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# Structural validation and cross-cultural robustness of the European Cyberbullying Intervention Project Questionnaire

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## ABSTRACT

During the last decade, cyberbullying has become an increasing concern which has been addressed by diverse theoretical and methodological approaches. As a result there is a debate about its nature and rigorously validated assessment instruments have not yet been validated. In this context, in the present study an instrument composed of 22 items representing the different types of behaviours and actions that define cyberbullying has been structurally validated and its cross-cultural robustness has been calculated for the two main dimensions: cyber-victimization and cyber-aggression. To this end, 5679 secondary school students from six European countries (Spain, Germany, Italy, Poland, United Kingdom, and Greece) were surveyed through this self-report questionnaire which was designed based on previously existing instruments and the most relevant conceptual elements. Exploratory and confirmatory factor analyses were conducted and the global internal consistency was computed for the instrument and its two dimensions. Identical factor structures were found across all of the six subsamples. The results contribute to existing research by providing an instrument, the European Cyberbullying Intervention Project Questionnaire, which has been structurally validated in a wide sample from six different countries and that is useful to evaluate psycho-educative interventions against cyberbullying.

## 1. Introduction

In recent years, society has shown a growing interest in the phenomenon named cyberbullying frequently appearing in the online social relationships among youngsters and adolescents (Fenaughty & Harré, 2013). Nowadays, we are immersed in the process of elaborating a solid theoretical approximation and an agreed definition of the phenomenon (Berne et al., 2013;

Tokunaga, 2010). Thus, one of the main guides to follow is the research developed around traditional bullying (Olweus, 2013) as cyberbullying is defined as bullying developed through electronic media (Vivolo-Kantor, Martell, Holland, & Westby, 2014). Traditional bullying has been defined as physical, verbal, social and/or psychological aggression by a pupil against another, whom is chosen to be a victim of repeated attacks (Olweus, 1993, 1999). Such a negative and intentioned action puts the victim in a situation that is difficult to get out of. Bullying is neither an isolated aggression nor a simple individual behaviour but an interactive phenomenon in which several subjects are involved in at least three roles: bully, victim and bully-victim. Its distinctive characteristics are: the intentionality to hurt someone else, the imbalance of

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power between the aggressor and the victim and the repetition of the aggressive conducts by the aggressors over their victims. Such scientific evidences have clarified the nature of the bullying phenomenon and determined its standardization, hence, the appearance of instruments to measure it (Greif & Furlong, 2006). However, the nature of the electronic means that characterizes cyberbullying has made it necessary to investigate not only its conceptualization but also in order to provide instruments suitable to its nature with the aim of showing the levels of prevalence among adolescent population (Vivolo-Kantor et al., 2014).

## 2. Cyberbullying: Definition and characteristics

Behaviours such as verbal attacks through digital devices, publication and exhibition of embarrassing pictures, and the exclusion from online communication are some examples of how traditional bullying brings to life cyberbullying. Other behaviours such as virtual identity theft (i.e. to impersonate someone else or to hack personal accounts with the aim of obtaining personal information) are not included in traditional forms of bullying but are considered as cyberbullying (Perren et al., 2012). Instead, virtual behaviours accompanied by certain nuances such as the presence of adults take advantage of minors, or the intentionality of sexual nature to obtain embarrassing pictures (Smith, Thompson, & Davidson, 2014), are linked to other phenomena different from cyberbullying such as grooming or sexting (Den Hamer & Konijn, 2015).

Cyberbullying, by mainly referring to traditional bullying researches (Slonje & Smith, 2008), is defined as a clearly intentional aggression or hostile or harmful act carried out through an electronic device repeatedly over time by setting up an imbalance of powers between the aggressor and the victim (Tokunaga, 2010). Accordingly, both the aggressor and the victim are, a priori, substantial characters of the phenomenon, but there are also those that are aggressors and victims at the same time, the bully-victims (Yang & Salmivalli, 2013). In addition, there are researches that identifies cyberbullying exclusively with cyber-aggression (Calvete, Orue, Estévez, Villardón, & Padilla, 2010) or with cyber-victimization (Müller, Pfetsch, & Ittel, 2014), leaving out the dynamic existing between the roles and how the criteria of intentionality, repetition and imbalance of powers takes place between them (Olweus, 2013). Criteria that otherwise are not as evident in cyberbullying as they are in traditional bullying (Dehue, 2013; Slonje, Smith, & Frisén, 2013; Smith, Del Barrio, & Tokunaga, 2013). In this respect, some authors suggest that a single image or any other humiliating audio-visual material can be comparable to the repetition of traditional bullying, since the content can be perpetual on the Internet and is available for any person seeking access to it, or may even be downloaded and stored on personal devices (Heirman & Walrave, 2008), this suggests that digital aggression is equally harmful. Concerning the imbalance of power, high levels of technological knowledge and the difficulties that the victims may have in identifying the aggressors, can be interpreted as inferiority before the aggressor (Menesini & Nocentini, 2009; Vandebosch & Van Cleemput, 2008).

Despite these considerations, from our point of view, repetition shall be considered as a requirement for cyberbullying as for the victim or even for both the aggressor and the victim the experience is a repeated behaviour. In fact, certain existing qualitative studies have shown the need to maintain such a criterion (Nocentini et al., 2010). With regard to the imbalance of power, there is no doubt of its relevance in the dynamic of the phenomenon as the lack of competence for keeping personal data secure in digital scenarios imply or may imply that the victim faces inferiority with respect to the aggressor when communicating through digital devices (Vandebosch & Van Cleemput, 2008).

## 3. Prevalence of cyberbullying

There are currently more than 300 articles published about cyberbullying which offer figures about its prevalence (Arsène & Raynaud, 2014). However, there are differences in these results which make it difficult to know how many people are affected by this problem (Modecki, Minchin, Harbaugh, Guerra, & Runions, 2014). Thus, among the researches referring only to cyberaggression, figures range from 5.3% to 31.5% (Gradinger, Strohmeier, & Spiel, 2009; Pornari & Wood, 2010). Among those referring only to cyber-victimization figures oscillate between 2.2% to 56.2% (Perren, Dooley, Shaw, & Cross, 2010). These differences are also present in researches that analyze both cyber-aggression and cyber-victimization. For instance, a Greek study found 28.3% of cyber-victims of cyberbullying and 14.6% of perpetrators (Floros, Siomos, Fisoun, Dafouli, & Geroukalis, 2013). However, another in Sweden found that 5% were cyber-victims and 4% were cyberbullies (Läftman, Modin, & Östberg, 2013). The differences can be found even in studies developed with a population of similar characteristics and in the same regions or countries (Baek & Bullock, 2014).

Such a diversity in the rates of prevalence may have its origin, among others, in the plethora of perspectives from which cyberbullying is analyzed (Sabella, Patchin, & Hinduja, 2013) and therefore, in the great differences existing between the instruments used (Modecki et al., 2014), this diversity is inherited from traditional bullying measurements (Greif & Furlong, 2006).

## 4. Measures of cyberbullying

As mentioned above, cyberbullying involves a diversity of elements and behaviours making research and the development of valid assessment instruments more complicated (Ybarra, Mitchell, & Korchmaros, 2011). This difficulty has been partially overcome in recent studies which focus on the development of measurement instruments, but have exclusively dealt with only one of the dimensions so far: cyber-aggression (Calvete et al., 2010; Law, Shapka, Hymel, Olson & Waterhouse, 2012) or cyber-victimization (Tynes, Rose, & Williams, 2010), separately. Although this is a great contribution to existing research, studying only part of the phenomenon omits the dynamic nature of cyberbullying, and means it is not sufficient to test its complexity with the respective instruments (Dempsey, Sulkowski, Nichols, & Storch, 2009).

In the latest systematic review of measurement instruments of cyberbullying, Berne et al. (2013) have found 44 different instruments (until October 2010) which have been used to evaluate, measure or analyse this phenomenon. The vast majority of these instruments are self-reports focussing on different aspects of cyberbullying, with 56% of them assessing cyber-aggression or cyber-victimization separately (Berne et al., 2013). These two aspects are addressed together only in the Italian studies of Menesini, Nocentini, and Calussi (2011). However, the whole model did not fit and, therefore, the validation of the instrument had to be made separately for cyber-aggression and cyber-victimization. The 40% of remaining instruments focus on other elements such as the type of devices that can carry out an attack or on cyber-abuse (Law, Shapka, Domene, & Gagné, 2012).

In regards to specific psychometric properties, only 22% ( $N = 10$ ) of all the instruments analysed used statistical methods to empirically examine the underlying theory, such as confirmatory factor analysis (Berne et al., 2013). In most cases, these structural validations have been performed with samples of approximately 500 participants, with the exception of the study by Ybarra and Mitchell (2008) which was carried out with 1700 people through an online survey about cyber-victimization.

Moreover, these examples each come from one particular locality, region or country. There are no records of structural validations in which both dimensions (cyber-aggression and cyber-victimizations) are included, which is essential if it is assumed that cyberbullying is a dynamic between the aggressor and the victim. There is not, the cross-cultural robustness analysis in researches in which there has been collaboration from different countries (Berne et al., 2013; Byrne & Campbell, 1999). Therefore, we still face significant limitations to describe the cyberbullying phenomenon as no comparable figures of prevalence can be shown (Baek & Bullock, 2014). In this regard, it is still pending the task of creating instruments, validated with samples from different countries, which allow us to know if the phenomenon is more prevalent in certain countries. In addition, it is also pending and analysis of whether the dynamic between the cyber-aggressor and the cyber-victim become a part or cyberbullying or not.

## 5. Objective

In this respect, the goal of the present study was to look for the structural validation and the cross-cultural robustness of an instrument for cyberbullying that, considering the dynamic nature of the phenomenon, includes cyber-aggression and cyber-victimization behaviours. We understand that this is a modest step in the research of study and intervention against cyberbullying.

## 6. Materials and methods

### 6.1. The sample group

The whole sample was composed of 5679 students from six European countries: Poland [ $N=900$ ; 15.8% (52.3% girls), 8 schools], Spain [ $N=859$ ; 15.1% (52.3% girls), 3 schools], Italy [ $N=1,430$ ; 25.2% (56.3% girls), 15 schools], United Kingdom [ $N=737$ ; 13% (44.6% girls), 5 schools], Germany [ $N=846$ ; 14.9% (49% girls), 5 schools] and Greece [ $N=907$ ; 16% (54.2% girls), 13 schools]. 49% percent were girls and participants' age was between 11 and 23 years ( $M=14.77$ ;  $SD=1.58$ ; 7th to 10th grades). The selection of the sample was made through an incidental design by accessibility of those schools that accepted to participate in this study.

### 6.2. Measure

The European Cyberbullying Intervention Project Questionnaire (ECIPQ; Brighi et al., 2012) composed 22 Likert type items with five options for frequency ranging from *Never* to *more, 1 once or twice, 2 once a month, 3 once a week, 4 more times a week*. This questionnaire was designed on the basis of previous measures by Dooley, Pyżalski, and Cross (2009) and Pyżalski (2012) to evaluate the two dimensions of cyberbullying – victimization and aggression. It also tries to include the cyberbullying criteria of repetition and, although implicitly, the imbalance of power as it includes items of behaviours which imply a technological domain of the aggressor and a lack of the use of security measures by the victim. For instance, *Someone hacked into my account and pretended to be me*. The 22 items are given in the Appendix.

### 6.3. Procedure and data analysis

A purposive sampling procedure was used. The first version of the questionnaire was designed in English, translated into each of the five other languages, and a back translation checked against the original.

At first, descriptive analyses in the total sample and by participant countries were performed. Additionally, due to the ordinal nature of the variables rank-base comparative analyses were also performed with the aim of contrasting the differences among the subsamples.

A cross-validation process was performed by randomly dividing the sample into two same-size subsamples: S1 ( $n=2820$ ) and S2 ( $n=2859$ ). The first subsample was employed to identify the number of underlying factors; in this study the two factors of victimization and aggression. This was done by means of an exploratory factor analysis (EFA) carried out using FACTOR 9.2 (Lorenzo-Seva & Ferrando, 2006). This program was used to generate a matrix of polychoric inter-item correlations which were then subjected to EFA. The decision to compute polychoric correlations was based on recognition that the ECIPQ employs a polytomous response format and provides ordinal rather than interval-level measurement (Flora, Finkel, & Foshee, 2003). A polychoric correlation is a more appropriate measure of the relationship between two Likert-type items than Pearson correlation (Flora & Curran, 2004; Jöreskog, 1994; Olsson, 1979). This is because a set of Likert-type items does not have a multivariate-normal distribution and the normal-theory maximum likelihood method of factor estimation that is typically employed in EFA is not the best method of estimation for assessing the factor structure of such variables, even when applied to a matrix of polychoric correlations (Bollen, 1989; Browne & Cudeck, 1993). Unweighted least squares (ULS) served as the extraction technique because it is robust for use with data that are not normally distributed. Finally, the rotation was set to *direct oblimin* because total scores on the victimization and aggression subscales were inter-correlated.

The second subsample was used to verify the goodness of fit of the factor structure that emerged from the first exploratory analysis and to test for invariance across the different countries. Confirmatory factor analysis (CFA) was carried out using LISREL 9.1 (Jöreskog & Sörbom, 2012). As the data are ordinal and the correlations between factors are high, the Diagonally Weighted Least Squares (DWLS) method was used (Flora & Curran, 2004; Holgado-Tello, Chacón-Moscoso, Barbero-García, & Vila-Abad, 2010; Morata-Ramírez & Holgado-Tello, 2013). We followed the recommendations of Hu and Bentler (1999) to compare the suitability of the models proposed, using the comparative fit index (CFI), the goodness of fit index (GFI), and the non-normed fit index (NNFI), with values above .95 (Bentler & Bonett, 1980). The root mean square error of approximation (RMSEA) should be less than .05; SRMR values less than .08 were considered acceptable and those below .05 as adequate (Hu & Bentler, 1995). In order to investigate the stability of the confirmatory factor analysis under cross-validation, and to avoid splitting the available data, the expected cross-validation index (ECVI) was used (Browne & Cudeck, 1993). The ECVI measures the discrepancy between the fitted covariance matrix in the analysed sample and the expected covariance matrix that would be obtained in another sample of equivalent size. Application of the ECVI assumes a comparison of models, whereby the model having the smallest ECVI exhibits the greatest potential for replication (Byrne, 2006). The selection of these indices and their cut off values was based on their performance with large samples and DWLS as estimation method (Wu, West, & Taylor, 2009).

For calculating the prevalence of cyberbullying, the roles of behaviours participation and repetition were considered. Thus, *Cyber-victims* have been identified with scores equal or higher than 2 (*once a month*) in any of the items of cyber-victimization and with scores equal or lower than 1 (*once or twice*) in all of the items of aggression. In addition, *Cyber-aggressors* are those subjects with scores equal or higher than 2 (*once a month*) in any of the items of cyber-aggression and equal or lower than 1 (*once or twice*) in all of

the items of cyber-victimization. *Cyber-bully/Victim* have been identified those subjects with a score in any of the items of both cyber-aggression and cyber-victimization with a score equal or higher than 2 (*once a month*).

## 7. Results

### 7.1. Descriptive analysis

The Kruskal–Wallis test has been performed with the aim of comparing the differences among the participant countries with regard to the two dimensions of the questionnaire, cyber-aggression and cyber-victimization, and to the total questionnaire. Differences among participant countries were found (see [Table 1](#)).

### 7.2. Exploratory factor analysis

In the EFA with subsample S1 the variables violated the assumption of multivariate normality (see [Table 2](#)), Mardia's coefficient = 777.18. The Kaiser–Meyer–Olkin of Sampling Adequacy (KMO) value was .832 and Bartlett's Test of Sphericity  $\chi^2$ : 71170.0 ( $df = 231$ ;  $p < .001$ ). The factors were subjected to a direct oblimin rotation and the analysis revealed that the extraction of two factors was appropriate (see [Table 2](#)). The variance accounted for by the two factors was 65.4%. Also, Bentler's simplicity (S) index (1977) and the loading simplicity (LS) index ([Bentler, 1977](#); [Lorenzo-Seva, 2003](#)) were computed with values of .99 and .75, respectively, which means that each item was mainly related to only one dimension, and the overall solution showed high factor simplicity. The inter-factor correlation is .765.

### 7.3. Confirmatory factor analysis

According to the EFA results, a two-factor model could be an adequate approximation of the data. The CFA was conducted with the conditions outlined above because of the violation of the assumptions of normality and kurtosis with Mardia's coefficient = 629.17. The results of fitting the two-factor solution are suitable in sample S2 and in each subsample from the six countries in which the instrument was evaluated (see [Table 3](#), [Fig. 1](#)).

The results show that all models have CFI values greater than 0.95 and the ECVI index value does not exceed 0.6. Based on these low values of ECVI and adequacy of CFI values, the suitability of the model for different samples can be assumed. Both these indices are used to measure the comparative fit between two or more models, and the smaller the obtained values the better the fit ([Bandalos, 1993](#)). Our study, with the large sample employed and the need to compare more than three groups, is unsuited for the chi square difference test ([Browne & Cudeck, 1993](#); [Satorra & Bentler, 2001](#)). This test has the same limitations as the likelihood ratio test in general, so that very large samples lead to very trivial difference tests. As to the reliability of the instrument, the indices obtained (McDonald's Omega = .99 and Standardized Cronbach's alpha = .96) exhibit a suitable over all reliability and also adequate

**Table 2**  
Descriptive univariate analysis and factor loadings.

	M	SD	Skewness	Kurtosis	Factor 1	Factor 2
CybV1	.27	.645	3.305	13.362	.359	
CybV2	.23	.606	3.441	14.298	.498	
CybV3	.12	.471	5.347	34.466	.427	
CybV4	.13	.492	5.015	30.112	.408	
CybV5	.22	.489	3.051	14.493	.520	
CybV6	.08	.384	6.101	43.589	.849	
CybV7	.10	.448	5.646	36.979	.866	
CybV8	.13	.463	4.629	26.313	.698	
CybV9	.09	.398	6.318	47.837	.793	
CybV10	.11	.424	5.447	36.627	.687	
CybV11	.15	.509	4.704	26.398	.883	
Cybb1	.22	.623	3.934	17.955		.782
Cybb2	.17	.562	4.427	22.986		.800
Cybb3	.10	.492	6.096	40.367		.945
Cybb4	.09	.487	6.284	42.252		.910
Cybb5	.09	.471	6.499	46.054		.882
Cybb6	.10	.490	6.147	41.200		.828
Cybb7	.08	.463	6.788	48.921		.828
Cybb8	.09	.443	6.500	47.670		.819
Cybb9	.07	.408	7.486	62.431		.859
Cybb10	.22	.678	3.881	16.329		.787
Cybb11	.08	.443	6.746	49.208		.865

**Table 3**  
Model Fit subsample S2 and six countries.

	$\chi^2$	df	GFI	CFI	NNFI	RMSEA	SRMR	ECVI
Total Sample S2	1484.15	208	.986	.993	.993	.030	.080	.505
Poland	452.42	208	.992	.994	.993	.010	.058	.664
Spain	375.35	208	.973	.978	.975	.020	.055	.677
Italy	581.63	208	.966	.988	.987	.051	.087	.699
UK	525.71	208	.958	.952	.947	.010	.072	.684
Germany	389.59	208	.989	.996	.994	.047	.021	.600
Greece	395.56	208	.991	.989	.981	.013	.010	.654

reliability of each of the two factors making up the scale, the cyber-victimization factor with  $\alpha = .97$  and the aggression factor with  $\alpha = .93$ .

Finally, for calculating the prevalence of involvement in cyber-bullying, subjects were selected on the basis of the above-cited theoretical criterion. Such a selection was developed with the whole sample and with the subsamples comprising each of the participant countries (see [Table 4](#)).

## 8. Discussion

The study achieved the intended goal of the structural validation of an instrument on cyberbullying that integrates the traditional bullying characteristics, the dimensional nature of the phenomenon including both cyber-victimization and cyber-aggression bearing in mind the diversity of cyberbullying behaviours based on the concepts of [Olweus \(2012\)](#) and [Smith et al.](#)

**Table 1**  
Kruskal–Wallis test .

	Country	N	Mean rank		Mean Rank		Mean Rank		
Total sacale	Poland	900	2697.09	Cyber aggression	2906.57	Cyber Victim	2599.78		
	Spain	859	2509.55		2681.66		2528.01		
	Italy	1430	3183.37		3230.88		3072.31		
	UK	737	2334.24		2312.35		2549.22		
	Germany	846	2545.05		2830.72		2497.55		
	Greece	907	3439.48		2745.03		3563.27		
		$\chi^2 = 350.60$	$p = .00$		$\chi^2 = 247.21$	$p = .00$		$\chi^2 = 363.90$	$p = .00$

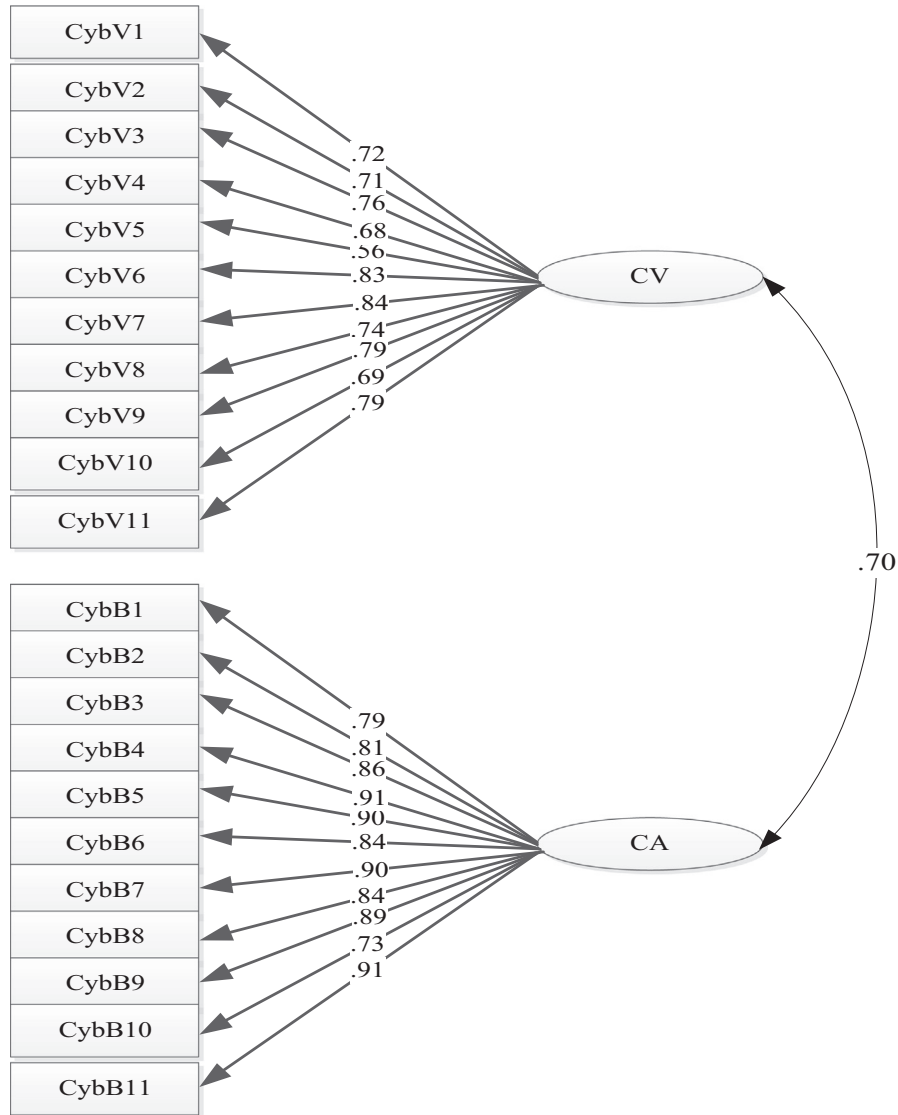


Fig. 1. AFC Graphic solution.

**Table 4**  
Prevalence of cyberbullying.

Country	N	Victims (N, %)	Aggressors (N, %)	Bully/Victims (N, %)	Total implication (N, %)
Poland	900	55 (6.11%)	61 (6.77%)	36 (4%)	152 (16.88%)
Spain	859	40 (4.65%)	44 (5.12%)	18 (2.09%)	102 (11.87%)
Italy	1430	115 (8.04%)	79 (5.52%)	75 (5.52%)	269 (18.81%)
UK	737	47 (6.37%)	7 (.94%)	15 (2.03%)	69 (9.36%)
Germany	846	35 (4.13%)	58 (6.85%)	26 (3.07%)	119 (14.06%)
Greece	907	92 (10.14%)	71 (7.82%)	56 (6.17%)	219 (24.14%)
Total Sample	5679	384 (6.76%)	320 (5.63%)	226 (3.97%)	930 (16.37%)

(2013). Results indicate that the structure of the European Cyberbullying Intervention Project Questionnaire (ECIPQ) is composed of two factors and the model provided good fit to the data and its scales have good internal consistency, exceeding the minimum limits of Cronbach index established. Thus, we could affirm that cyber-aggression and cyber-victimization become a part of cyberbullying not only from the theoretical point of view (Völlink, Bolman, Dehue, & Jacobs, 2013), but also with empirical

support. It has reinforced the conception of cyberbullying as a phenomenon of dynamic nature in which both the action of attacking and being attacked are relevant (Casas, Del Rey, & Ortega-Ruiz, 2013; Guckin, Cummins, & Lewis, 2010).

The different instruments which have been developed and used to measure or study cyberbullying previously are mostly limited only to cyber-aggression (Calvete et al., 2010) or only to cyber-victimization (Müller et al., 2014). Only one existing study by

Menesini et al. (2011), shared the bi-dimensional perspective. However cyber-aggression and cyber-victimization were considered independent patterns and the study was only carried out for an exclusively Italian sample. Therefore, this questionnaire could have an advantage: the approach to present very specific actions that clearly show the conceptual boundaries of the phenomenon of cyberbullying and, therefore, showing a possible involvement as cyber-victim, cyber-aggressor, or both.

In addition, the obtained psychometric properties show a good adjustment of the instrument in a sample consisting of participants from six European countries and on a country by country basis. This fact provides robustness to the consideration that the bi-dimensionality of cyberbullying is present in samples from different countries, which reinforces the nature of the phenomenon. This is especially relevant bearing in mind that each country has its own particular habits in the use of the information and communication technologies, with differences in the type of media and virtual social networking sites used (Drabowicz, 2012). The diversity in the extent of the phenomenon is shown in the descriptive results. The utilization of one common instrument not only points out that some behaviours appear more repeatedly than others in the different countries, but also that such differences would be explained by other factors that are not covered by this study. Accordingly, it has opened a new line of cross-cultural investigation that shall cope with other differences of psychosocial nature. As a consequence of this study, we know that, at least among the participant adolescents, the prevalence of cyberbullying in diverse contexts is not that different than is shown in the studies available so far (Modecki et al., 2014). This leads us to hypothesize that said heterogeneity may be owed, as announced when reviewing the literature, to the instruments used in previous investigations (Berne et al., 2013). Thus, this instrument turns out to be a useful and validated tool to continue progressing in the knowledge of the phenomenon, mainly in two directions. On the one hand, for assessing future comparative studies with population of the participant countries and, hence, analyzing the possible differences of prevalence and the variables associated to them. On the other hand, for evaluating the interventions that are being developed to cope with cyberbullying (Malti, Ribeaud, & Eisner, 2012) and, therefore, to develop knowledge about which practices are effective or not. As a consequence of the above, we could have evidence based practices that serve as a guarantee to adopt the necessary measures against this phenomenon which has devastating consequences for both children and adolescents and that can be even worse than traditional bullying (Gámez-Guadix, Orue, Smith, & Calvete, 2013; Pettalia, Levin, & Dickinson, 2013).

When interpreting the results of this study there are limitations to consider such as distortions and biases, for instance social desirability (Stone, Bachrach, Jobe, Kurtzman, & Cain, 1999). Also, although a large sample has participated in the study, the sub-samples may not be representative samples of the six countries which are included in the study. Nevertheless, one of the major contributions of the present study is the integration into a single measurement instrument across different countries, in which ICT use may differ, but where cyberbullying is conceived in a similar way. This contribution implies a modest advance in seeking evidence based practices against cyberbullying as it offers an instrument, which can be easily administered, and that could be used for contrasting the existing cyberbullying prior to and after the psycho-educative intervention, accordingly, it can also be used for evaluating the effectiveness of the actions.

Finally, more efforts should be dedicated to achieving the validation of the ECIPQ which in turn would be an important advance for the scientific community and, particularly, to the educational policies against cyberbullying.

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## Appendix A. Supplementary material

Supplementary data associated with this article can be found, in the online version, at <http://dx.doi.org/10.1016/j.chb.2015.03.065>.

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