therefore, given $s \in \mathbb{R}^+$,

$$(3.5) \quad \left(\bar{T}^{(q^{2\alpha/(\alpha+1)})} \right)^{-1}(s) = q^{-2\alpha/(\alpha+1)} \bar{n},$$

for some $\bar{n} \equiv \bar{n}(s, q) \in \mathbb{N}$. The previous two equations imply that $T_{\bar{n}-1} < sq \leq T_{\bar{n}}$, whence, in view of (2.2) and (2.5),

$$(3.6) \quad \bar{X}^{(q)}(s) \in \left[\frac{\omega_{S_{\bar{n}-1}}}{q^{1/(\alpha+1)}}, \frac{\omega_{S_{\bar{n}}}}{q^{1/(\alpha+1)}} \right].$$

Denoting

$$(3.7) \quad p \equiv p(q) := q^{2\alpha/(\alpha+1)}$$

and using (3.5) and (3.2), we see that

$$(3.8) \quad S_{\bar{n}} = \sqrt{p} \bar{S}^{(p)} \left((\bar{T}^{(p)})^{-1}(s) \right),$$

whence, in view of (3.1),

$$(3.9) \quad \omega_{S_{\bar{n}}} = p^{1/2\alpha} \bar{\omega}^{(p)} \left(\bar{S}^{(p)} \left((\bar{T}^{(p)})^{-1}(s) \right) \right).$$

Since $p^{1/2\alpha} = q^{1/(\alpha+1)}$, (3.6) becomes

$$(3.10) \quad \bar{X}^{(q)}(s) \in \left[\bar{\omega}^{(p)} \circ \bar{S}^{(p)} \left((\bar{T}^{(p)})^{-1}(s) - p^{-1} \right), \bar{\omega}^{(p)} \circ \bar{S}^{(p)} \circ (\bar{T}^{(p)})^{-1}(s) \right].$$

Let $(q_k)_{k \in \mathbb{Z}^+}$ be a positive sequence of real numbers such that $q_k \rightarrow +\infty$ as $k \rightarrow \infty$, and let $p_k := p(q_k)$ as defined in (3.7) above.

Assumption. From now on we assume that the convergence in the statement of Lemma 3.1, restricted to $q = p_k$, $k \rightarrow \infty$, holds almost everywhere. If this is not the case, by virtue of the Skorokhod representation theorem [D, Thm. 3], there exists a probability space where it does (for processes which have the same joint distribution as the ones used in Lemma 3.1): as we shall see, the specifics of the probability space are irrelevant. In particular, up to discarding a null set of realizations, we may assume that $B : [0, +\infty) \rightarrow \mathbb{R}$ is continuous and $\Delta : [0, +\infty) \rightarrow [0, +\infty)$ is continuous, strictly increasing and bijective.

Lemma 3.2. *(Almost) surely, for all $s > 0$,*

$$\lim_{k \rightarrow \infty} (\bar{T}^{(p_k)})^{-1}(s) = \Delta^{-1}(s).$$

Proof. Since Δ is continuous and strictly increasing, so is the inverse Δ^{-1} . For a given $s > 0$ and for every sufficiently small $\theta > 0$, we set $u := \Delta^{-1}(s)$, $x_\theta := \Delta^{-1}(s - \theta/2)$ and $y_\theta := \Delta^{-1}(s + \theta/2)$, so that

$$(3.11) \quad \Delta(x_\theta) = s - \frac{\theta}{2} < s = \Delta(u) < \Delta(y_\theta) = s + \frac{\theta}{2}.$$

Then, for any given $\varepsilon > 0$, we choose $\theta > 0$ such that $u - x_\theta < \varepsilon/2$ and $y_\theta - u < \varepsilon/2$. Moreover, we set $\varsigma := \min\{\theta/2, u - x_\theta, y_\theta - u, \varepsilon/2\}$ and notice that for k big enough, we have $d_{J_1, [0, u+1]}(\bar{T}^{(p_k)}, \Delta) < \varsigma$, where the l.h.s. of the inequality denotes the Skorokhod J_1 -distance on the interval $[0, u+1]$. In other words, there exists a strictly increasing homeomorphism $\varphi_k : [0, u+1] \rightarrow [0, u+1]$ such that, for all $v \in [0, u+1]$,

$$(3.12) \quad |v - \varphi_k(v)| < \varsigma;$$

$$(3.13) \quad |\Delta(v) - \bar{T}^{(p_k)}(\varphi_k(v))| < \varsigma.$$

Hence, for k big enough, one has

$$(3.14) \quad \bar{T}^{(p_k)}(\varphi_k(y_\theta)) > \Delta(y_\theta) - \varsigma = s + \frac{\theta}{2} - \varsigma \geq s;$$

$$(3.15) \quad \bar{T}^{(p_k)}(\varphi_k(x_\theta)) < \Delta(x_\theta) + \varsigma = s - \frac{\theta}{2} + \varsigma \leq s,$$

whence

$$(3.16) \quad (\bar{T}^{(p_k)})^{-1}(s) \geq \varphi_k(x_\theta) > x_\theta - \varsigma > u - \frac{\varepsilon}{2} - \varsigma \geq u - \varepsilon;$$

$$(3.17) \quad (\bar{T}^{(p_k)})^{-1}(s) \leq \varphi_k(y_\theta) < y_\theta + \varsigma < u + \frac{\varepsilon}{2} + \varsigma \leq u + \varepsilon.$$

Since ε was arbitrary, this proves the stated convergence. \square

Observe now that Z and B are independent and that the discontinuities of Z occur at random points whose distribution has no atom in \mathbb{R} . Therefore, for every $t \geq 0$, Z is continuous at $B(t)$, equivalently, Z does not have a jump at $B(t)$, with probability 1. This ensures that the following lemma holds true for almost every realization of (B, Z) .

Lemma 3.3. *Let $t > 0$ and consider a realization of our processes such that Z is continuous at $B(t)$. For any sequence $(a_k) \subset \mathbb{R}^+$ with $\lim_{k \rightarrow \infty} a_k = t$, we have:*

$$\lim_{k \rightarrow \infty} \bar{\omega}^{(p_k)}(\bar{S}^{(p_k)}(a_k)) = Z(B(t)).$$

Proof. Let $\varepsilon \in (0, 1)$ and $\eta \in (0, \varepsilon)$ be such that

$$(3.18) \quad \sup_{x: |x-B(t)| < 2\eta} |Z(B(t)) - Z(x)| < \frac{\varepsilon}{2}.$$

Also choose $\varsigma \in (0, \eta/2)$ so that

$$(3.19) \quad \sup_{v: |v-t| < \varsigma} |B(t) - B(v)| < \frac{\eta}{2}.$$

Let k be big enough so that $|a_k - t| < \varsigma/2$ and $d_{J_1, [0, t+1]}(\bar{S}^{(p_k)}, B) < \varsigma/2$. Thus there exists an increasing homeomorphism φ_k of $[0, \Delta^{-1}(s) + 1]$ such that, for all $v \in [0, t + 1]$,

$$(3.20) \quad |v - \varphi_k(v)| < \frac{\varsigma}{2};$$

$$(3.21) \quad |\bar{S}^{(p_k)}(v) - B(\varphi_k(v))| < \frac{\varsigma}{2}.$$

Hence

$$(3.22) \quad |\bar{S}^{(p_k)}(a_k) - B(t)| \leq \frac{\varsigma}{2} + |B(\varphi_k(a_k)) - B(t)| < \frac{\varsigma}{2} + \frac{\eta}{2} < \eta,$$

since $|\varphi_k(a_k) - t| \leq |\varphi_k(a_k) - a_k| + |a_k - t| < \varsigma$.

Assume moreover that k is so big that

$$(3.23) \quad d_{J_1, [0, |B(t)|+1]}(\bar{\omega}^{(p_k)}, Z) < \frac{\eta}{2};$$

$$(3.24) \quad d_{J_1, [0, |B(t)|+1]}(\bar{\omega}^{(p_k)}(-\cdot), Z(-\cdot)) < \frac{\eta}{2}.$$

Then there exists a strictly increasing homeomorphism ψ_k of $[-|B(t)| - 1, |B(t)| + 1]$, with $\psi_k(0) = 0$, such that, for all $w \in [-|B(t)| - 1, |B(t)| + 1]$,

$$(3.25) \quad |w - \psi_k(w)| < \frac{\eta}{2};$$

$$(3.26) \quad |\bar{\omega}^{(p_k)}(w) - Z(\psi_k(w))| < \frac{\eta}{2}.$$

Therefore $|\bar{\omega}^{(p_k)}(\bar{S}^{(p_k)}(a_k)) - Z(\psi_k(\bar{S}^{(p_k)}(a_k)))| < \eta/2$ and

$$(3.27) \quad \begin{aligned} & |\psi_k(\bar{S}^{(p_k)}(a_k)) - B(t)| \\ & \leq |\psi_k(\bar{S}^{(p_k)}(a_k)) - \bar{S}^{(p_k)}(a_k)| + |\bar{S}^{(p_k)}(a_k) - B(t)| \\ & < 2\eta. \end{aligned}$$

whence, by (3.18),

$$\begin{aligned}
(3.28) \quad & |\bar{\omega}^{(p_k)}(\bar{S}^{(p_k)}(a_k)) - Z(B(t))| \\
& \leq \frac{\eta}{2} + |Z(\psi_k(\bar{S}^{(p_k)}(a_k))) - Z(B(t))| \\
& < \frac{\eta}{2} + \frac{\varepsilon}{2} < \varepsilon.
\end{aligned}$$

We conclude that $\lim_{k \rightarrow \infty} \bar{\omega}^{(p_k)}(\bar{S}^{(p_k)}(a_k)) = Z(B(t))$. \square

We are now ready to complete the proofs of Theorems 2.1 and 2.2.

Proof of Theorem 2.1. Let $m \in \mathbb{Z}^+$ and $s_1, \dots, s_m \in [0, +\infty)$ and notice that with probability one, Z is continuous at $B(\Delta^{-1}(s_i))$, for every $i \in \{1, \dots, m\}$.

From the convergence stated in Lemma 3.2, for every $i \in \{1, \dots, m\}$ we can apply Lemma 3.3 twice: first with $a_k = (\bar{T}^{(p_k)})^{-1}(s_i)$ and $t = \Delta^{-1}(s_i)$, then with $a_k = (\bar{T}^{(p_k)})^{-1}(s_i) - q_k^{-2\alpha/(\alpha+1)}$ and again $t = \Delta^{-1}(s_i)$. We then obtain, for all $i = 1, \dots, m$,

$$\begin{aligned}
(3.29) \quad \lim_{k \rightarrow \infty} \bar{\omega}^{(p_k)}\left(\bar{S}^{(p_k)}\left((\bar{T}^{(p_k)})^{-1}(s_i)\right)\right) &= \lim_{k \rightarrow \infty} \bar{\omega}^{(p_k)}\left(\bar{S}^{(p_k)}\left((\bar{T}^{(p_k)})^{-1}(s_i) - p_k^{-1}\right)\right) \\
&= Z(B(\Delta^{-1}(s_i))).
\end{aligned}$$

Hence, from (3.10), it follows that, almost surely,

$$(3.30) \quad \lim_{k \rightarrow \infty} (\bar{X}^{(q_k)}(s_1), \dots, \bar{X}^{(q_k)}(s_m)) = (Z(B(\Delta^{-1}(s_1))), \dots, Z(B(\Delta^{-1}(s_m))))).$$

The convergence then holds in distribution, for any choice of (q_k) . This concludes the proof of Theorem 2.1. \square

Proof of Theorem 2.2. Let $m \in \mathbb{Z}^+$ and $t_1, \dots, t_m \in [0, +\infty)$. Fix an arbitrary sequence $q_k \rightarrow +\infty$. It suffices to prove the assertion of the theorem for $q = q_k$.

For every $i \in \{1, \dots, m\}$ we get by definition (3.3) that

$$(3.31) \quad \frac{T_{[t_i q]}}{q^{(\alpha+1)/2\alpha}} = \bar{T}^{(q)}(t_i).$$

Also, combining definitions (3.1)-(3.2) as in (3.9) gives

$$(3.32) \quad \frac{Y_{[t_i q]}}{q^{1/2\alpha}} = \bar{\omega}^{(q)}(\bar{S}^{(q)}(t_i)).$$

Now, with probability one Z is continuous at $B(t_1), \dots, B(t_m)$. We restrict to such realizations. Fix a sequence $q_k \rightarrow +\infty$. We use (3.31) with the almost sure version of the Main Lemma 3.1, for $q = q_k$, and (3.32) with Lemma 3.3, for $p_k = q_k$ and $a_k = t = t_i$. This proves the a.s. convergence of the l.h.sides of (3.31)-(3.32), along the sequence (q_k) , to $\Delta(t_i)$ and $Z(B(t_i))$, respectively. Considering the joint convergence for all $i \in \{1, \dots, m\}$ and passing to distributional convergence proves the desired assertion. \square

4. PROOF OF THE MAIN LEMMA

The proof of Lemma [3.1](#) follows a standard argument for the convergence of processes and will be split into two parts: convergence of the finite-dimensional distributions and tightness. Beyond technicalities, the main ingredient of the proof is a representation of the collision time as a suitable random walk in random scenery. Before getting to the proofs proper, it is worth clarifying this idea, introducing certain quantities and giving some estimates that will be used further on.

4.1. Collision time and random scenery. Let us first rewrite T_n as

$$(4.1) \quad T_n := \sum_{y \in \mathbb{Z}} \mathcal{N}_n(y) \zeta_y,$$

where $\mathcal{N}_n(y)$ denotes the *local time on the bonds* for the underlying random walk:

$$(4.2) \quad \mathcal{N}_n(y) := \#\{k \in \{1, \dots, n\} : [y-1, y] \subseteq [S_{k-1}, S_k]\}.$$

In other words, $\mathcal{N}_n(y)$ is the number of times, up to the n^{th} jump, that the walk X travels the y^{th} gap $[\omega_{y-1}, \omega_y]$. Notice that, given S , T_n is a sum of independent random variables. In particular, eq. [\(4.1\)](#) provides an interpretation of the collision time as a random walk in random scenery, and suggests the convergence stated in the Main Lemma. To formalize this connection it is useful to introduce more quantities pertaining to the underlying random walk S , such as the standard local time, or *local time on the sites*:

$$(4.3) \quad N_n(y) := \#\{k \in \{0, \dots, n-1\} : S_k = y\};$$

the *range*:

$$(4.4) \quad R_n := \#\{y \in \mathbb{Z} : N_n(y) > 0\};$$

and the *self-intersection*:

$$(4.5) \quad V_n := \sum_{k, \ell=0}^{n-1} \mathbf{1}_{\{S_k = S_\ell\}} = \sum_{y \in \mathbb{Z}} \sum_{k, \ell=0}^{n-1} \mathbf{1}_{\{S_k = S_\ell = y\}} = \sum_{y \in \mathbb{Z}} (N_n(y))^2.$$

We recall (see, e.g., [\[DE\]](#)) that

$$(4.6) \quad \mathbb{E}[R_n] = O(\sqrt{n}).$$

Moreover, as noticed in [\[KS\]](#), the local limit theorem for S implies that

$$(4.7) \quad \mathbb{E}[V_n] = \sum_{k, \ell=0}^{n-1} \mathbb{P}(S_k = S_\ell) = n + 2 \sum_{0 \leq k < \ell \leq n-1} \mathbb{P}(S_{\ell-k} = 0) = O(n^{3/2}).$$

Therefore, by Hölder's inequality, for every $\beta \in (0, 2]$ we have

$$(4.8) \quad \begin{aligned} \sum_{y \in \mathbb{Z}} \mathbb{E}[(N_n(y))^\beta] &\leq \left(\mathbb{E} \left[\sum_{y \in \mathbb{Z}} \mathbf{1}_{\{N_n(y) > 0\}} \right] \right)^{1-\beta/2} \left(\mathbb{E} \left[\sum_{y \in \mathbb{Z}} (N_n(y))^2 \right] \right)^{\beta/2} \\ &= (\mathbb{E}[R_n])^{1-\beta/2} (\mathbb{E}[V_n])^{\beta/2} = O(n^{(1+\beta)/2}). \end{aligned}$$

For later purposes we study certain properties of the local time N_n . To state the first result, we need to introduce $\tau_n(x)$, the time of the n^{th} visit to $x \in \mathbb{Z}$. We define $\tau_0(x) := 0$ and, for $n \geq 1$,

$$(4.9) \quad \tau_n(x) := \inf\{m > \tau_{n-1}(x) : S_m = x\}.$$

Then we have the following:

Lemma 4.1. *For every $\beta \in (0, 1]$, there exists $K_0 > 0$ such that, for every $x, y \in \mathbb{Z}$,*

$$\begin{aligned} \mathbb{E}\left[|N_n(x) - N_n(y)|^\beta\right] &\leq K_0 \left(|x - y|^{\beta/2} \mathbb{E}\left[(N_n(x))^{\beta/2} + (N_n(y))^{\beta/2}\right] \right. \\ &\quad \left. + \left(1 + (|x - y| \log n)^\beta\right) (\mathbb{P}(\tau_1(x) \leq n) + \mathbb{P}(\tau_1(y) \leq n))\right). \end{aligned}$$

Proof. As in [KS], for $x, y \in \mathbb{Z}$, we let $M_j(x, y)$ denote the number of visits of the random walk S to y between the j^{th} and the $(j + 1)^{\text{th}}$ visit to x :

$$(4.10) \quad M_j(x, y) := \sum_{m=\tau_j(x)+1}^{\tau_{j+1}(x)} \mathbf{1}_{\{S_m=y\}}.$$

Note that $M_j(x, y)$ is independent of the σ -algebra $\mathcal{F}_{\tau_j(x)}$, where $\mathcal{F}_n := \sigma(S_1, \dots, S_n)$. For reasons of symmetry, it is enough to study

$$(4.11) \quad \mathbb{E}\left[|N_n(x) - N_n(y)|^\beta \mathbf{1}_{\{\tau_1(x) \leq \tau_1(y)\}}\right].$$

Let us observe that

$$(4.12) \quad \begin{aligned} &(N_n(x) - N_n(y)) \mathbf{1}_{\{\tau_1(x) \leq \tau_1(y)\}} \\ &= \left(\mathbf{1}_{\{\tau_1(x) \leq n\}} + \sum_{j=1}^{N_n(x)-1} (1 - M_j(x, y)) - \sum_{k=\tau_{N_n(x)}(x)}^n \mathbf{1}_{\{S_k=y\}} \right) \mathbf{1}_{\{\tau_1(x) \leq \tau_1(y)\}}. \end{aligned}$$

Using that, for all $a, b \in \mathbb{C}$ and $\beta \in [0, 1]$,

$$(4.13) \quad |a + b|^\beta \leq |a|^\beta + |b|^\beta,$$

we obtain

$$(4.14) \quad \mathbb{E}\left[(N_n(x) - N_n(y))^\beta \mathbf{1}_{\{\tau_1(x) \leq \tau_1(y)\}}\right] \leq \mathbb{P}(\tau_1(x) \leq n) + \mathcal{A}_n(x, y) + \mathcal{B}_n(x, y),$$

where

$$(4.15) \quad \mathcal{A}_n(x, y) := \mathbb{E}\left[\left(\sum_{j=1}^{N_n(x)-1} (1 - M_j(x, y))\right)^\beta \mathbf{1}_{\{\tau_1(x) \leq \tau_1(y)\}}\right];$$

$$(4.16) \quad \mathcal{B}_n(x, y) := \mathbb{E}\left[\left(\sum_{k=\tau_{N_n(x)}(x)}^n \mathbf{1}_{\{S_k=y\}}\right)^\beta \mathbf{1}_{\{\tau_1(x) \leq \tau_1(y) \leq n\}}\right].$$

By the independence of $(M_j(x, y), j \geq 1)$ and $\tau_1(x)$ we get

$$(4.17) \quad \begin{aligned} \mathcal{B}_n(x, y) &\leq \mathcal{B}'_n(x, y) := \mathbb{E} \left[\max_{1 \leq j \leq n} |M_j(x, y)|^\beta \mathbf{1}_{\{\tau_1(x) \leq n\}} \right] \\ &\leq \mathbb{P}(\tau_1(x) \leq n) \mathbb{E} \left[\max_{1 \leq j \leq n} |M_j(x, y)|^\beta \right]. \end{aligned}$$

Moreover, by [KS, proof of Lem. 2], it turns out that the random variables $M_j(x, y)$ are mutually independent, with distribution given by the product of a geometric random variable and a Bernoulli random variable, independent and both with parameter $p(|x-y|)$, where $p : \mathbb{N} \rightarrow [0, 1]$ is a function such that $p(r) \sim cr^{-1}$, for some $c > 0$, as $r \rightarrow +\infty$. In particular, the variables $M_j(x, y)$ are stochastically dominated by independent geometric variables $\widetilde{M}_j(x, y)$ with parameter $p(|x-y|)$, that in turn are dominated by independent exponential variables $M'_j(x, y)$ with parameter $\lambda := \log(1 - p(|x-y|))$. Therefore, by Jensen's inequality and since $\beta \leq 1$,

$$(4.18) \quad \begin{aligned} \mathbb{E} \left[\max_{1 \leq j \leq n} |M_j(x, y)|^\beta \right] &\leq 1 + \mathbb{E} \left[\max_{1 \leq j \leq n} |M'_j(x, y)|^\beta \right] \\ &\leq 1 + \mathbb{E} \left[\max_{1 \leq j \leq n} M'_j(x, y) \right]^\beta. \end{aligned}$$

But

$$(4.19) \quad \begin{aligned} \mathbb{E} \left[\max_{1 \leq j \leq n} M'_j(x, y) \right] &= \int_0^{+\infty} \mathbb{P} \left(\max_{1 \leq j \leq n} M'_j(x, y) \geq t \right) dt \\ &= \int_0^{+\infty} \left(1 - \mathbb{P} \left(\max_{1 \leq j \leq n} M'_j(x, y) < t \right) \right) dt \\ &= \int_0^{+\infty} (1 - (1 - e^{-\lambda t})^n) dt \\ &= \int_0^{+\infty} e^{-\lambda t} \sum_{k=0}^{n-1} (1 - e^{-\lambda t})^k dt \\ &= \sum_{k=0}^{n-1} \left[\frac{(1 - e^{-\lambda t})^{k+1}}{\lambda(k+1)} \right]_0^{+\infty} \\ &= \sum_{k=1}^n \frac{1}{k\lambda} \leq \frac{\log n + 1}{\lambda}. \end{aligned}$$

Combining (4.17), (4.18) and (4.19) and using the fact that, when $|x-y| \rightarrow +\infty$, $\lambda = \log(1 - p(|x-y|)) \sim p(|x-y|) \sim c|x-y|^{-1}$, we obtain the existence of $K_1 > 0$ such that

$$(4.20) \quad \mathcal{B}_n(x, y) \leq \mathcal{B}'_n(x, y) \leq K_1 \mathbb{P}(\tau_1(x) \leq n) (\log n |x-y|)^\beta.$$

Concerning the term $\mathcal{A}_n(x, y)$ in (4.15), observe that

$$(4.21) \quad \sum_{j=1}^{N_n(x)-1} (1 - M_j(x, y)) \mathbf{1}_{\{\tau_1(x) \leq \tau_1(y)\}} = \sum_{j=1}^n (1 - M_j(x, y)) \mathbf{1}_{\{\tau_j(x) \leq n\}} \mathbf{1}_{\{\tau_1(x) \leq \tau_1(y)\}},$$

and recall that, by [KS, Lem. 2], the process $(\mathcal{M}_m := \sum_{j=1}^m h_j, m \geq 1)$, with

$$(4.22) \quad h_j := (1 - M_j(x, y)) \mathbf{1}_{\{\tau_j(x) \leq n\}} \mathbf{1}_{\{\tau_1(x) \leq \tau_1(y)\}},$$

is a martingale with respect to $(\mathcal{F}_{\tau_{m+1}(x)}, m \geq 1)$, thus implying that $\mathbb{E}[h_j^2 | \mathcal{F}_{\tau_j(x)}] = \mathbb{E}[M_j(x, y)^2 | \mathcal{F}_{\tau_j(x)}]$. Therefore, using Burkholder's inequality [HH, Thm. 2.11], there exists $C_\beta > 0$ such that

$$(4.23) \quad \begin{aligned} \mathbb{E}[|\mathcal{M}_n|^\beta] &\leq C_\beta \left(\mathbb{E} \left[\left(\sum_{j=1}^n \mathbb{E}[h_j^2 | \mathcal{F}_{\tau_j(x)}] \right)^{\beta/2} \right] + \mathbb{E} \left[\max_{j \leq n} |h_j|^\beta \right] \right) \\ &\leq C_\beta \left(\mathbb{E} \left[\left(\sum_{j=1}^n \mathbb{E}[(M_j(x, y))^2 | \mathcal{F}_{\tau_j(x)}] \mathbf{1}_{\{\tau_j(x) \leq n\}} \right)^{\beta/2} \right] \right. \\ &\quad \left. + \mathbb{E} \left[\left(1 + \max_{1 \leq j \leq n} |M_j|^\beta \right) \mathbf{1}_{\{\tau_1(x) \leq n\}} \right] \right). \end{aligned}$$

From [KS, Lem. 2], which ensures the existence of a constant $K_2 > 0$ such that $\mathbb{E}[(M_j(x, y))^2 | \mathcal{F}_{\tau_j(x)}] \leq K_2 |x - y|$, we finally get

$$(4.24) \quad \begin{aligned} \mathbb{E}[|\mathcal{M}_n|^\beta] &\leq K_3 \left(\mathbb{E} \left[\left(\sum_{j=1}^n |x - y| \mathbf{1}_{\{\tau_j(x) \leq n\}} \right)^{\beta/2} \right] + \mathcal{B}'_n(x, y) + \mathbb{P}(\tau_1(x) \leq n) \right) \\ &\leq K_3 \left(|x - y|^{\beta/2} \mathbb{E}[(N_n(x))^{\beta/2}] + \mathcal{B}'_n(x, y) + \mathbb{P}(\tau_1(x) \leq n) \right), \end{aligned}$$

for some $K_3 > 0$. Combined with (4.14), (4.20) and (4.21), this ends the proof of Lemma 4.1. \square

Lemma 4.2. *For any $\beta \leq 1$ such that $\gamma > 1 + 1/\beta$,*

$$\sum_{y \in \mathbb{Z}} \mathbb{E} \left[\left| \sum_{r \geq 0} \mathbb{P}(S_1 < -r) (N_n(y) - N_n(y + r)) \right|^\beta \right] = O(n^{1/2 + \beta/4}),$$

as $n \rightarrow \infty$.

Proof. Using (4.13) we have

$$(4.25) \quad \sum_{y \in \mathbb{Z}} \mathbb{E} \left[\left| \sum_{r \geq 0} \mathbb{P}(S_1 < -r) (N_n(y) - N_n(y+r)) \right|^\beta \right] \\ \leq \sum_{y \in \mathbb{Z}} \sum_{r \geq 0} (\mathbb{P}(S_1 < -r))^\beta \mathbb{E} \left[|N_n(y) - N_n(y+r)|^\beta \right].$$

By Lemma 4.1, this quantity is dominated by

$$(4.26) \quad 2K_0 \sum_{y \in \mathbb{Z}} \sum_{r \geq 0} (\mathbb{P}(S_1 < -r))^\beta \left(r^{\beta/2} \mathbb{E} \left[|N_n(y)|^{\beta/2} \right] + (r^\beta (\log n)^\beta + 1) \mathbb{P}(\tau_1(y) \leq n) \right) \\ \leq 2K_0 \sum_{r \geq 0} (\mathbb{P}(S_1 < -r))^\beta (r^\beta + 1) \left(\sum_{y \in \mathbb{Z}} \mathbb{E} \left[|N_n(y)|^{\beta/2} \right] + (\log n)^\beta \mathbb{E}[R_n] \right) \\ = O(n^{1/2+\beta/4}),$$

where the last estimate follows from (4.6) and (4.8), with β replaced by $\beta/2$, and the following application of the Markov inequality:

$$(4.27) \quad \sum_{r \geq 0} (\mathbb{P}(S_1 < -r))^\beta (r^\beta + 1) \leq (\mathbb{E}[|S_1|^\gamma])^\beta \sum_{r \geq 0} r^{-\gamma\beta} (r^\beta + 1),$$

because $\gamma\beta - \beta > 1$ by hypothesis. \square

Later we will need certain inequalities for the local time on the bonds \mathcal{N}_n that are analogous to the ones already seen for the local time on the sites N_n . For this we will exploit the following lemma, which estimates the α -norm distance between the two types of local times.

Lemma 4.3.

$$\sum_{y \in \mathbb{Z}} \mathbb{E}[|\mathcal{N}_n(y) - \mu_\xi N_n(y)|^\alpha] = O(n^{1/2+\alpha/4}).$$

Proof. We write

$$(4.28) \quad \mathcal{N}_n(y) = \mathcal{N}_n^-(y) + \mathcal{N}_n^+(y),$$

where

$$(4.29) \quad \mathcal{N}_n^-(y) := \#\{k = 0, \dots, n-1 : S_k \geq y, S_{k+1} \leq y-1\};$$

$$(4.30) \quad \mathcal{N}_n^+(y) := \#\{k = 0, \dots, n-1 : S_{k+1} \geq y, S_k \leq y-1\}$$

denote the number of crossings of $[y-1, y]$, respectively, from right to left and from left to right.

Observe that, for fixed y , the process

$$(4.31) \quad \left(\mathcal{N}_n^-(y) - \sum_{r \geq 0} N_n(y+r) \mathbb{P}(S_1 < -r), n \in \mathbb{Z}^+ \right)$$

is a martingale with k^{th} increment given by

$$(4.32) \quad d_k^-(y) := \sum_{r \geq 0} (\mathbf{1}_{\{S_k \leq y-1\}} - \mathbb{P}(S_1 < -r)) \mathbf{1}_{\{S_{k-1} = y+r\}}.$$

By Burkholder's inequality [HH, Thm. 2.11], there exists $C_\alpha > 0$ such that

$$(4.33) \quad \mathbb{E} \left[\left| \mathcal{N}_n^-(y) - \sum_{r \geq 0} N_n(y+r) \mathbb{P}(S_1 < -r) \right|^\alpha \right] \leq C_\alpha (A_y^- + B_y^-),$$

where

$$(4.34) \quad A_y^- := \mathbb{E} \left[\left(\sum_{k=1}^n \mathbb{E}[(d_k^-(y))^2 | S_{k-1}] \right)^{\alpha/2} \right];$$

$$(4.35) \quad B_y^- := \mathbb{E} \left[\max_{k \leq n} |d_k^-(y)|^\alpha \right].$$

Let us estimate these terms. Using (4.13) with $\beta = \alpha/2 < 1$, we get for the first term:

$$(4.36) \quad \begin{aligned} \sum_{y \in \mathbb{Z}} A_y^- &= \sum_{y \in \mathbb{Z}} \mathbb{E} \left[\left(\sum_{k=0}^{n-1} \sum_{r \geq 0} \mathbb{P}(S_1 < -r) (1 - \mathbb{P}(S_1 < -r)) \mathbf{1}_{\{S_k = y+r\}} \right)^{\alpha/2} \right] \\ &\leq \sum_{y \in \mathbb{Z}} \mathbb{E} \left[\left(\sum_{k=0}^{n-1} \sum_{r \geq 0} \mathbb{P}(S_1 < -r) \mathbf{1}_{\{S_k = y+r\}} \right)^{\alpha/2} \right] \\ &\leq \sum_{y \in \mathbb{Z}} \mathbb{E} \left[\left(\sum_{r \geq 0} \mathbb{P}(S_1 < -r) N_n(y+r) \right)^{\alpha/2} \right] \\ &\leq \sum_{y \in \mathbb{Z}} \mathbb{E} \left[\sum_{r \geq 0} (\mathbb{P}(S_1 < -r))^{\alpha/2} (N_n(y+r))^{\alpha/2} \right] \\ &\leq (\mathbb{E}[|S_1|^\gamma])^{\alpha/2} \sum_{r \geq 0} r^{-\alpha\gamma/2} \sum_{y \in \mathbb{Z}} \mathbb{E}[(N_n(y))^{\alpha/2}] \\ &= O(n^{1/2+\alpha/4}), \end{aligned}$$

where in the last estimates we have used the Markov inequality for the variable $|S_1|^\gamma$, with $\gamma > 2/\alpha$, and estimate (4.8) with $\beta = \alpha/2$.

As for the second term, in view of (4.35) and (4.32), we write

$$(4.37) \quad \sum_{y \in \mathbb{Z}} B_y^- \leq \sum_{y \in \mathbb{Z}} \mathbb{E}[D_y^- + E_y^-],$$

where

$$(4.38) \quad D_y^- := \max_{k \leq n} \left| \sum_{r \geq 0} \mathbf{1}_{\{S_k \leq y-1, S_{k-1} = y+r\}} \right|^\alpha = \max_{k \leq n} \mathbf{1}_{\{S_k \leq y-1 < y \leq S_{k-1}\}};$$

$$(4.39) \quad E_y^- := \max_{k \leq n} \left| \sum_{r \geq 0} \mathbb{P}(S_1 < -r) \mathbf{1}_{\{S_{k-1} = y+r\}} \right|^\alpha \\ = \max_{k \leq n} |F_{S_1}(y - S_{k-1})|^\alpha \mathbf{1}_{\{S_{k-1} \geq y\}}.$$

Here F_{S_1} is the distribution function of $S_1 = \xi_1$. The expectation of the first of the two terms above is easily estimated as follows, upon a relabeling of the index k :

$$(4.40) \quad \sum_{y \in \mathbb{Z}} \mathbb{E}[D_y^-] = \sum_{y \in \mathbb{Z}} \mathbb{E} \left[\max_{k \leq n-1} \mathbf{1}_{\{S_{k+1} < y \leq S_k\}} \right] \\ \leq \mathbb{E}[\#\{y \in \mathbb{Z} : \exists k \leq n-1 \text{ such that } S_{k+1} < y \leq S_k\}] \\ \leq \mathbb{E} \left[\max_{k \leq n-1} S_k - \min_{\ell \leq n-1} S_\ell \right] = O(\sqrt{n}),$$

having used Doob's maximal inequality and the fact that $\mathbb{E}[S_n^2] = v_\xi n$. As concerns the expectation of (4.39),

$$(4.41) \quad \sum_{y \in \mathbb{Z}} \mathbb{E}[E_y^-] \leq \sum_{y \in \mathbb{Z}} \mathbb{E} \left[\left(F_{S_1} \left(y - \min_{\substack{S_{k-1} \geq y \\ k \leq n}} S_{k-1} \right) \right)^\alpha \mathbf{1}_{\{\exists k \leq n : S_{k-1} \geq y\}} \right] \\ \leq \mathbb{E} \left[\sum_{y = \min_{k \leq n-1} S_k}^{\max_{k \leq n-1} S_k} 1 \right] + \mathbb{E} \left[\sum_{y < \min_{k \leq n-1} S_k} \left(F_{S_1} \left(y - \min_{\substack{S_k \geq y \\ k \leq n-1}} S_k \right) \right)^\alpha \right] \\ \leq \mathbb{E} \left[\max_{k \leq n-1} S_k - \min_{k \leq n-1} S_k + 1 \right] + \sum_{\xi < 0} (F_{S_1}(\xi))^\alpha,$$

where again we have redefined the index k in the second inequality. Now, in the last bound above, the first term is $O(\sqrt{n})$, as in (4.40), and the second term is $O(1)$, because $(F_{S_1}(\xi))^\alpha = O(|\xi|^{-\alpha\gamma})$ by the Markov inequality and the fact that $\alpha\gamma > 1$ by hypothesis; cf. (4.27).

Together with (4.37) and (4.40) we conclude that

$$(4.42) \quad \sum_{y \in \mathbb{Z}} B_y^- = O(\sqrt{n}),$$

whence, in light of (4.33) and (4.36),

$$(4.43) \quad \sum_{y \in \mathbb{Z}} \mathbb{E} \left[\left| \mathcal{N}_n^-(y) - \sum_{r \geq 0} N_n(y+r) \mathbb{P}(S_1 < -r) \right|^\alpha \right] = O(n^{1/2+\alpha/4}).$$

Applying (4.13) and Lemma 4.2, both with $\beta = \alpha$ (since $\gamma > 2/\alpha > 1 + 1/\alpha$), we obtain

$$(4.44) \quad \sum_{y \in \mathbb{Z}} \mathbb{E} \left[\left| \mathcal{N}_n^-(y) - N_n(y) \sum_{r \geq 0} \mathbb{P}(S_1 < -r) \right|^\alpha \right] = O(n^{1/2 + \alpha/4}).$$

In the same way one can prove that

$$(4.45) \quad \sum_{y \in \mathbb{Z}} \mathbb{E} \left[\left| \mathcal{N}_n^+(y) - N_n(y) \sum_{r \geq 0} \mathbb{P}(S_1 > r) \right|^\alpha \right] = O(n^{1/2 + \alpha/4}).$$

Lemma 4.3 then follows from the two previous estimates and (4.13) with $\beta = \alpha$, because $\mu_\xi = \mathbb{E}[|S_1|] = \sum_{r \geq 0} \mathbb{P}(|S_1| > r)$. \square

Observation. Lemma 4.3 is the only part of the paper where the assumption $\gamma > 2/\alpha$ is used.

4.2. Finite-dimensional distributions. To prove the convergence of the finite-dimensional joint distribution $(\bar{\omega}^{(q)}, \bar{S}^{(q)}, \bar{T}^{(q)})$, as $q \rightarrow \infty$, we use the method of the characteristic function, as in [KS].

Notation. In what follows, the asymptotic equivalence $a(\theta) \sim b(\theta)$, as $\theta \rightarrow \theta_0$, means that $\lim_{\theta \rightarrow \theta_0} a(\theta)/b(\theta) = 1$. In case $b(\theta_j) = 0$, for $\theta_j \rightarrow \theta_0$, the equivalence means that $a(\theta_j) = 0$, for all j . A similar notation will be used w.r.t. to other limits, such as $q \rightarrow \infty$.

The assumption on the distribution of the gaps ζ_j , see beginning of Section 2, implies that the following limit exists

$$(4.46) \quad c_0 := \lim_{z \rightarrow +\infty} z^\alpha \mathbb{P}(\zeta_1 \geq z) > 0.$$

Equivalently, denoting by $\phi_{\mathcal{X}}$ the characteristic function of a given random variable \mathcal{X} , the following relation holds for $\theta \rightarrow 0$:

$$(4.47) \quad 1 - \phi_{\zeta_1}(\theta) \sim -\log(\phi_{Z_1}(\theta)) = c_1 |\theta|^\alpha \left(1 - i \tan \frac{\pi\alpha}{2} \operatorname{sgn}(\theta) \right),$$

where Z_1 is the α -stable variable whose basin of attraction contains ζ_1 , cf. Section 2, and $c_1 := \Gamma(1-\alpha)c_0 \cos(\alpha\pi/2)$ [II, §2.2]. Hence, there exists a continuous increasing function $\epsilon_0 : \mathbb{R}_0^+ \rightarrow \mathbb{R}_0^+$ such that $\epsilon_0(0) = 0$ and, for all $\theta \in \mathbb{R}$,

$$(4.48) \quad |\phi_{\zeta_1}(\theta) - \phi_{Z_1}(\theta)| \leq |\theta|^\alpha \epsilon_0(|\theta|).$$

Proof of the convergence of the finite-dimensional distributions. For a given integer $m \geq 1$, we take $3m$ real numbers $\kappa_1, \dots, \kappa_m, \mu_1, \dots, \mu_m, \theta_1, \dots, \theta_m$, and $3m$ real numbers $t_1, \dots, t_m, s_1, \dots, s_m, v_1, \dots, v_m$ satisfying $0 < t_1 < \dots < t_m$, $0 < v_1 < \dots < v_m$ and $s_1 < \dots < s_m$. Let us denote by $\varphi^{(q)}$ the characteristic function of

$$(4.49) \quad (\bar{\omega}^{(q)}(s_1), \dots, \bar{\omega}^{(q)}(s_m), \bar{S}^{(q)}(v_1), \dots, \bar{S}^{(q)}(v_m), \bar{T}^{(q)}(t_1), \dots, \bar{T}^{(q)}(t_m))$$

evaluated in $(\kappa_1, \dots, \kappa_m, \mu_1, \dots, \mu_m, \theta_1, \dots, \theta_m)$. Conditioning with respect to S , we have

$$(4.50) \quad \varphi^{(q)} = \mathbb{E} \left[\mathbb{E} \left[e^{i \sum_{j=1}^m (\kappa_j \bar{\omega}^{(q)}(s_j) + \theta_j \bar{T}^{(q)}(t_j))} \middle| S \right] e^{i \sum_{j=1}^m \mu_j \bar{S}^{(q)}(v_j)} \right].$$

Setting $\tilde{\kappa}_j := \text{sgn}(s_j) \kappa_j$, the inner expectation can be rewritten as

$$(4.51) \quad \begin{aligned} & \mathbb{E} \left[e^{i \sum_{j=1}^m (\kappa_j \bar{\omega}^{(q)}(s_j) + \theta_j \bar{T}^{(q)}(t_j))} \middle| S \right] \\ &= \mathbb{E} \left[\exp \left(i \sum_{y \in \mathbb{Z}} \sum_{j=1}^m \left(\theta_j q^{-(1+\alpha)/2\alpha} \mathcal{N}_{[t_j q]}(y) + \tilde{\kappa}_j q^{-1/2\alpha} \mathbf{1}_{I_{s_j \sqrt{q}}}(y) \right) \zeta_y \right) \middle| S \right] \\ &= \mathbb{E} \left[\prod_{y \in \mathbb{Z}} \phi_{\zeta_1} \left(\sum_{j=1}^m \left(\theta_j q^{-(1+\alpha)/2\alpha} \mathcal{N}_{[t_j q]}(y) + \tilde{\kappa}_j q^{-1/2\alpha} \mathbf{1}_{I_{s_j \sqrt{q}}}(y) \right) \right) \middle| S \right], \end{aligned}$$

with

$$(4.52) \quad I_s := \begin{cases} (0, s], & s \geq 0; \\ (s, 0], & s < 0. \end{cases}$$

Altogether we get

$$(4.53) \quad \varphi^{(q)} = \mathbb{E} \left[\prod_{y \in \mathbb{Z}} \phi_{\zeta_1} \left(\sum_{j=1}^m \left(\theta_j q^{-(1+\alpha)/2\alpha} \mathcal{N}_{[t_j q]}(y) + \tilde{\kappa}_j q^{-1/2\alpha} \mathbf{1}_{I_{s_j \sqrt{q}}}(y) \right) \right) e^{i \sum_{j=1}^m \mu_j \bar{S}^{(q)}(v_j)} \right].$$

Convention. In the remainder of this proof, for $a \in \mathbb{R}$, we shall use the notation

$$(4.54) \quad |a|_{\text{sgn}}^\alpha := |a|^\alpha \text{sgn}(a).$$

Also, we shall introduce pairs of quantities, say A, A_{sgn} , which are defined similarly except that a certain expression $|\cdot|^\alpha$ in the first definition is replaced by the corresponding expression $|\cdot|_{\text{sgn}}^\alpha$ in the second one. If A, A_{sgn} and B, B_{sgn} are two such pairs, we shall use the imprecise notation

$$(4.55) \quad A_\sigma = B_\sigma, \quad \text{for } \sigma \in \{, \text{sgn}\}$$

to mean that $A = B$ and $A_{\text{sgn}} = B_{\text{sgn}}$.

If we set

$$(4.56) \quad V^{(q)}(y) := \sum_{j=1}^m \left(\theta_j q^{-(1+\alpha)/2\alpha} \mathcal{N}_{[t_j q]}(y) + \tilde{\kappa}_j q^{-1/2\alpha} \mathbf{1}_{I_{s_j \sqrt{q}}}(y) \right),$$

the following is an identity by (4.47):

$$(4.57) \quad \begin{aligned} & \prod_{y \in \mathbb{Z}} \phi_{\zeta_1} \left(\sum_{j=1}^m \left(\theta_j q^{-(1+\alpha)/2\alpha} \mathcal{N}_{[t_j q]}(y) + \tilde{\kappa}_j q^{-1/2\alpha} \mathbf{1}_{I_{s_j \sqrt{q}}}(y) \right) \right) \\ &= \exp \left(-c_1 \sum_{y \in \mathbb{Z}} \left(|V^{(q)}(y)|^\alpha - i |V^{(q)}(y)|_{\text{sgn}}^\alpha \tan \frac{\pi \alpha}{2} \right) \right). \end{aligned}$$

Therefore, in view of (4.48), we obtain

$$\begin{aligned}
(4.58) \quad & \left| \varphi^{(q)} - \mathbb{E} \left[\exp \left(-c_1 \sum_{y \in \mathbb{Z}} \left(|V^{(q)}(y)|^\alpha - i |V^{(q)}(y)|_{\text{sgn}}^\alpha \tan \frac{\pi\alpha}{2} \right) + i \sum_{j=1}^m \mu_j \bar{S}^{(q)}(v_j) \right) \right] \right| \\
& \leq \mathbb{E} \left[\sum_{y \in \mathbb{Z}} |V^{(q)}(y)|^\alpha \epsilon_0(|V^{(q)}(y)|) \right] \\
& \leq \mathbb{E} \left[\sum_{y \in \mathbb{Z}} |V^{(q)}(y)|^\alpha \right] \epsilon_0 \left(\sum_{j=1}^m (|\theta_j| t_m q^{(\alpha-1)/2\alpha} + |\kappa_j| q^{-1/2\alpha}) \right),
\end{aligned}$$

where in the last inequality we have used that $\mathcal{N}_{[t_j q]} \leq t_m q$. Now, from the definition of $V^{(q)}$, using (4.13) with $\beta = \alpha < 1$, we see that the last term in the above inequality is bounded above by

$$(4.59) \quad \mathbb{E} \left[\sum_{j=1}^m \left(|\theta_j|^\alpha q^{-(\alpha+1)/2} \sum_{y \in \mathbb{Z}} (\mathcal{N}_{[t_j q]}(y))^\alpha + |\tilde{\kappa}_j|^\alpha q^{-1/2} |s_j|^\alpha q^{\alpha/2} \right) \right] \epsilon_q,$$

where $\epsilon_q := \epsilon_0(\sum_{j=1}^m (|\theta_j| t_j q^{(\alpha-1)/2\alpha} + |\kappa_j| q^{-1/2\alpha}))$. Notice that $\epsilon_q \rightarrow 0$, as $q \rightarrow \infty$. Now, by Lemma 4.3 and (4.8),

$$(4.60) \quad \sum_{y \in \mathbb{Z}} \mathbb{E}[(\mathcal{N}_n(y))^\alpha] = \sum_{y \in \mathbb{Z}} \mu_\xi^\alpha \mathbb{E}[(N_n(y))^\alpha] + O(n^{1/2+\alpha/4}) = O(n^{1/2+\alpha/4}).$$

We conclude that, for $q \rightarrow \infty$,

$$(4.61) \quad \varphi^{(q)} \sim \mathbb{E} \left[\exp \left(-c_1 \sum_{y \in \mathbb{Z}} \left(|V^{(q)}(y)|^\alpha - i |V^{(q)}(y)|_{\text{sgn}}^\alpha \tan \frac{\pi\alpha}{2} \right) + i \sum_{j=1}^m \mu_j \bar{S}^{(q)}(v_j) \right) \right].$$

In order to take the limit of the above r.h.s. we define

$$(4.62) \quad \tilde{V}^{(q)}(y) := \sum_{j=1}^m \left(\theta_j q^{-(1+\alpha)/2\alpha} \mu_\xi N_{[t_j q]}(y) + \tilde{\kappa}_j q^{-1/2\alpha} \mathbf{1}_{I_{s_j \sqrt{q}}}(y) \right).$$

Applying

$$(4.63) \quad ||a|_\sigma^\alpha - |b|_\sigma^\alpha| \leq 2^{1-\alpha} |a - b|^\alpha,$$

which holds true for both $\sigma \in \{, \text{sgn}\}$, since $\alpha \leq 1$, and using (4.56), we obtain

$$\begin{aligned}
(4.64) \quad & \left| \sum_{y \in \mathbb{Z}} |V^{(q)}(y)|_\sigma^\alpha - \sum_{y \in \mathbb{Z}} |\tilde{V}^{(q)}(y)|_\sigma^\alpha \right| \\
& \leq 2^{1-\alpha} \sum_{y \in \mathbb{Z}} |V^{(q)}(y) - \tilde{V}^{(q)}(y)|^\alpha \\
& \leq 2^{1-\alpha} \sum_{y \in \mathbb{Z}} \left| \sum_{j=1}^m \theta_j q^{-(1+\alpha)/2\alpha} (\mathcal{N}_{[t_j q]}(y) - \mu_\xi N_{[t_j q]}(y)) \right|^\alpha \\
& \leq 2^{1-\alpha} \sum_{y \in \mathbb{Z}} \sum_{j=1}^m |\theta_j|^\alpha q^{-(1+\alpha)/2} |\mathcal{N}_{[t_j q]}(y) - \mu_\xi N_{[t_j q]}(y)|^\alpha,
\end{aligned}$$

where in the last step we have again used (4.13) with $\beta = \alpha$. Therefore, by Lemma 4.3,

$$(4.65) \quad \mathbb{E} \left[\left| \sum_{y \in \mathbb{Z}} |V^{(q)}(y)|_\sigma^\alpha - \sum_{y \in \mathbb{Z}} |\tilde{V}^{(q)}(y)|_\sigma^\alpha \right| \right] = O(q^{-\alpha/4}).$$

Setting

$$(4.66) \quad G_\sigma^{(q)} := \sum_{y \in \mathbb{Z}} |\tilde{V}^{(q)}(y)|_\sigma^\alpha$$

and combining (4.65) with (4.61), we see, by means of the Lebesgue's Dominated Convergence Theorem, that, for $q \rightarrow \infty$,

$$(4.67) \quad \varphi^{(q)} \sim \mathbb{E} \left[\exp \left(-c_1 \left(G^{(q)} - i G_{\text{sgn}}^{(q)} \tan \frac{\pi\alpha}{2} \right) + i \sum_{j=1}^m \mu_j \bar{S}^{(q)}(v_j) \right) \right].$$

We then define

$$(4.68) \quad G_\sigma := \int_{\mathbb{R}} \left| \sum_{j=1}^m \left(\theta_j \mu_\xi L_{t_j}(y) + \tilde{\kappa}_j \mathbf{1}_{I_{s_j}}(y) \right) \right|_\sigma^\alpha dy.$$

The following lemma will be proved below.

Lemma 4.4. *For fixed values of the parameters $\{\kappa_i, \mu_i, \theta_i, t_i, s_i, v_i\}_{i=1}^m$, as described earlier, the random vector $(G^{(q)}, G_{\text{sgn}}^{(q)}, \sum_{j=1}^m \mu_j \bar{S}^{(q)}(v_j))$ converges in distribution, as $q \rightarrow \infty$, to the vector $(G, G_{\text{sgn}}, \sum_{j=1}^m \mu_j B(v_j))$.*

It follows from (4.67), Lemma 4.4 and dominated convergence that

$$(4.69) \quad \lim_{q \rightarrow \infty} \varphi^{(q)} = \mathbb{E} \left[\exp \left(-c_1 \left(G - i G_{\text{sgn}} \tan \frac{\pi\alpha}{2} \right) + i \sum_{j=1}^m \mu_j B(v_j) \right) \right],$$

which is the characteristic function of

$$(4.70) \quad (Z(s_1), \dots, Z(s_m), B(v_1), \dots, B(v_m), \Delta(t_1), \dots, \Delta(t_m))$$

evaluated in $(\kappa_1, \dots, \kappa_m, \mu_1, \dots, \mu_m, \theta_1, \dots, \theta_m)$. In fact,

$$(4.71) \quad \sum_{j=1}^m (\kappa_j Z(s_j) + \theta_j \Delta(t_j)) = \sum_{j=1}^m \int_0^{+\infty} (\kappa_j \mathbf{1}_{I_j}(x) + \theta_j L_{t_j}(x)) dZ_+(x) \\ + \int_0^{+\infty} (\kappa_j \mathbf{1}_{I_j}(-x) + \theta_j L_{t_j}(-x)) dZ_-(x)$$

and so, conditionally to B ,

$$(4.72) \quad \mathbb{E} \left[e^{i \sum_{j=1}^m (\kappa_j Z(s_j) + \theta_j \Delta(t_j))} \middle| B \right] = \exp \left(-c_1 \left(G - iG_{\text{sgn}} \tan \frac{\pi\alpha}{2} \right) \right).$$

□

Proof of Lemma 4.4. We follow the proof of [KS, Lem. 6]. For all $q, \vartheta \in \mathbb{R}^+$, $M \in \mathbb{Z}^+$ and $\sigma \in \{, \text{sgn}\}$, we define

$$(4.73) \quad \mathcal{V}_\sigma(\vartheta, M, q) := \vartheta^{1-\alpha} \sum_{|k| \leq M} |\mathcal{T}(k, q)|_\sigma^\alpha,$$

where

$$(4.74) \quad \mathcal{T}(k, q) := \sum_{j=1}^m \left(\sum_{x=\lceil k\vartheta\sqrt{q} \rceil + 1}^{\lceil (k+1)\vartheta\sqrt{q} \rceil} \left(\frac{1}{q} \theta_j \mu_\xi N_{\lfloor t_j q \rfloor}(x) + \frac{1}{\sqrt{q}} \tilde{\kappa}_j \mathbf{1}_{I_{s_j\sqrt{q}}}(x) \right) \right),$$

where $\lceil a \rceil$ is the smallest integer larger than or equal to a .

As in [KS], we decompose $G_\sigma^{(q)} - \mathcal{V}_\sigma(\vartheta, M, q)$ as follows:

$$(4.75) \quad G_\sigma^{(q)} - \mathcal{V}_\sigma(\vartheta, M, q) = U_\sigma(\vartheta, M, q) + W_\sigma^1(\vartheta, M, q) + W_\sigma^2(\vartheta, M, q),$$

where

$$(4.76) \quad U_\sigma(\vartheta, M, q) := \sum_{x \in A_{\vartheta, M, q}} \left| \sum_{j=1}^m \left(q^{-(1+\alpha)/2\alpha} \theta_j \mu_\xi N_{\lfloor t_j q \rfloor}(x) + q^{-1/2\alpha} \tilde{\kappa}_j \mathbf{1}_{I_{s_j\sqrt{q}}}(x) \right) \right|_\sigma^\alpha,$$

with $A_{\vartheta, M, q} := \mathbb{Z} \setminus \{ \lceil -M\vartheta\sqrt{q} \rceil + 1, \dots, \lceil (M+1)\vartheta\sqrt{q} \rceil \}$;

$$(4.77) \quad W_\sigma^1(\vartheta, M, q) := \sum_{|k| \leq M} \sum_{x \in E_{k, q}} q^{-(1+\alpha)/2} W_\sigma^{1, k, q}(x),$$

with $E_{k, q} := \{ \lceil k\vartheta\sqrt{q} \rceil + 1, \dots, \lceil (k+1)\vartheta\sqrt{q} \rceil \}$ and

$$(4.78) \quad W_\sigma^{1, k, q}(x) := \left| \sum_{j=1}^m \theta_j \mu_\xi N_{\lfloor t_j q \rfloor}(x) + \sqrt{q} \tilde{\kappa}_j \mathbf{1}_{I_{s_j\sqrt{q}}}(x) \right|_\sigma^\alpha - q^\alpha (\#E_{k, q})^{-\alpha} |\mathcal{T}(k, q)|_\sigma^\alpha;$$

and finally

$$(4.79) \quad W_\sigma^2(\vartheta, M, q) := \sum_{|k| \leq M} \left(q^{(\alpha-1)/2} (\#E_{k, q})^{1-\alpha} - \vartheta^{1-\alpha} \right) |\mathcal{T}(k, q)|_\sigma^\alpha.$$

Convention. With reference to definition (4.77), we convene that for those values of ϑ, k, q such that $\lceil k\vartheta\sqrt{q} \rceil = \lceil (k+1)\vartheta\sqrt{q} \rceil$, $E_{k, q} := \emptyset$.

By [KS, Lem. 1, eq. (2.11)] applied with $s = \lfloor t_m q \rfloor$ and $A = M\vartheta/\sqrt{t_m}$, there exists a function $\eta : \mathbb{R}^+ \rightarrow [0, 1]$, vanishing at $+\infty$, such that

$$(4.80) \quad \begin{aligned} & \sup_{q \geq 1/t_m} \mathbb{P}(U_\sigma(\vartheta, M, q) \neq 0) \\ & \leq \sup_{q \geq 1/t_m} \mathbb{P}(\exists x : |x| \geq M\vartheta\sqrt{q}, N_{\lfloor t_m q \rfloor}(x) \neq 0) + \mathbf{1}_{\{\max\{|s_1|, |s_m|\} \geq M\vartheta\}} \\ & = \eta(M\vartheta). \end{aligned}$$

Let us prove that there exist $K_1, K_2 > 0$ such that, for every $M \in \mathbb{Z}^+$,

$$(4.81) \quad \sup_{q > 1} \mathbb{E}[|W_\sigma^1(\vartheta, M, q)|] \leq K_1 M \vartheta^{1+\alpha/2} + K_2 \vartheta.$$

Using (4.63) and (4.13) (with $\beta = \alpha$), we have

$$(4.82) \quad \begin{aligned} & 2^{\alpha-1} q^{-(1+\alpha)/2} \mathbb{E}[|W_\sigma^{1,k,q}(x)|] \\ & \leq q^{-(1+\alpha)/2} \mathbb{E} \left[\left| \sum_{j=1}^m \left(\theta_j \mu_\xi N_{\lfloor t_j q \rfloor}(x) + \sqrt{q} \tilde{\kappa}_j \mathbf{1}_{I_{s_j \sqrt{q}}}(x) \right) - q \frac{\mathcal{T}(k, q)}{\#E_{k,q}} \right|^\alpha \right] \\ & \leq q^{-(1+\alpha)/2} \mathbb{E} \left[\left| \sum_{j=1}^m \theta_j \mu_\xi \left(N_{\lfloor t_j q \rfloor}(x) - \frac{1}{\#E_{k,q}} \sum_{y \in E_{k,q}} N_{\lfloor t_j q \rfloor}(y) \right) \right|^\alpha \right] \\ & \quad + q^{-(1+\alpha)/2} \mathbb{E} \left[\left| \sum_{j=1}^m \sqrt{q} \tilde{\kappa}_j \left(\mathbf{1}_{I_{s_j \sqrt{q}}}(x) - \frac{1}{\#E_{k,q}} \sum_{y \in E_{k,q}} \mathbf{1}_{I_{s_j \sqrt{q}}}(y) \right) \right|^\alpha \right] \end{aligned}$$

Let us sum the r.h.s. of the last inequality over $|k| \leq M$ and $x \in E_{k,q}$. By [KS, eq. (3.10)], the sum of the first term is bounded above by some constant times $M\vartheta^{1+\alpha/2}$. As for the second term, we observe that the function $E_{k,q} \ni x \mapsto \mathbf{1}_{I_{s_j \sqrt{q}}}(x)$ is constant for all values of $|k| \leq M$ except at most one (corresponding to certain cases where the convex hull of $E_{k,q}$ contains $s_j \sqrt{q}$). Thus, for all non-exceptional values of k and $x \in E_{k,q}$, the term within parentheses is zero. This and a cancelation of the exponents of q show that the sum over $|k| \leq M$ and $x \in E_{k,q}$ of the last term above is of the order of $\vartheta \sum_{j=1}^m |\kappa_j|$. So (4.81) is in force.

Now fix $a, b, c \in \mathbb{R}$ and $\varepsilon > 0$. By (4.81) and the inequality $|e^{iz} - 1| \leq |z|$, for $z \in \mathbb{R}$, we see that one can find values of M, ϑ such that $M\vartheta$ is arbitrarily large but ϑ is so small that

$$(4.83) \quad \sup_{q > 1} \mathbb{E}[\exp(i(aW^1 + bW_{\text{sgn}}^1)(\vartheta, M, q)) - 1] < \varepsilon.$$

Also, let us verify that, for all M, ϑ , the variable $W_\sigma^2(\vartheta, M, q)$ converges in probability to 0 as $q \rightarrow \infty$. For the contribution to $W_\sigma^2(\vartheta, M, q)$ coming from the first term in $\mathcal{T}(k, q)$, cf. (4.79) and (4.74), this is proved in [KS] (see the argument after eq. (3.7) therein). The contribution from the second term of $\mathcal{T}(k, q)$ is bounded by

$$(4.84) \quad \sum_{|k| \leq M} \left(|q^{(\alpha-1)/2} (\#E_{k,q})^{1-\alpha} - \vartheta^{1-\alpha}| \vartheta^\alpha \sum_{j=1}^m |\kappa_j|^\alpha \right),$$

but, for all k , $(q^{(\alpha-1)/2} (\#E_{k,q})^{1-\alpha} - \vartheta^{1-\alpha}) \rightarrow 0$, as $q \rightarrow \infty$. Hence, for all large enough q (depending on M, ϑ),

$$(4.85) \quad \mathbb{E}[\exp(i(aW^2 + bW_{\text{sgn}}^2)(\vartheta, M, q)) - 1] < \varepsilon.$$

Let us introduce the random function

$$(4.86) \quad H_\sigma(\vartheta, M) := \vartheta^{1-\alpha} \sum_{|k| \leq M} \left| \int_{k\vartheta}^{(k+1)\vartheta} \sum_{j=1}^m \left(\theta_j \mu_\xi L_{t_j}(x) + \tilde{\kappa}_j \mathbf{1}_{I_{s_j}}(x) \right) dx \right|^\alpha.$$

It is easy to check that, when $M \rightarrow \infty$ and $\vartheta \rightarrow 0^+$ in such a way that $M\vartheta \rightarrow +\infty$, $H_\sigma(\vartheta, M)$ converges almost surely to G_σ , for both labels σ .

Therefore, by (4.80), (4.83), (4.85) and the limit just discussed, one can find values of M, ϑ such that, for all sufficiently large q ,

$$(4.87) \quad \mathbb{E}[\left| \exp(i(aG^{(q)} + bG_{\text{sgn}}^{(q)})) - \exp(i(a\mathcal{V} + b\mathcal{V}_{\text{sgn}})(\vartheta, M, q)) \right|] < \varepsilon;$$

$$(4.88) \quad \mathbb{E}[\left| \exp(i(aH + bH_{\text{sgn}})(\vartheta, M)) - \exp(i(aG + bG_{\text{sgn}})) \right|] < \varepsilon.$$

Moreover, by [KS, eqs. (2.4)-(2.8)], when $q \rightarrow \infty$,

$$(4.89) \quad \left(\left(\frac{1}{q} \sum_{x=\lceil k\vartheta\sqrt{q} \rceil}^{\lceil (k+1)\vartheta\sqrt{q} \rceil} N_{\lfloor t_j q \rfloor}, \bar{S}^{(q)}(v_j) \right), j = 1, \dots, m \right)$$

converges in distribution to

$$(4.90) \quad \left(\left(\int_{k\vartheta}^{(k+1)\vartheta} L_{t_j(x)} dx, B(v_j) \right), j = 1, \dots, m \right).$$

We also have the following convergence of Riemann sums:

$$(4.91) \quad \lim_{q \rightarrow \infty} \frac{1}{\sqrt{q}} \sum_{x=\lceil k\vartheta\sqrt{q} \rceil}^{\lceil (k+1)\vartheta\sqrt{q} \rceil} \mathbf{1}_{I_{s_j\sqrt{q}}}(x) = \int_{k\vartheta}^{(k+1)\vartheta} \mathbf{1}_{I_{s_j}}(x) dx.$$

Therefore the vector $(\mathcal{V}(\vartheta, M, q), \mathcal{V}_{\text{sgn}}(\vartheta, M, q), \sum_{j=1}^m \mu_j \bar{S}^{(q)}(v_j))$ converges in distribution to $(H(\vartheta, M), H_{\text{sgn}}(\vartheta, M), \sum_{j=1}^m \mu_j B(v_j))$. This means that there exists $q_0 = q_0(M, \vartheta) > 0$ such that, for all $q > q_0$, (4.87) holds together with the following:

$$(4.92) \quad \mathbb{E} \left[\left| \exp \left(i(a\mathcal{V} + b\mathcal{V}_{\text{sgn}})(\vartheta, M, q) + ic \sum_{j=1}^m \mu_j \bar{S}^{(q)}(v_j) \right) - \exp \left(i(aH + bH_{\text{sgn}})(\vartheta, M) + ic \sum_{j=1}^m \mu_j B(v_j) \right) \right| \right] < \varepsilon.$$

Putting the last three inequalities together we obtain that, for all $q > q_0$,

$$(4.93) \quad \left| \mathbb{E} \left[\exp \left(i \left(aG^{(q)} + bG_{\text{sgn}}^{(q)} + c \sum_{j=1}^m \mu_j \bar{S}^{(q)}(v_j) \right) \right) \right] - \mathbb{E} \left[\exp \left(i \left(aG + bG_{\text{sgn}} + c \sum_{j=1}^m \mu_j B(v_j) \right) \right) \right] \right| < 3\varepsilon.$$

In other words, as $q \rightarrow \infty$, $(G^{(q)}, G_{\text{sgn}}^{(q)}, \sum_{j=1}^m \mu_j \bar{S}^{(q)}(v_j))$ converges in distribution to $(G, G_{\text{sgn}}, \sum_{j=1}^m \mu_j B(v_j))$. \square

4.3. Tightness. It is a standard result [S] that the sequences $(\bar{\omega}^{(q)}(x), x > 0)_{q \in \mathbb{R}^+}$ and $(-\bar{\omega}^{(q)}(-x), x > 0)_{q \in \mathbb{R}^+}$, cf. (3.1), converge in distribution w.r.t. $(\mathbb{D}(\mathbb{R}_0^+), J_1)$ to the process Z_+ (or Z_-) defined in Section 2. In particular they are tight. The same is true for $(\bar{S}^{(q)}(t), t \geq 0)_{q \in \mathbb{R}^+}$, which, due to the functional central limit theorem, converges to B .

Hence it remains to prove the tightness of $(\bar{T}^{(q)}(t), t \geq 0)_{q \in \mathbb{R}^+}$ on $(\mathbb{D}(\mathbb{R}_0^+), J_1)$. Let us fix $T > 0$ and prove the tightness of $(\bar{T}^{(q)}(t), t \in [0, T])_{q \in \mathbb{R}^+}$ in $(\mathbb{D}[0, T], J_1)$ following the proof of [KS, Lem. 7].

For a given $\rho > 0$, we set $\zeta_{y,q} = \zeta_{y,q,\rho} := \zeta_y \mathbf{1}_{\{\zeta_y \leq \rho q^{1/2\alpha}\}}$ and define the process

$$(4.94) \quad \bar{T}_0^{(q)}(t) = \bar{T}_0^{(q,\rho)}(t) := q^{-(1+\alpha)/2\alpha} \sum_{y \in \mathbb{Z}} \mathcal{N}_{[tq]}(y) \zeta_{y,q},$$

that will approximate the process $\bar{T}^{(q)}$. To control the approximation error first notice that, for any $c > 0$,

$$(4.95) \quad \begin{aligned} & \mathbb{P}(\exists t \in [0, T] : \bar{T}^{(q)}(t) \neq \bar{T}_0^{(q)}(t)) \\ &= \mathbb{P}\left(\exists t \in [0, T] : \sum_{y \in \mathbb{Z}} \mathcal{N}_{[tq]}(y) (\zeta_y - \zeta_{y,q}) \neq 0\right) \\ &\leq \mathbb{P}(\exists y : |y| \leq c\sqrt{q}, \zeta_y \neq \zeta_{y,q}) + \mathbb{P}(\exists t \in [0, T], \exists y : |y| > c\sqrt{q}, \mathcal{N}_{[tq]}(y) > 0). \end{aligned}$$

Now, by virtue of standard results on the maximum of a random walk, we fix $c > 0$ such that

$$(4.96) \quad \mathbb{P}(\exists t \in [0, T], \exists y : |y| > c\sqrt{q}, \mathcal{N}_{[tq]}(y) > 0) = \mathbb{P}\left(\max_{k \leq [Tq]} |S_k| > c\sqrt{q}\right) < \frac{\varepsilon}{4}.$$

We then choose ρ such that, using (4.46) as well,

$$(4.97) \quad \mathbb{P}(\exists y : |y| \leq c\sqrt{q}, \zeta_y \neq \zeta_{y,q}) \leq 3c\sqrt{q} \mathbb{P}(\zeta_1 \neq \zeta_{1,q}) \leq 3cc_0 \rho^{-\alpha} < \frac{\varepsilon}{4}.$$

Inserting these estimates in (4.95) we obtain

$$(4.98) \quad \mathbb{P}(\exists t \in [0, T] : \bar{T}^{(q)}(t) \neq \bar{T}_0^{(q)}(t)) < \frac{\varepsilon}{2},$$

showing that it is enough to prove the tightness of $(\bar{T}_0^{(q)}(t), t \in [0, T])_{q \in \mathbb{R}^+}$ in $(\mathbb{D}[0, T], J_1)$.

We observe that, due to [B, Thm. 13.5, eq. (13.14)], it is enough to prove the existence of $K_1 > 0$ such that, for every s, t with $0 < s < t < T$,

$$(4.99) \quad \mathbb{E} \left[\left| \bar{T}_0^{(q)}(t) - \bar{T}_0^{(q)}(s) \right| \right] \leq K_1 (q^{-1}(\lfloor tq \rfloor - \lfloor sq \rfloor))^{3/4}.$$

Indeed, together with Hölder's inequality, this would imply that for every $0 < s < r < t < T$,

$$(4.100) \quad \mathbb{E} \left[\left| \bar{T}_0^{(q)}(r) - \bar{T}_0^{(q)}(s) \right|^{1/2} \left| \bar{T}_0^{(q)}(t) - \bar{T}_0^{(q)}(r) \right|^{1/2} \right] \leq K_1 (2(t-s))^{3/4}.$$

The above uses that, when $t-s \geq 1/q$, $\lfloor tq \rfloor - \lfloor sq \rfloor \leq 2q(t-s)$ and, when $t-s < 1/q$, either $\lfloor rq \rfloor = \lfloor sq \rfloor$ or $\lfloor rq \rfloor = \lfloor tq \rfloor$ making the l.h.s. of (4.100) null.

To verify (4.99), we decompose $\bar{T}_0^{(q)}$ in $\bar{T}_1^{(q)} + \bar{T}_2^{(q)}$, where

$$(4.101) \quad \bar{T}_1^{(q)}(t) := q^{-(1+\alpha)/2\alpha} \sum_{y \in \mathbb{Z}} \mathcal{N}_{\lfloor tq \rfloor}(y) \mathbb{E}[\zeta_{1,q}];$$

$$(4.102) \quad \bar{T}_2^{(q)}(t) := q^{-(1+\alpha)/2\alpha} \sum_{y \in \mathbb{Z}} \mathcal{N}_{\lfloor tq \rfloor}(y) \bar{\zeta}_{y,q},$$

with $\bar{\zeta}_{y,q} := \zeta_{y,q} - \mathbb{E}[\zeta_{1,q}]$. Using (4.46), we see that, as $q \rightarrow \infty$,

$$(4.103) \quad \begin{aligned} \mathbb{E}[\zeta_{1,q}] &= \int_0^{+\infty} \mathbb{P}(\zeta_{1,q} > u) du = \int_0^{\rho q^{1/2\alpha}} (\mathbb{P}(\zeta_1 > u) - \mathbb{P}(\zeta_1 > \rho q^{1/2\alpha})) du \\ &\sim c_0 \frac{\alpha}{1-\alpha} \rho^{1-\alpha} q^{(1-\alpha)/2\alpha} =: K_2 q^{(1-\alpha)/2\alpha}, \end{aligned}$$

for a certain positive constant K_2 . Analogously there exists $K_3 > 0$ such that

$$(4.104) \quad \mathbb{E}[\zeta_{1,q}^2] = \int_0^{+\infty} \mathbb{P}(\zeta_{1,q}^2 > u) du \leq \int_0^{\rho^2 q^{1/\alpha}} \mathbb{P}(\zeta_1 > \sqrt{u}) du \sim K_3 q^{(2-\alpha)/2\alpha}.$$

Now, on one hand, since S and ζ are independent, we obtain

$$(4.105) \quad \begin{aligned} \mathbb{E} \left[\left| \bar{T}_1^{(q)}(t) - \bar{T}_1^{(q)}(s) \right| \right] &\leq q^{-(1+\alpha)/2\alpha} \sum_{y \in \mathbb{Z}} \mathbb{E} \left[\left| \mathcal{N}_{\lfloor tq \rfloor}(y) - \mathcal{N}_{\lfloor sq \rfloor}(y) \right| \right] \mathbb{E}[\zeta_{1,q}] \\ &\leq q^{-(1+\alpha)/2\alpha} \sum_{j=\lfloor sq \rfloor+1}^{\lfloor tq \rfloor} \mathbb{E}[|S_j - S_{j-1}|] \mathbb{E}[\zeta_{1,q}] \\ &\leq 2K_2 q^{-1}(\lfloor tq \rfloor - \lfloor sq \rfloor) \mathbb{E}[|S_1|], \end{aligned}$$

for every q large enough, by (4.103). On the other hand, using in addition that the variables $\{\bar{\zeta}_{y,q}\}_{y \in \mathbb{Z}}$ are mutually orthogonal, we write

$$(4.106) \quad \begin{aligned} \mathbb{E} \left[\left(\bar{T}_2^{(q)}(t) - \bar{T}_2^{(q)}(s) \right)^2 \right] &\leq q^{-(1+\alpha)/\alpha} \sum_{y \in \mathbb{Z}} \mathbb{E} \left[\left(\mathcal{N}_{[tq]}(y) - \mathcal{N}_{[sq]}(y) \right)^2 \right] \mathbb{E} \left[(\bar{\zeta}_{1,q})^2 \right] \\ &\leq q^{-(1+\alpha)/\alpha} \sum_{y \in \mathbb{Z}} \mathbb{E} \left[\left(\mathcal{N}_{[tq]-[sq]}(y) \right)^2 \right] \mathbb{E} \left[(\zeta_{1,q})^2 \right]. \end{aligned}$$

At this point observe that (4.33)-(4.36) and (4.42) are still valid if $\alpha = 2$. Hence, using the inequality $2ab \leq a^2 + b^2$ and (4.8) with $\beta = 2$, we have, for $n \rightarrow \infty$,

$$(4.107) \quad \begin{aligned} &\sum_{y \in \mathbb{Z}} \mathbb{E} \left[(\mathcal{N}_n(y))^2 \right] \\ &\leq 2 \sum_{y \in \mathbb{Z}} \mathbb{E} \left[\left(\sum_{r \in \mathbb{Z}} \mathbb{P}(|S_1| > r) N_n(y+r) \right)^2 \right] + O(n) \\ &\leq \sum_{y \in \mathbb{Z}} \sum_{r,s \in \mathbb{Z}} \mathbb{P}(|S_1| > r) \mathbb{P}(|S_1| > s) \mathbb{E} \left[(N_n(y+r))^2 + (N_n(y+s))^2 \right] + O(n) \\ &\leq 2\mu_\xi^2 \sum_{y \in \mathbb{Z}} \mathbb{E} \left[(N_n(y))^2 \right] + O(n) \\ &= O(n^{3/2}). \end{aligned}$$

Applying the above and (4.104) in (4.106) yields

$$(4.108) \quad \mathbb{E} \left[\left(\bar{T}_2^{(q)}(t) - \bar{T}_2^{(q)}(s) \right)^2 \right] \leq K_4 q^{-3/2} ([tq] - [sq])^{3/2},$$

for some $K_4 > 0$. Therefore, by Jensen's inequality,

$$(4.109) \quad \mathbb{E} \left[\left| \bar{T}_2^{(q)}(t) - \bar{T}_2^{(q)}(s) \right| \right] \leq \sqrt{K_4} q^{-3/4} ([tq] - [sq])^{3/4},$$

which, combined with (4.105), gives (4.99), concluding the proof of the tightness of $(\bar{T}^{(q)})_{q \in \mathbb{R}^+}$ and therefore of Lemma 3.1.

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