Alma Mater Studiorum Università di Bologna Archivio istituzionale della ricerca

Global observables for RW: Law of large numbers

This is the final peer-reviewed author's accepted manuscript (postprint) of the following publication:

Published Version:

Dolgopyat D., Lenci M., Nandori P. (2021). Global observables for RW: Law of large numbers. ANNALES DE L'INSTITUT HENRI POINCARE-PROBABILITES ET STATISTIQUES, 57(1), 94-115 [10.1214/20-AIHP1072].

Availability:

This version is available at: https://hdl.handle.net/11585/820507 since: 2021-05-19

Published:

DOI: http://doi.org/10.1214/20-AIHP1072

Terms of use:

Some rights reserved. The terms and conditions for the reuse of this version of the manuscript are specified in the publishing policy. For all terms of use and more information see the publisher's website.

This item was downloaded from IRIS Università di Bologna (https://cris.unibo.it/). When citing, please refer to the published version.

(Article begins on next page)

This is the final peer-reviewed accepted manuscript of:

Dolgopyat, D., Lenci, M., & Nándori, P. (2021). Global observables for RW: Law of large numbers. *Annales De l'Institut Henri Poincare (B) Probability and Statistics*, 57(1), 94-115

The final published version is available online at https://dx.doi.org/10.1214/20-AIHP1072

Terms of use:

Some rights reserved. The terms and conditions for the reuse of this version of the manuscript are specified in the publishing policy. For all terms of use and more information see the publisher's website.

This item was downloaded from IRIS Università di Bologna (https://cris.unibo.it/)

When citing, please refer to the published version.

GLOBAL OBSERVABLES FOR RANDOM WALKS: LAW OF LARGE NUMBERS

DMITRY DOLGOPYAT, MARCO LENCI, AND PÉTER NÁNDORI

ABSTRACT. We consider the sums $T_N = \sum_{n=1}^N F(S_n)$ where S_n is a random walk on

 \mathbb{Z}^d and $F: \mathbb{Z}^d \to \mathbb{R}$ is a global observable, that is, a bounded function which admits an average value when averaged over large cubes. We show that T_N always satisfies the weak Law of Large Numbers but the strong law fails in general except for one dimensional walks with drift. Under additional regularity assumptions on F, we obtain the Strong Law of Large Numbers and estimate the rate of convergence. The growth exponents which we obtain turn out to be optimal in two special cases: for quasiperiodic observables and for random walks in random scenery.

1. Introduction

1.1. **Motivation.** Ergodic theory was created in the beginning of the last century motivated by the needs of homogenization (more specifically the quest to justify the kinetic equations of statistical mechanics). By now ergodic theory is a flourishing subject. Namely, ergodic theorems are established under very general conditions and ergodic properties of a large number of smooth systems are known (see e.g. [20]). Moreover, ergodicity turns out to be useful in the questions of averaging and homogenization (see e.g. [19, 29, 31, 34]). However, many dynamical systems appearing in applications preserve infinite invariant measure and ergodic theory of infinite-measure-preserving systems is much less developed. In fact, most of the work in infinite ergodic theory (see e.g. [1]) deals with local (L^1) observables while from physical point of view it is more natural to consider extensive observables ([22, 33]) which admit an infinite-volume average. One explanation for this is that while for local observables ergodic theorems can be obtained with minimal regularity assumptions on the observable, this is not the case for global observables as the present paper shows. The study of ergodic properties of infinite measure transformations with respect to extensive functions started relatively recently [24]. In particular, mixing properties of several systems with respect to global observables were obtained in [6, 7, 15, 24, 25]. A natural question is thus to investigate the law of large numbers for global observables. A first step in this direction was recently taken in [28]. In this paper we carry out a detailed analysis in the simplest possible setting: random walks on \mathbb{Z}^d . Our goal is to ascertain the correct spaces for the law of large numbers in various cases.

²⁰⁰⁰ Mathematics Subject Classification. Primary 60F15, 60G50; Secondary 37A40, 60K37.

Key words and phrases. Random walks, law of large numbers, arcsine law, random walks in random scenery, global observables.

1.2. **Results.** Let $X_1, X_2, ...$ be an iid sequence of \mathbb{Z}^d valued random variables. Let $S_0 = 0$ and $S_N = \sum_{n=1}^N X_n$ be the corresponding random walk. We assume that

- (1) (non-degeneracy) the smallest group supporting the range of X_1 is \mathbb{Z}^d ,
- (2) (aperiodicity) $g.c.d.\{n > 0 : \mathbb{P}(S_n = 0) > 0\} = 1.$

We will also assume that S_n is in the normal domain of attraction of a stable law with some index α . That is, there is a non-degenerate d dimensional random variable Y such that

$$\frac{S_n}{n^{1/\alpha}} \Rightarrow Y \text{ if } \alpha \in (0,1) \text{ and } \frac{S_n - n\mathbb{E}(X_1)}{n^{1/\alpha}} \Rightarrow Y \text{ if } \alpha \in (1,2].$$

To avoid uninteresting minor technical difficulties, we will mostly assume that $\alpha \neq 1$.

We define several function spaces which proved to be useful in the previous studies of global observables [15]. Without further notice, we always assume that all functions are bounded.

Given a non-empty subset $V \subset \mathbb{Z}^d$ and $F \in L^{\infty}(\mathbb{Z}^d, \mathbb{R})$, we write

$$\bar{F}_V = \frac{1}{|V|} \sum_{v \in V} F(v).$$

Given $(a_1, b_1, \dots a_d, b_d)$ with $a_i < b_i$, for $j = 1, \dots d$, let

$$(1.1) V(a_1, b_1, \dots a_d, b_d) = \{x \in \mathbb{Z}^d : x_j \in [a_j, b_j] \text{ for } j = 1, \dots d\}.$$

Let \mathbf{G}_+ be the space of bounded functions on \mathbb{Z} such that the limit $\bar{F}_+ = \lim_{v \to \infty} \bar{F}_{[0,v]}$ exists and \mathbf{G}_- be the space of bounded functions on \mathbb{Z} such that the limit $\bar{F}_- = \lim_{v \to \infty} \bar{F}_{[-v,0]}$ exists. Set $\mathbf{G}_{\pm} = \mathbf{G}_+ \cap \mathbf{G}_-$. Define

$$\mathbf{G}_0 = \{ F \in L^{\infty}(\mathbb{Z}^d, \mathbb{R}) : \exists \bar{F} \ \forall a_1 < b_1, \dots a_d < b_d \ \lim_{L \to \infty} \bar{F}_{V(a_1 L, b_1 L, \dots, a_d L, b_d L)} = \bar{F} \}.$$

Note that in dimension 1, $\mathbf{G}_0 = \{ F \in \mathbf{G}_{\pm} : \bar{F}_+ = \bar{F}_- \}$. Let \mathbf{G}_U be the space of functions such that for each ε there is L such that for all cubes V with side larger than L we have

$$(1.2) |\bar{F}_V - \bar{F}| \le \varepsilon.$$

Let \mathbf{G}_{γ} be the set of functions where (1.2) only holds if the center of V is within distance L^{γ} of the origin. Thus $\mathbf{G}_{U} \subset \mathbf{G}_{\gamma}$. Also, $\mathbf{G}_{\gamma} \subset \mathbf{G}_{0}$ if $\gamma > 1$. Finally, let

$$\mathbf{G}_{\gamma}^{\beta} = \{ F \in L^{\infty}(\mathbb{Z}^{d}, \mathbb{R}) : \exists \bar{F} \ \forall a_{1} < b_{1}, ..., a_{d} < b_{d} \ \exists C : \forall L, \forall z \in \mathbb{Z}^{d}, |z| < L^{\gamma}, |\bar{F}_{z+V(a_{1}L,b_{1}L,...,a_{d}L,b_{d}L)} - \bar{F}| < CL^{d(\beta-1)} \}.$$

Clearly,
$$\mathbf{G}_{\gamma}^{\beta} \subset \mathbf{G}_{\gamma}$$
 for any $\beta < 1$. Also, let $\mathbf{G}_{\infty}^{\beta} = \bigcap_{\gamma > 0} \mathbf{G}_{\gamma}^{\beta}$.

Our goal is to study Birkhoff sums

$$T_N = \sum_{n=1}^N F(S_n).$$

In particular we would like to know if $\frac{T_N}{N}$ converges to \bar{F} for F in each of the spaces \mathbf{G}_* introduced above.

Our results could be summarized as follows.

Theorem 1.1. Suppose that $\mathbb{E}(X) = 0$ and that S_N is in the normal domain of attraction of a stable law of some index $\alpha > 1$. Then for all $F \in \mathbf{G}_0$, $\frac{T_N}{N} \Rightarrow \bar{F}$ in law as $N \to \infty$.

Theorem 1.2. Suppose that d=1, $\mathbb{E}(X)=0$ and $V(X)<\infty$. Then for all $F\in \mathbf{G}_{\pm}$, $\frac{T_N}{N}$ converges in law as $N\to\infty$. In particular, if $\bar{F}_-=0$ and $\bar{F}_+=1$ then the limiting law has arcsine distribution: for $z\in[0,1]$

$$\lim_{N \to \infty} \mathbb{P}\left(\frac{T_N}{N} \le z\right) = \frac{2}{\pi} \arcsin\sqrt{z}.$$

Note that Theorem 1.2 is a simple homogenization result: it says that the limit distribution of $\frac{T_N}{N}$ remains the same if the oscillatory function F is replaced by a more regular function $\bar{F}_{-}1_{x<0} + \bar{F}_{+}1_{x\geq0}$ (see [27]). This confirms the usefulness of global observables in applications.

Theorem 1.3. Suppose that S_N is in the normal domain of attraction of the stable law of some index α . Suppose that either

- (i) $1 < \alpha \le 2$, $\mathbb{E}(X_1) \ne 0$ and $\gamma > 1$ or
- (ii) $1 < \alpha \le 2$, $\mathbb{E}(X_1) = 0$ and $\gamma > 1/\alpha$ or
- (iii) $\alpha \leq 1$ and $\gamma > 1/\alpha$.

Then, for all $F \in \mathbf{G}_{\gamma}$, $\frac{T_N}{N} \to \bar{F}$ almost surely.

Theorem 1.4. Suppose $E(X_1) = 0$ and $E(|X_1|^k) < \infty$ for all $k \in \mathbb{N}$. For $d \in \mathbb{N}$, let

$$\rho_d(\beta) := \begin{cases} \frac{1}{2} & \text{if } \beta \leq \frac{d-1}{d} \\ \frac{d}{2}(\beta - 1) + 1 & \text{if } \beta > \frac{d-1}{d}, \end{cases} \qquad \gamma(d, \beta, \varepsilon) := \begin{cases} \frac{2}{\beta} & \text{if } d = 1 \\ \frac{1}{\varepsilon} & \text{if } d \geq 2. \end{cases}$$

Then for every $d \in \mathbb{N}$, for every $\beta \in [0,1)$ every $\varepsilon > 0$ and any $F \in \mathbf{G}_{\gamma(d,\beta,\varepsilon)}^{\beta}$ with $\overline{F} = 0$, we have

$$\frac{T_N}{N^{\rho_d(\beta)+\varepsilon}} \to 0 \text{ almost surely as } N \to \infty.$$

Corollary 1.5. If d = 1 and $F \in \mathbf{G}_{2/\beta}^{\beta}$ or $d \geq 2$ and $F \in \mathbf{G}_{\infty}^{\beta}$ then with probability 1, for all ε

$$\lim_{N\to\infty}\frac{T_N}{N^{\rho_d(\beta)+\varepsilon}}=0.$$

Remark 1.6. Let us discuss two special cases.

(A) (Random walk in random scenery) If F(x), $x \in \mathbb{Z}^d$ are bounded and iid with expectation 0, then by moderate deviation estimates, almost surely for every $\gamma < \infty$ and for every $\varepsilon > 0$, $F \in \mathbf{G}_{\gamma}^{\frac{1}{2}+\varepsilon}$ and $\bar{F} = 0$.

Now assuming that the random walker has zero expectation and finite moments of every order, Theorem 1.4 implies $\frac{T_N}{N^{\frac{3}{4}+\varepsilon}} \to 0$ almost surely in dimension d=1. Note

that in this case, $\frac{T_N}{N^{\frac{3}{4}}}$ has a non-trivial weak limit by [21]. If $d \geq 2$, Theorem 1.4 gives $\frac{T_N}{N^{\frac{1}{2}+\varepsilon}} \to 0$ almost surely while $\frac{T_N}{N^{\frac{1}{2}}}$ ($\frac{T_N}{\sqrt{N \ln N}}$ if d=2) has a non-trivial weak limit ([21, 8, 4]). We note that Theorem 1.4 is not new for F as above (see [23, 18]) however, we would like to emphasize that our space $\mathbf{G}_{\gamma}^{\frac{1}{2}+\varepsilon}$ includes many more functions than just realizations of iid process, so both the result and the proof of Theorem 1.4 are new even for $\mathbf{G}_{\gamma}^{\frac{1}{2}+\varepsilon}$.

(B) If F(x) is periodic, $\bar{F} = 0$, then $F \in \mathbf{G}_{\infty}^{\frac{d-1}{d}}$. Thus, assuming that the random walker has zero expectation and finite moments of every order, Corollary 1.5 implies

$$\frac{T_N}{N^{\frac{1}{2}+\varepsilon}} \to 0$$

almost surely for all d. Note that by the central limit theorem for finite Markov chains, $\frac{T_N}{\sqrt{N}}$ has a Gaussian weak limit.

In fact, our results also give (1.3) for quasi-periodic observables. That is, given $d \in \mathbb{N}$ and a C^{∞} function $\mathfrak{F}: \mathbb{T}^d \to \mathbb{R}$, let $\hat{\mathfrak{F}}: \mathbb{R}^d \to \mathbb{R}$ be the $[0,1]^d$ -periodic extension of \mathfrak{F} . Furthermore, given d vectors $\alpha_{(1)}, \ldots \alpha_{(d)} \in \mathbb{R}^d$ and an initial phase $\omega \in [0,1]^d$, let

$$F(x) = \hat{\mathfrak{F}}\left(\omega + \sum_{j=1}^{d} x_j \alpha_{(j)}\right)$$

where $(x_1, \ldots x_d)$ are coordinates of vector $x \in \mathbb{Z}^d$. We say that a vector $\alpha \in \mathbb{Z}^d$ is Diophantine, if there are constants K and σ such that for each $m \in \mathbb{Z}^d$,

$$\left| e^{2\pi\langle m,\alpha\rangle} - 1 \right| \ge \frac{K}{|m|^{\sigma}}.$$

If $\alpha_{(j)}$ is Diophantine for all j=1,...,d, then $F \in \mathbf{G}_{\infty}^{0}$ (see e.g. [20, §2.9]) so (1.3) holds.

Thus in both cases (A) and (B) our results give an optimal exponent for the growth rate of T_N .

Remark 1.7. Periodic (and quasi-periodic) observables are special case of stationary ergodic observables. More precisely, let $\mathfrak{T}_1, \ldots \mathfrak{T}_d$ be commuting measurable maps of a space Ω preserving a probability measure ν . Given a bounded measurable function \mathfrak{F} on Ω and an initial condition $\omega \in \Omega$, define

$$(1.4) F_{\omega}(k) = \mathfrak{F}(\mathfrak{T}^k \omega),$$

where for $k = (k_1, \dots k_d) \in \mathbb{Z}^d$ we let $\mathfrak{T}^k = \mathfrak{T}_d^{k_d} \dots \mathfrak{T}_1^{k_1}$. If the family \mathfrak{T}^k is ergodic, then the ergodic theorem tells us that for almost all ω , $F_{\omega} \in \mathbf{G}_0$ and $\bar{F} = \nu(\mathfrak{F})$. For the observables given by (1.4) the strong law of large numbers for almost every ω follows from ergodicity of the environment viewed by the particle process ([5]). Theorem 1.1 only gives a weak law of large numbers, (except in dimension 1 in the ballistic case, see Theorem 1.8 below). On the other hand our result gives valuable additional information even for stationary ergodic environments. Namely, the set of full measure where the weak law of large numbers holds contains all environments where ergodic averages of \mathfrak{F}

exist. We also note that Theorem 1.4 provides new and non-trivial information even in the stationary ergodic case.

Theorem 1.8. Suppose that $d=1, v=\mathbb{E}(X_1)>0$ and for all $t\geq 1, \mathbb{P}(|X_1|>t)\leq C/t^{\beta}$ for some C>0 and $\beta>1$. If $F\in \mathbf{G}_+$, then $\frac{T_N}{N}\to \bar{F}_+$ almost surely.

The next theorem shows that in general the strong law of large numbers fails in G_0 .

Theorem 1.9. Suppose that S_N is in the normal domain of attraction of the stable law of some index α . Moreover assume that one of the following assumptions is satisfied

- (a) $\alpha > 1$ and $\mathbb{E}(X_1) = 0$; or
- (b) $\alpha < 1$.

Then there exists $F \in \mathbf{G}_0$ such that, with probability 1, $\frac{T_N}{N}$ does not converge as $N \to \infty$.

Remark 1.10. The same conclusion holds in case (a) even if $\mathbb{E}(X_1) \neq 0$. However, in this case \mathbf{G}_0 is not an appropriate space to look at since we even do not have a weak law of large numbers in \mathbf{G}_0 .

2. Weak convergence.

Here we prove Theorems 1.1 and 1.2.

2.1. **Preliminaries.** First, we recall two useful results.

Theorem 2.1. ([17, Section 50]) Under the assumptions of Theorem 1.1, S_n satisfies the local limit theorem, i.e. there is a continuous probability density g such that

$$\lim_{n \to \infty} \sup_{l \in \mathbb{Z}^d} |n^{d/\alpha} \mathbb{P}(S_n = l) - g(l/n^{1/\alpha})| = 0.$$

Theorem 2.2 (Local global mixing, [15]). Under the assumptions of Theorem 1.1, S_n is local global mixing, i.e.

$$\lim_{n\to\infty} \mathbb{E}(F(S_n)) = \bar{F}.$$

2.2. **Proof of Theorem 1.1.** Replacing F by $F - \bar{F}$, we can assume that $\bar{F} = 0$. By Theorem 2.2, we have $\lim_{N \to \infty} \frac{\mathbb{E}(T_N)}{N} = 0$. Thus in order to prove Theorem 1.1, it suffices to verify that $\lim_{N \to \infty} \frac{\mathbb{E}(T_N^2)}{N^2} = 0$. Let us fix some $\varepsilon > 0$ and prove that $\mathbb{E}(T_N^2) < \varepsilon N^2$ for all sufficiently large N. We have

$$\mathbb{E}(T_N^2) = 2 \sum_{0 \le n_1 < n_2 \le N} \mathbb{E}(F(S_{n_1})F(S_{n_2})) + \sum_{n=1}^N \mathbb{E}(F^2(S_n)).$$

Now writing $\varepsilon_1 = \frac{\varepsilon}{50||F||_{\infty}^2}$, we have

$$\mathbb{E}(T_N^2) \le \frac{\varepsilon}{10} N^2 + \left| 2 \sum_{\varepsilon_1 N < n_1 < n_1 + \varepsilon_1 N < n_2 \le N} \mathbb{E}(F(S_{n_1}) F(S_{n_2})) \right|.$$

Choose a constant K such that $\mathbb{P}(|S_N| > KN^{1/\alpha}/2) < \frac{\varepsilon}{20||F||_{\infty}^2}$ for all sufficiently large N. Thus we have

$$\mathbb{E}(T_N^2) \le \frac{2\varepsilon}{10} N^2 + \left| 2 \sum_{\varepsilon_1 N < n_1 < n_1 + \varepsilon_1 N < n_2 \le N} \mathbb{E}(1_{\{|S_{n_1}|, |S_{n_2}| < KN^{1/\alpha}\}} F(S_{n_1}) F(S_{n_2})) \right|.$$

By the Markov property, we have

$$|\mathbb{E}(1_{\{|S_{n_1}|,|S_{n_2}|< KN^{1/\alpha}\}}F(S_{n_1})F(S_{n_2}))| \le \sum_{|x|< KN^{1/\alpha}} ||F||_{\infty} \mathbb{P}(S_{n_1} = x)|\mathbb{E}(1_{\{|S_{n_2-n_1}-x|< KN^{1/\alpha}\}}F(S_{n_2-n_1} - x))|.$$

Thus it is sufficient to prove that for every x with $|x| < KN^{1/\alpha}$.

(2.1)
$$|\mathbb{E}(1_{\{|S_{n_2-n_1}-x| < KN^{1/\alpha}\}} F(S_{n_2-n_1}-x))| \le \frac{\varepsilon}{5||F||_{\infty}}.$$

Recall that g is the density function of the limiting distribution of $S_n/n^{1/\alpha}$ as in Theorem 2.1. Now we choose δ so that the oscillation of g on any cube of side length $\delta/\varepsilon_1^{1/\alpha}$ within distance $2K/\varepsilon_1^{1/\alpha}$ from the origin is less than

(2.2)
$$\eta := \frac{\varepsilon \varepsilon_1^{d/\alpha}}{20K^d \|F\|_{\infty}^2}.$$

Let us partition the box $[-K,K]^d \subset \mathbb{R}^d$ into boxes of side length δ . Denote these boxes by $B_k, k = 1, ..., \bar{k}$ and let $B_{k,M} = MB_k \cap \mathbb{Z}^d$ for positive integers M. Since, $F \in \mathbf{G}_0$ we obtain from the definition of \mathbf{G}_0 that

$$\lim_{M \to \infty} \bar{F}_{B_{k,M}} = 0.$$

Note that the convergence in (2.3) is uniform in k as there is a finite number of k's.

Now let $m = n_2 - n_1$ and $M = N^{1/\alpha}$. Then $m^{1/\alpha} > \varepsilon_1^{1/\alpha} M$. Recalling (2.1), for every $x \in \mathbb{Z}^d$ with |x| < KM, we have

$$(2.4) |\mathbb{E}(1_{\{|S_m+x|< KM\}}F(S_m+x))| \le \sum_{k=1}^{\bar{k}} |\mathbb{E}(1_{\{S_m+x\in B_{k,M}\}}F(S_m+x))|$$

Now by Theorem 2.1 and the choice of δ , we have for all $x \in \mathbb{Z}^d$ with |x| < KM

(2.5)
$$|\mathbb{E}(1_{\{S_m+x\in B_{k,M}\}}F(S_m+x))| = |\sum_{y\in B_{k,M}} (p_{k,x,m}+e_{x,k,m,y})m^{-d/\alpha}F(y)|,$$

where $p_{k,x,m} = g\left(\frac{z_k M - x}{m^{1/\alpha}}\right)$, z_k is the center of B_k and for m sufficiently large $e_{x,k,m,y} < 2\eta$ (where η is given by (2.2)) uniformly in x, k, y as above. Consequently,

$$\sum_{k} \sum_{y} e_{x,k,m,y} m^{-d/\alpha} F(y) \le m^{-d/\alpha} (KN^{1/\alpha})^d 2\eta \|F\|_{\infty} \le \frac{\varepsilon}{10 \|F\|_{\infty}}$$

for sufficiently large m (that is for sufficiently large N). Thus dropping $e_{x,k,m,y}$ from the right hand side of (2.5) gives a negligible error. The remaining term is $p_{k,x,m}m^{-d/\alpha}\sum_{y\in B_k}F(y)$,

which when summed over k, is small by (2.3). Thus the absolute value of (2.4) is smaller than $\varepsilon/(5||F||_{\infty})$ for N sufficiently large, which completes the proof of (2.1). Theorem 1.1 follows.

2.3. **Proof of Theorem 1.2.** We prove the second statement. The first one is a trivial corollary. Indeed, given $F \in \mathbf{G}_{\pm}$ with $\bar{F}_{-} = \bar{F}_{+}$, the convergence follows from Theorem

1.1. On the other hand if $\bar{F}_- \neq \bar{F}_+$, then we can consider $\tilde{F}(x) = \frac{F(x) - \bar{F}_-}{\bar{F}_+ - \bar{F}_-}$ and note

that
$$\bar{\tilde{F}}_- = 0$$
, $\bar{\tilde{F}}_+ = 1$ and $\tilde{T}_N = \sum_{n=1}^N \tilde{F}(S_n) = \frac{T_N - N\bar{F}_-}{\bar{F}_+ - \bar{F}_-}$, whereby one derives the limit distribution of T_N .

Thus we assume $\bar{F}_{-}=0$, $\bar{F}_{+}=1$. Denote H(x)=1 if x>0, otherwise H(x)=0. Let us also write G(x)=F(x)-H(x).

Then $\bar{G}_{-} = \bar{G}_{+} = 0$ and so $G \in \mathbf{G}_{0}$. Decompose

$$T_N = \frac{1}{N} \sum_{n=1}^{N} G(S_n) + \frac{1}{N} \sum_{n=1}^{N} H(S_n).$$

The first sum on the RHS converges to zero by Theorem 1.1 and the second sum on the RHS converges weakly to the arcsine law by classical theory [16, §XII.8]. Then the LHS also converges weakly to the arcsine law by Slutsky's theorem.

3. SLLN IN
$$\mathbf{G}_{\gamma}$$
.

Here we prove Theorem 1.3.

Fix $F \in \mathbf{G}_{\gamma}$. As before, we can assume w.l.o.g. that F = 0. Since $F \in \mathbf{G}_{\gamma}$, the proof of Theorem 2.3 of [15] shows that, for any $\eta \in (0, 1)$,

(3.1)
$$\lim_{k \to \infty} \sup_{|x| < k^{\eta \gamma}} \left| \frac{\mathbb{E}_x(T_k)}{k} \right| = 0,$$

where \mathbb{E}_x denotes the expectation in the case where $S_0 = x$.

Lemma 3.1. Suppose that $1 < \gamma_1$ in case (i) or $1/\alpha < \gamma_1$ in cases (ii) and (iii). Then with probability one we have that, for large N,

$$\max_{0 \le k \le N} |S_k| \le N^{\gamma_1}.$$

Proof. In case (i) the statement follows from the Law of Large Numbers, so we only need to consider cases (ii) and (iii). We have for any $\varepsilon > 0$ that $|S_N| > N^{1/\alpha + \varepsilon}$ holds only finitely many times almost surely by [30] in case (ii) and by [26] in case (iii).

Choose $\gamma_1 < \gamma$ as in Lemma 3.1 and $\eta < 1$ such that $\gamma_1 < \gamma \eta$. For $j = 0, 1, \ldots, \lfloor N^{1-\eta} \rfloor$, set $\tilde{T}_j := \tilde{T}_{N,j} := N^{-\eta} T_{\lfloor j N^{\eta} \rfloor}$ (with the convention $T_0 \equiv 0$) and denote by $\tilde{\mathcal{F}}_j := \tilde{\mathcal{F}}_{N,j}$ the σ -algebra generated by $\{S_k\}_{k=0}^{\lfloor j N^{\eta} \rfloor}$. Denote $\mathcal{A}_{j,N} = \{|S_{\lfloor j N^{\eta} \rfloor}| \leq N^{\gamma_1}\}$. Fix $\varepsilon > 0$. We claim that there exists $N_0 = N_0(\varepsilon)$ such that, for all $N \geq N_0$ and $j < \lfloor N^{1-\eta} \rfloor$,

(3.2)
$$\left| \mathbb{E} \left(1_{\mathcal{A}_{j,N}} \left(\tilde{T}_{j+1} - \tilde{T}_{j} \right) \middle| \tilde{\mathcal{F}}_{j} \right) \right| \leq \varepsilon.$$

Indeed if $\mathcal{A}_{j,N}$ occurs then (3.2) holds due to (3.1), otherwise it holds since the LHS is zero.

Setting

$$(3.3) Y_j := 1_{\mathcal{A}_{j,N}} \left(\tilde{T}_{j+1} - \tilde{T}_j \right) - D_j$$

where

(3.4)
$$D_j := \mathbb{E}\left(1_{\mathcal{A}_{j,N}}\left(\tilde{T}_{j+1} - \tilde{T}_j\right) \middle| \tilde{\mathcal{F}}_j\right)$$

defines a martingale difference, w.r.t. $\{\tilde{\mathcal{F}}_j\}$, with $|Y_j| \leq ||F||_{\infty} + \varepsilon$. Applying Azuma's inequality we get that, for all $\delta > 0$,

$$\mathbb{P}\left(\left|\sum_{j=0}^{N^{1-\eta}-1} Y_j\right| \ge \delta N^{1-\eta}\right) \le 2 \exp\left(-\frac{\delta^2 N^{1-\eta}}{2(\|F\|_{\infty} + \varepsilon)^2}\right).$$

Therefore, by Borel-Cantelli,

$$\limsup_{N \to \infty} \left. \frac{1}{N^{1-\eta}} \middle| \sum_{j=0}^{N^{1-\eta}-1} Y_j \right| \le \delta \quad \text{a.s.}$$

Since δ is arbitrary, with probability one we have

(3.5)
$$\lim_{N \to \infty} \frac{1}{N^{1-\eta}} \sum_{j=0}^{N^{1-\eta}-1} Y_j = 0.$$

On the other hand, definitions (3.3)–(3.4) and Lemma 3.1 show that, with probability one, for all large N depending on the realization of the walk,

(3.6)
$$T_N = T_{\lfloor N^{1-\eta} \rfloor N^{\eta}} + O(N^{\eta}) = N^{\eta} \left(\sum_{j=0}^{N^{1-\eta}-1} Y_j + \sum_{j=0}^{N^{1-\eta}-1} D_j \right) + O(N^{\eta}).$$

In view of (3.2), (3.5), and (3.6) we have:

$$\limsup_{N \to \infty} \left| \frac{T_N}{N} \right| = \limsup_{N \to \infty} \left| \frac{1}{N^{1-\eta}} \right| \sum_{j=0}^{N^{1-\eta}-1} (Y_j + D_j) \right|$$

$$= \limsup_{N \to \infty} \frac{1}{N^{1-\eta}} \left| \sum_{j=0}^{N^{1-\eta}-1} D_j \right| \le \varepsilon \quad \text{a.s.}$$

Since ε is arbitrary, $\lim_{N\to\infty}\frac{T_N}{N}=0$ almost surely.

4. Speed of convergence in $\mathbf{G}_{\gamma}^{\beta}$.

Here we prove Theorem 1.4.

Note that $\mathbf{G}_{\gamma_1}^{\beta_1} \subset \mathbf{G}_{\gamma_2}^{\beta_2}$ whenever $\beta_1 \leq \beta_2$ and $\gamma_2 \leq \gamma_1$. Since $\rho_d(\beta)$ is constant for $\beta \in [0, (d-1)/d]$ and is continuous at (d-1)/d, it is sufficient to prove the theorem for

$$\beta > \frac{d-1}{d}.$$

Let $\mathbb{P}_x(.) = \mathbb{P}(.|S_0 = x), \mathbb{E}_x(.) = \mathbb{E}(.|S_0 = x).$

We start with the following

Proposition 4.1. Under the conditions of Theorem 1.4, for every $d \in \mathbb{N}$, every $\beta \in ((d-1)/d, 1)$ and every $\varepsilon > 0$ there exists some $\delta > 0$ so that

(4.2)
$$\sup_{x_0:|x_0| \le N^{1/2+\delta}} \mathbb{E}_{x_0}(T_N^2) < CN^{2\rho_d(\beta) + 2\varepsilon}.$$

Note that Proposition 4.1 combined with Chebyshev's inequality implies that

$$\frac{T_N}{N^{\rho_d(\beta)+\varepsilon}} \Rightarrow 0 \text{ in law as } N \to \infty.$$

Section 4 is divided into three parts. In §4.1 we derive Theorem 1.4 from Proposition 4.1. In §4.2 we prove Proposition 4.1 for d=1. In §4.3, we extend the proof of Proposition 4.1 to arbitrary dimension d.

4.1. **Proof of Theorem 1.4.** Here, we derive the theorem from Proposition 4.1. For simplicity we write $\rho = \rho_d(\beta)$.

We will show that

$$(4.3) \mathbb{P}(\exists n \le N : |T_n| > 2N^{\rho + \varepsilon/2}) \le CN^{-\varepsilon/2}$$

If (4.3) holds, then writing $N_k = 2^k$, we find

$$\mathbb{P}(\exists n = N_{k-1}, ...N_k : |T_n| > 2N_k^{\rho + \varepsilon/2}) \le CN_k^{-\varepsilon/2}.$$

and the theorem follows from Borel Cantelli lemma. To prove (4.3), let us write

$$\tau_N = \min\{\min\{n : |T_n| > 2N^{\rho + \varepsilon/2}\}, N\}.$$

Then

$$\mathbb{P}(\exists n \leq N : |T_n| > 2N^{\rho + \varepsilon/2})$$

$$\leq \mathbb{P}(|T_N| > N^{\rho + \varepsilon/2}) + \mathbb{P}(|T_{\tau_N}| > 2N^{\rho + \varepsilon/2}, |T_N| \leq N^{\rho + \varepsilon/2}) =: p_1 + p_2.$$

By Proposition 4.1 for $x_0=0$ and by Chebyshev's inequality, we have $p_1 \leq CN^{-\varepsilon/2}$. To bound p_2 , we distinguish two cases: $S_{\tau_N} > N^{1/2+\delta}$ and $S_{\tau_N} \leq N^{1/2+\delta}$. The first case has negligible probablity by moderate deviation bound for random walks (see formula (4.5) below). In the second case we compute

$$\mathbb{P}(|T_{\tau_N}| > 2N^{\rho + \varepsilon/2}, |T_N| \le N^{\rho + \varepsilon/2}, |S_{\tau_N}| \le N^{1/2 + \delta})$$

$$\le \sup_{x_0: |x_0| \le N^{1/2 + \delta}} \max_{n = 1, \dots, N} \mathbb{P}_{x_0}(|T_n| \ge N^{\rho + \varepsilon/2})$$

which is again bounded by $CN^{-\varepsilon/2}$ by Proposition 4.1 and Chebyshev's inequality. We have verified (4.3) and finished the proof of the theorem.

4.2. **Proof of Proposition 4.1 for** d = 1. We have $\beta \in (0,1)$ and $2\rho_d(\beta) = \beta + 1$. We start by recalling some results on expansions in the LLT in case all moments are finite (a.k.a. Edgeworth expansion).

Theorem 4.2. ([17, Section 51]) Under the assumptions of Theorem 1.4, there are polynomials $Q_1, Q_2, ...$ so that for any $M \in \mathbb{N}$

(4.4)
$$\mathbb{P}(S_n = l) = \frac{1}{\sqrt{n}} \mathfrak{g}(l/\sqrt{n}) (1 + \sum_{m=1}^M Q_m(l/\sqrt{n}) n^{-m/2}) + e_{n,l,M}$$

where \mathfrak{g} is a Gaussian density and

$$\limsup_{n\to\infty} \sup_{l\in\mathbb{Z}} e_{n,l,M} n^{M/2+1} < \infty.$$

Next, we claim that for any $\eta > 0$ there exists $C < \infty$ such that

$$(4.5) \mathbb{P}(|S_n| > n^{1/2+\eta}) < Cn^{-1/\eta}.$$

To prove this claim, first observe that

$$\mathbb{E}(S_n^{2k}) = \sum_{n_1=1}^n \dots \sum_{n_{2k}=1}^n \mathbb{E}(X_{n_1} \dots X_{n_{2k}}) \le C_{2k} n^k.$$

Indeed, since $\mathbb{E}(X_1) = 0$, for every i = 1, ..., 2k we need to have $j \neq i$ so that $n_i = n_j$ to make the expectation non zero. This estimate for $k > 1/(2\eta^2)$ combined with the Markov inequality gives (4.5).

Given a function $h = h(n, x) : \mathbb{Z}_+ \times \mathbb{Z} \to \mathbb{R}$, we write ∇h for the discrete derivative in the second coordinate, i.e.

$$\nabla h(n,x) = h(n,x) - h(n,x-1).$$

Note that

(4.6)
$$\nabla(gh)(n,x) = (\nabla g)(n,x)h(n,x) + g(n,x-1)(\nabla h)(n,x).$$

We will also write $\nabla^k h$ for the kth discrete derivative.

Denote

(4.7)
$$H(n,x) = \mathbb{P}(S_n = x).$$

With this notation, (4.5) can be rewritten as

(4.8)
$$\sum_{x:|x|>n^{1/2+\eta}} H(n,x) < C_{\eta} n^{-1/\eta} \text{ for any } \eta > 0.$$

Also (4.4) implies that there is a constant c so that, for every k = 0, 1, 2,

(4.9)
$$\sup_{x \in \mathbb{Z}} |\nabla^k H(n, x)| \le c n^{-\frac{k+1}{2}}.$$

Observe that

(4.10)
$$\mathbb{E}_{x_0}(T_N^2) = \sum_{0 \le n_1 \le n_2 \le N} c_{n_1, n_2} E_{n_1, n_2}(x_0)$$

where $c_{n_1,n_2} = 1$ if $n_1 = n_2$ and $c_{n_1,n_2} = 2$ otherwise and

$$E_{n_1,n_2}(x_0) = \mathbb{E}_{x_0}(F(S_{n_1})F(S_{n_2}))$$

$$= \sum_{x_1,x_2 \in \mathbb{Z}} \mathbb{P}(S_{n_2-n_1} = x_2 - x_1)\mathbb{P}_{x_0}(S_{n_1} = x_1)F(x_1)F(x_2)$$

We will show the following: for any $0 \le n_1 \le n_2 \le N$ such that

(4.11)
$$n_1 > N^{\alpha}, \quad n_2 - n_1 > N^{\alpha} \text{ where } \alpha = 1/\gamma = \beta/2$$

and for any x_0 with $|x_0| < N^{1/2+\delta}$, we have

(4.12)
$$|E_{n_1,n_2}(x_0)| \le$$

$$Cn_1^{\frac{\beta-1}{2}+\varepsilon} (n_2 - n_1)^{\frac{\beta-1}{2}+\varepsilon} + Cn_1^{\frac{\beta}{2}+\varepsilon} (n_2 - n_1)^{\frac{\beta-2}{2}+\varepsilon}$$

Summing the estimate (4.12) for n_1, n_2 satisfying (4.11) we obtain $N^{\beta+1+2\varepsilon}$ as needed. To complete the proof of the proposition, it remains to

- (I) prove (4.12);
- (II) verify that the contribution of (n_1, n_2) 's that do not satisfy (4.11) is also negligible.

We start with (I).

We will use the following lemma:

Lemma 4.3. There is a constant \hat{C} such that for any positive integer n and any constants A, B, the following holds. If a bounded function $g(x): \mathbb{Z} \to \mathbb{R}$ satisfies

$$(H1) \sup_{x:|x| < n^{1/2+\delta'}} |g(x)| \le A$$

(H1)
$$\sup_{\substack{x:|x|\leq n^{1/2+\delta'}\\x:|x|\leq n^{1/2+\delta'}}} |g(x)| \leq A$$
(H2)
$$\sup_{\substack{x:|x|\leq n^{1/2+\delta'}\\x:|x|\leq n^{1/2+\delta'}}} |\nabla g(x)| \leq B$$

for some sufficiently small δ' , then, for i = 0, 1,

(4.13)

$$\sup_{y:|y|\leq \frac{1}{2}n^{\frac{1/2+\delta'}{\alpha}}}\left|\sum_{z\in\mathbb{Z}}\nabla^{i}H(n,z-y)g(z)F(z)\right|\leq \hat{C}(\|g\|_{\infty}n^{-10}+An^{\frac{\beta-i-1}{2}+\varepsilon}+Bn^{\frac{\beta-i}{2}+\varepsilon}).$$

Note that y is allowed to be of order $n^{\frac{1/2+\delta'}{\alpha}} \gg n^{1/2+\delta'}$ which is the range for x in the hypothesis. This is important in the application of the lemma later, especially when n_1 is big and $n_2 - n_1$ is small.

Proof. For the rest of the section C will denote a constant (independent of A and B) whose value may change from line to line. By (4.8) and since F is bounded, the sum for

z's with
$$|z - y| > n^{1/2 + \delta'}$$
 is bounded by $C ||g||_{\infty} n^{-11}$. Denote $I(x) = \sum_{w = y - 2n^{1/2 + \delta'}}^{x} F(w)$.

Using summation by parts and (H1), we find

(4.14)
$$\sum_{z:|z-y| \le n^{1/2+\delta'}} \nabla^{i} H(n,z-y) g(z) F(z) = O\left(An^{-10}\right) + \sum_{z:|z-y| \le n^{1/2+\delta'}} I(z-1) \nabla_{z} \left(\nabla^{i} H(n,z-y) g(z)\right).$$

Using (4.6), (4.9), (H1) and (H2) we find that

(4.15)
$$\nabla_z \left(\nabla^i H(n, z - y) g(z) \right) \le C \left(A n^{-\frac{i+2}{2}} + B n^{-\frac{i+1}{2}} \right).$$

Next, we estimate |I(z)| for z satisfying $|z-y| \leq n^{1/2+\delta'}$. For any such z, I(z) is defined as a sum over the interval $[a,z] \cap \mathbb{Z}$, where $a=y-2n^{1/2+\delta'}$ and consequently the length of this interval satisfies $L=z-a \in [n^{1/2+\delta'},3n^{1/2+\delta'}]$. Using $\alpha=1/\gamma<1$ and our assumption $|y| \leq \frac{1}{2}n^{\frac{1/2+\delta'}{\alpha}}$, we find that $|a| < L^{\gamma}$ for n sufficiently large. Since $F \in \mathbf{G}_{\gamma}^{\beta}$, with $\bar{F}=0$, we use the definition of $\mathbf{G}_{\gamma}^{\beta}$ to conclude

$$(4.16) |I(z)| \le CL^{\beta} \le Cn^{\frac{\beta}{2} + \varepsilon}$$

(assuming that $\delta' = \delta'(\varepsilon)$ is small enough). The last two estimates imply that the sum in (4.14) is bounded by $C(An^{\frac{\beta-1-i}{2}+\varepsilon} + Bn^{\frac{\beta-i}{2}+\varepsilon})$.

Now we are ready to estimate $E_{n_1,n_2}(x_0)$. First, let

(4.17)
$$g_1(x_1) := g_{1,n_2-n_1}(x_1) := \sum_{x_2 \in \mathbb{Z}} H(n_2 - n_1, x_2 - x_1) F(x_2).$$

By definition, $||g_1||_{\infty} \le ||F||_{\infty}$. Applying Lemma 4.3 with i = 0, g = 1, $n = n_2 - n_1$, A = 1, B = 0 and using $n_2 - n_1 > N^{\alpha}$, we find

(4.18)
$$\sup_{x_1:|x_1|\leq \frac{1}{2}N^{1/2+\delta'}} |g_1(x_1)| \leq C(n_2-n_1)^{\frac{\beta-1}{2}+\varepsilon}.$$

Using Lemma 4.3 the same way but now with i = 1, we find

(4.19)
$$\sup_{x_1:|x_1|\leq \frac{1}{2}N^{1/2+\delta'}} |\nabla g_1(x_1)| \leq C(n_2-n_1)^{\frac{\beta-2}{2}+\varepsilon}$$

Next, set

(4.20)
$$g_2(x_0) := g_{2,n_1}(x_0) := \sum_{x_1 \in \mathbb{Z}} H(n_1 - n_0, x_1 - x_0) g_1(x_1) F(x_1)$$

Now we use Lemma 4.3 with $i=0,\ n=n_1,\ g=g_1,\ A=(n_2-n_1)^{\frac{\beta-1}{2}+\varepsilon},\ B=(n_2-n_1)^{\frac{\beta-2}{2}+\varepsilon}.$ Since $n\leq N,\ (4.18)$ and (4.19) give (H1) and (H2) (with $\delta'/2$ instead of δ'). Also using that $\|g_1\|_{\infty}\leq \|F\|_{\infty}^2$ and $n_1-n_2>N^{\alpha}$, we get

(4.21)
$$\sup_{x_0:|x_0|\leq \frac{1}{2}N^{1/2+\delta'/2}}|g_2(x_0)|\leq$$

$$Cn_1^{\frac{\beta-1}{2}+\varepsilon}(n_2-n_1)^{\frac{\beta-1}{2}+\varepsilon}+Cn_1^{\frac{\beta}{2}+\varepsilon}(n_2-n_1)^{\frac{\beta-2}{2}+\varepsilon}$$

which gives (4.12) (with $\delta = \delta'/4$).

It remains to verify (II), that is that the contribution of pairs (n_1, n_2) 's that do not satisfy (4.11) is negligible.

First, assume that $n_1 > N^{\alpha}$ and $n_2 - n_1 \le N^{\alpha}$. Then we derive as in (4.21) but using the trivial bounds $A = 1 + ||F||_{\infty}$, $B = 2(1 + ||F||_{\infty})$ that

$$\sup_{x_0:|x_0| \le N^{1/2+\delta}} |g_2(x_0)| \le C n_1^{\frac{\beta-1}{2}+\varepsilon} + C n_1^{\frac{\beta}{2}+\varepsilon}.$$

Summing this estimate for $n_1 = N^{\alpha}, ..., N$ and multiplying by N^{α} for the number of choices of n_2 , we obtain

$$(4.22) O(N^{\alpha}N^{\frac{\beta+1}{2}}) = O(N^{\beta+\frac{1}{2}}) = o(N^{\beta+1}).$$

Next, assume that $n_1 < N^{\alpha}$, $n_2 - n_1 < N^{\alpha}$. Using the bound $|E_{n_1,n_2}(x_0)| \leq ||F||_{\infty}^2$ we obtain

$$\sum_{n_1=0}^{N^{\alpha}} \sum_{n_2=n_1}^{n_1+N^{\alpha}} |E_{n_1,n_2}(x_0)| \le CN^{2\alpha} = CN^{\beta} = o(N^{\beta+1}).$$

Finally, assume that $n_1 \leq N^{\alpha}$ and $n_2 - n_1 > N^{\alpha}$. By (4.8), we can assume that $S_{n_1} - S_{n_0} \leq N^{1/2+\delta}$. Then (4.18) still holds and we conclude that

$$(4.23) |E_{n_1,n_2}(x_0)| \le C(n_2 - n_1)^{\frac{\beta - 1}{2} + \varepsilon}.$$

Summing for $n_2 = n_1 + N^{\alpha}, ..., N$ and multiplying by N^{α} , we obtain the same error term as in (4.22).

This completes the proof.

4.3. Proof of Proposition 4.1 for $d \ge 2$. In dimension d, we have

(4.24)
$$\sup_{x \in \mathbb{Z}^d} |\nabla_{i_1} ... \nabla_{i_k} H(n, x)| \le c n^{-\frac{k+d}{2}}.$$

for any $i_1, ..., i_k = 1, ..., d$, where ∇_i denotes the discrete derivative with respect to x_i , the *i*-th component of x. We apply a similar approach to the case d = 1. That is, we perform summations by parts to estimate g_1 and g_2 defined by (4.17) and (4.20). Each time we need d summations by parts. For example, if d = 2, then with $m = n_2 - n_1$,

$$g_{1,m}(0) = \sum_{|x| < m^{1/2+\delta}} \sum_{|y| < m^{1/2+\delta}} H(m,(x,y)) F(x,y)$$

$$\approx \sum_{|x| < m^{1/2+\delta}} \sum_{|y| < m^{1/2+\delta}} \nabla_2 H(m,(x,y)) I_1(x,y)$$

$$\approx \sum_{|x| < m^{1/2+\delta}} \sum_{|y| < m^{1/2+\delta}} \nabla_1 \nabla_2 H(m,(x,y)) I(x,y)$$

$$(4.25)$$

where

$$I_1(x,y) = \sum_{z=-m^{1/2+\delta}}^{y} F(x,z)$$
 and $I(x,y) = \sum_{w=-m^{1/2+\delta}}^{x} \sum_{z=-m^{1/2+\delta}}^{y} F(w,z)$

and $a_m \approx b_m$ means that the $a_m - b_m$ is superpolynomially small in m. Recalling that $F \in \mathbf{G}_{1/\varepsilon}^{\beta}$ and $\bar{F} = 0$, we have $|I(x,y)| \leq C m^{(1/2+\delta)2\beta}$. Substituting this estimate and (4.24) into (4.25), we find

$$|g_{1,m}(0)| \le \sum_{|x| < m^{1/2+\delta}} \sum_{|y| < m^{1/2+\delta}} Cm^{-2} m^{\beta+2\delta\beta} \le Cm^{\beta-1+\varepsilon}.$$

assuming $2\delta(1+\beta) < \varepsilon$. Using that $|I(x,y)| \leq Cm^{\beta(\frac{1}{2}+\delta)}$, we find

$$|g_{1,m}(0)| \le Cm^{\beta - 1 + \varepsilon}.$$

To simplify formulas, we will use the notation

$$a_N \lesssim b_N \text{ if } a_N \leq Cb_N N^{\varepsilon}.$$

Lemma 4.4. For any $a \in (0,1]$, if $F \in \mathbf{G}_{1/a}^{\beta}$, then for all n_1, n_2 satisfying $n_1 \geq N^a$, $n_2 - n_1 \ge N^a$ we have

(4.26)
$$\sup_{x_0 \in \mathbb{Z}^d: |x_0| \le N^{1/2+\delta}} |E_{n_1, n_2}(x_0)| \lesssim \sum_{j=0}^d n_1^{\frac{d\beta-j}{2}} (n_2 - n_1)^{\frac{d\beta-2d+j}{2}}.$$

Proof. Since the proof of the lemma is similar to that of (4.12), we only mention the main difference. That is, now j can take values 0, 1, ..., d and in dimension d = 1 it could only take values 0, 1. This follows from the fact that when applying d summations by parts to the function $H(n_1 - n_0, x_1 - x_0)g_1(x_1)$, we obtain

$$\nabla_{1}...\nabla_{d}(Hg) = \sum_{\{i_{1},...,i_{I}\}\subset\{1,...,d\}} (\nabla_{i_{1}}...\nabla_{i_{I}}H)(\nabla_{j_{1}}...\nabla_{j_{J}}g)$$

where $\{j_1,...,j_J\} = \{1,...,d\} \setminus \{i_1,...,i_I\}$. In the proof of (4.12), we only used the definition of $\mathbf{G}_{\gamma}^{\beta}$ for boxes with side length $L \geq N^{1/2+\delta}$ (specifically in deriving (4.16)). In order to extend that proof to the present setting, we only need to replace F(.) by $F(.-x_0)$. Thus assuming $\gamma > 1/a$ and using the definition of $\mathbf{G}_{\gamma}^{\beta}$, we can repeat the previous proof.

Next we show that the extreme terms in the right hand side of (4.26) provide the main contribution.

Set $m_1 = n_1, m_2 = n_2 - n_1$. By Lemma 4.4, we have for $m_1 \ge N^a, m_2 \ge N^a$, $|x_0| \leq N^{1/2+\delta}$ that

(4.27)
$$|E_{n_1,n_2}(x_0)| \lesssim \begin{cases} m_1^{\frac{(\beta-1)d}{2}} m_2^{\frac{(\beta-1)d}{2}} & \text{if } m_2 \ge m_1\\ m_1^{\frac{\beta d}{2}} m_2^{\frac{d\beta-2d}{2}} & \text{if } m_2 < m_1. \end{cases}$$

Note that the second bound is quite bad if $m_2 \ll m_1$. However we can improve it by bootstrap. Namely we have

Lemma 4.5. If $N^a < m_2 < m_1$ then

$$|E_{n_1,n_2}(x_0)| \lesssim m_2^{(\beta-1)d}$$

Proof. If $m_1 \leq 2m_2$ then the result follows from (4.27). If $m_1 > 2m_2$, let $k = m_1 - m_2$ and note that $k > m_2$ and

$$E_{n_1,n_2}(x_0) = \sum_{y \in \mathbb{Z}^d} H(k, y - x_0) E_{m_2,2m_2}(y).$$

The sum of the terms where $|y| > 2N^{1/2+\delta}$ decays faster than N^{-r} for any r. The terms where $|y| \le 2N^{1/2+\delta}$ can be estimated by (4.27) with $m_1 = m_2$ giving the result. \square

We now combine the foregoing results in different regimes in case where $a = \varepsilon$. Then if $\gamma = 1/\varepsilon$, $F \in \mathbf{G}_{\gamma}^{\beta}$ we gather that

(4.28)
$$|E_{n_1,n_2}(x_0)| \lesssim \begin{cases} m_1^{\frac{(\beta-1)d}{2}} m_2^{\frac{(\beta-1)d}{2}} & \text{if } m_2 \ge m_1 \ge N^{\varepsilon}, \\ m_2^{(\beta-1)d} & \text{if } m_1 > m_2 \ge N^{\varepsilon}, \\ 1 & \text{if } \min(m_1, m_2) < N^{\varepsilon}. \end{cases}$$

Summing the bounds of (4.28) for $m_1, m_2 \in \{1...N\}$ we obtain Proposition 4.1.

5. SLLN in dimension 1.

5.1. Reduction to occupation times sum. Here we prove Theorem 1.8.

Let $\ell_n(x)$ be time spent by the walker at site x before time n. Set $\ell_\infty(x) := \lim_{n \to \infty} \ell_n(x)$. Thus $\ell_\infty(x)$ is the total time spent by the walker at site x.

Lemma 5.1. There exist $C, c > 0, \mathfrak{p} \in (0,1), \varepsilon_1$ such that for all $x \in \mathbb{N}$ and $m \in \mathbb{N}$

(5.1)
$$\mathbb{P}(\ell_{\infty}(x) > m) < Ce^{-cm}.$$

Furthermore

(5.2)
$$\left| \mathbb{E}(\ell_{\infty}(x)) - \frac{1}{v} \right| \le \frac{C}{x^{\varepsilon_1}}, \quad |\mathbb{P}(\ell_{\infty}(x) = 0) - \mathfrak{p}| \le \frac{C}{x^{\varepsilon_1}}.$$

Proof. (5.2) follows from quantitative renewal theorem [32]. (5.1) holds since for $k \geq 1$, $\mathbb{P}(\ell_{\infty}(x) = k) = \mathbb{P}(\ell_{\infty}(x) \neq 0) \, \mathfrak{p}_0^{k-1} (1 - \mathfrak{p}_0)$ where \mathfrak{p}_0 is the probability that S_n returns to the origin at some positive moment of time.

Let $\tilde{T}_N = \sum_{x=1}^N \ell_\infty(x) F(x)$. We will show that with probability 1

(5.3)
$$\frac{\tilde{T}_N}{N} \to \frac{\bar{F}_+}{v}.$$

We first deduce Theorem 1.8 from (5.3) and then prove (5.3). Denote $\mathbb{L}_N = \sum_{x=1}^N \ell_{\infty}(x)$. By the strong law of large numbers for S_N

$$\frac{\mathbb{L}_N}{N} \to \frac{1}{v}.$$

On the other hand for each ε and for almost every ω , there is some $N_0 = N_0(\varepsilon, \omega)$ so that for all $N > N_0$,

$$\left| T_N - \tilde{T}_{Nv(1-\varepsilon)} \right| \le ||F||_{\infty} \left(\mathbb{L}^- + \left[\mathbb{L}_{Nv(1+\varepsilon)} - \mathbb{L}_{Nv(1-\varepsilon)} \right] \right),$$

where $\mathbb{L}^- = \sum_{x=-\infty}^0 \ell_{\infty}(x)$ is the total time spent on the negative halfline. In view of

(5.4), $\frac{T_N - \tilde{T}_{Nv(1-\varepsilon)}}{N}$ can be made as small as we wish by taking ε small. Hence Theorem 1.8 follows from (5.3).

In order to prove (5.3) we observe that by Lemma 5.1

$$\mathbb{E}\left(\frac{\tilde{T}_N}{N}\right) = \frac{1}{N} \sum_{x=1}^N F(x) \mathbb{E}(\ell_\infty(x)) = \frac{1}{v} \left[\frac{1}{N} \sum_{x=1}^N F(x) \right] + O\left(N^{-\varepsilon_1}\right) = \frac{\bar{F}_+}{v} + o(1),$$

as $N \to \infty$.

We need the following bound, which will be proved in §5.2.

Lemma 5.2. There are constants C and ε_2 such that for each $n_1 < n_2$

$$|\operatorname{Cov}(\ell_{\infty}(n_1), \ell_{\infty}(n_2))| \le C\left(\frac{1}{n_1^{\varepsilon_2}} + \frac{1}{(n_2 - n_1)^{\varepsilon_2}}\right).$$

Lemma 5.2 implies that

$$\operatorname{Var}\left(\frac{\tilde{T}_N}{N}\right) \le \frac{C}{N^{\varepsilon_2}}$$

and so

$$\mathbb{P}\left(\frac{|\tilde{T}_N - \mathbb{E}(\tilde{T}_N)|}{N} \ge \delta\right) \le \frac{C}{\delta^2 N^{\varepsilon_2}}.$$

Set $r = 2/\varepsilon_2$. By Borel-Cantelli Lemma

$$\frac{\tilde{T}_{n^r}}{n^r} \to \frac{\bar{F}}{v} \text{ as } n \to \infty$$

almost surely. On the other hand, (5.1) and the Borel-Cantelli Lemma imply that, with probability 1, for all sufficiently large x, $\ell_{\infty}(x) \leq \ln^2 x$. Given N, take n such that $n^r \leq N < (n+1)^r$. Then

$$\left| \tilde{T}_N - \tilde{T}_{n^r} \right| \le ||F||_{\infty} \left(\mathbb{L}_N - \mathbb{L}_{n^r} \right) \le C N^{(r-1)/r} \ln^2 N.$$

It follows that $\frac{\tilde{T}_N}{N} = \frac{\tilde{T}_{n^r}}{n^r} + o(1)$, for $N \to \infty$, proving (5.3).

5.2. Covariance of occupation times. The proof of Lemma 5.2 relies on the following estimates.

Lemma 5.3. There are constants C and ε_3 such that for all $m \geq 1$,

$$\mathbb{P}(\min_{n}(S_n) \le -m) \le \frac{C}{m^{\varepsilon_3}}.$$

Lemma 5.3 (with $\varepsilon_3 = \beta - 1$) follows from Theorem 2(B) of [35].

Lemma 5.4. For each $\delta > 0$ there is a constant $C(\delta)$ such that the following holds. Consider a Markov chain with states $\{1, 2, 3\}$ and transition matrix

$$\begin{pmatrix}
p_1 & q_1 & \eta_1 \\
q_2 & p_2 & \eta_2 \\
0 & 0 & 1
\end{pmatrix}$$

and initial distribution (π_1, π_2, π_3) . Assume that

$$(5.5) q_1 > \delta, \quad and \quad \eta_2 > \delta.$$

Let l_1 and l_2 denote the occupation times of sites 1 and 2. Then

$$|\text{Cov}(\mathfrak{l}_1, \mathfrak{l}_2)| \le C(\delta) \left(\frac{q_1}{q_1 + \eta_1} (1 - \pi_1) - \pi_2 + q_2 \right).$$

In the special case where $\eta_1 = 0$, $\pi_1 = 1$, Lemma 5.4 follows from [10, Lemma 3.9(a)]. In this case the statement simplifies significantly since the first term in the RHS vanishes.

The proof of Lemma 5.4 will be given in §5.3.

We apply Lemma 5.4 to the states (n_1, n_2, ∞) with $n_1 < n_2$. This means that we define a 3-state Markov chain as a function of the random walk (S_k) , such that the chain starts in the state 1, if the random walk visits n_1 for the first time before it visits n_2 ; or in the state 2, if the random walk visits n_2 for the first time before it visits n_1 ; or in in the state 3 if the random walk never visits n_1 or n_2 . After that, the chain transitions to state 1, 2 or 3, respectively, if the next return of the random walk to the set $\{n_1, n_2\}$ occurs at n_1 , n_2 , or never does. Clearly 3 is an absorbing state for this chain. So

$$\pi_1 = \mathbb{P}(n_1 \text{ is visited before } n_2), \quad \pi_2 = \mathbb{P}(n_2 \text{ is visited before } n_1),$$

$$\pi_3 = \mathbb{P}(n_1 \text{ and } n_2 \text{ are not visited}).$$

Let V_n be the event that n is visited by our random walk. Note that Lemma 5.3 implies $q_2 = O\left((n_2 - n_1)^{-\varepsilon_3}\right)$. Hence the probability that both n_1 and n_2 are visited with n_2 being the first is also $O\left((n_2 - n_1)^{-\varepsilon_3}\right)$. Therefore

$$\pi_2 = \mathbb{P}(V_{n_2}) - \mathbb{P}(n_1 \text{ visited first, then } n_2 \text{ is visited}) \simeq \mathbb{P}(V_{n_2}) - \mathbb{P}(V_{n_1} \cap V_{n_2}),$$

where \approx means the difference between the LHS and the RHS is

$$O(n_1^{-\varepsilon_2}) + O((n_2 - n_1)^{-\varepsilon_2})$$
 where $\varepsilon_2 = \min(\varepsilon_1, \varepsilon_3)$.

Likewise,

$$\pi_1 \asymp \mathfrak{q}, \quad \frac{q_1}{q_1 + \eta_1} = \mathbb{P}_{n_1}(V_{n_2}) \asymp \mathfrak{q},$$

where $\mathfrak{q} = 1 - \mathfrak{p}$ (see (5.2)). Combining the last two displays, we obtain

$$\pi_2 \simeq \mathbb{P}(V_{n_2}) - \mathbb{P}(V_{n_1} \cap V_{n_2}) \simeq \mathfrak{q} - \mathfrak{q}^2,$$

These estimates, combined with Lemma 5.4 imply Lemma 5.2.

5.3. Analysis of three state chains.

Proof of Lemma 5.4. Under the assumptions of the lemma, l_1 , l_2 and l_1l_2 are uniformly integrable (the uniformity is over all chains satisfying (5.5)). Let \bar{l}_1 be the time spent at 1 before the first visit to another state. Then, by the uniform integrability,

$$\mathbb{E}(\mathfrak{l}_1 - \bar{\mathfrak{l}}_1) = O(q_2), \quad \mathbb{E}((\mathfrak{l}_1 - \bar{\mathfrak{l}}_1)\mathfrak{l}_2) = O(q_2), \quad \mathbb{E}_2(\mathfrak{l}_1) = O(q_2), \quad \mathbb{E}_3(\mathfrak{l}_i) = 0.$$

Hence

$$\mathbb{E}(\mathfrak{l}_1\mathfrak{l}_2) = \pi_1\mathbb{E}_1(\bar{\mathfrak{l}}_1)\frac{q_1}{q_1+\eta_1}\mathbb{E}_2(\mathfrak{l}_2) + O(q_2);$$

$$\mathbb{E}(\mathfrak{l}_1) = \pi_1 \mathbb{E}_1(\bar{\mathfrak{l}}_1) + O(q_2) \quad \text{and} \quad \mathbb{E}(\mathfrak{l}_2) = (\pi_1 \frac{q_1}{q_1 + \eta_1} + \pi_2) \mathbb{E}_2(\mathfrak{l}_2) + O(q_2).$$

Therefore

$$Cov(\mathfrak{l}_{1},\mathfrak{l}_{2}) = \pi_{1}\mathbb{E}_{1}(\bar{\mathfrak{l}}_{1})\frac{q_{1}}{q_{1}+\eta_{1}}\mathbb{E}_{2}(\mathfrak{l}_{2}) - \pi_{1}\mathbb{E}_{1}(\bar{\mathfrak{l}}_{1})(\pi_{1}\frac{q_{1}}{q_{1}+\eta_{1}}+\pi_{2})\mathbb{E}_{2}(\mathfrak{l}_{2}) + O(q_{2})$$

$$= \pi_{1}\mathbb{E}_{1}(\bar{\mathfrak{l}}_{1})\mathbb{E}_{2}(\mathfrak{l}_{2})\left[\frac{q_{1}}{q_{1}+\eta_{1}}(1-\pi_{1})-\pi_{2}\right] + O(q_{2})$$

as claimed. \Box

6. Counterexamples to the strong law.

Here we prove Theorem 1.9.

Consider first the case d = 1. Assume that we are given a strictly increasing sequence $d_k \in \mathbb{N}$, $d_1 = 1$, $d_k \to \infty$ sufficiently quickly. We will impose finitely many lower bounds on the growth of d_k and so one can take the biggest lower bound. Now, for $l = d_k, d_k + 1, ..., d_{k+1} - 1$, define $a_l = k$. Notice that there is a bijective correspondence between strictly increasing sequences (d_n) and diverging sequences (a_n) such that $a_n - a_{n-1} \in \{0, 1\}$.

Next, define $b_1 \gg 1$, $b_{n+1} = b_n + \lfloor b_n/a_n \rfloor$. By induction, we see that

(6.1)
$$a_n \le n < b_n < b_{n+1} \quad \text{and, for large } n, \quad b_n < 2^n.$$

Consider the function F defined by F(0) = 0,

(6.2)
$$F(x) = \begin{cases} 1 & \text{if } b_{2k} \le x < b_{2k+1} \text{ for some } k \\ 0 & \text{if } b_{2k+1} \le x < b_{2k+2} \text{ for some } k \end{cases}$$

for x > 0 and F(x) = F(-x) for x < 0. We start by verifying that $F \in \mathbf{G}_0$.

Lemma 6.1. If

$$\lim_{k \to \infty} \sqrt{d_k/d_{k-1}} = \infty,$$

then $F \in \mathbf{G}_0$ with $\bar{F} = 1/2$.

Note that (6.3) is satisfied if e.g. $d_1 = 1$ and $d_k = 2^{2^{2^k}}$ for $k \ge 2$.

Proof. Recall that $b_n < 2^n$, for large n (say for $n \ge n_0$). Next, we establish the a priori lower bound

for some c. To prove (6.4), first observe that, by (6.3), there is some C so that for all $n \ge n_0$, $a_n < C\sqrt{n}$. Thus

$$b_{n+1} \ge b_n + \lfloor b_n / C\sqrt{n} \rfloor \ge b_n (1 + 1/(C\sqrt{n})) - 1 \ge b_n (1 + 1/(2C\sqrt{n}))$$

where the last inequality follows from $n < b_n$. Thus $b_n \ge b_{n_0} \prod_{k=n_0}^{n-1} (1+1/(2C\sqrt{k}))$. Hence

$$b_n \ge \exp\left[\ln(b_{n_0}) + \sum_{k=1}^n \ln(1 + 1/(2C\sqrt{k}))\right] \ge 2^{c\sqrt{n}}$$

for $n \ge n_0$. Clearly, we can assume that (6.4) holds for all $n < n_0$ by further decreasing c if necessary.

We will prove that $\lim_{m\to\infty} S(m)/m = 0$, where $S(m) = \sum_{x=1}^m (F(x) - 1/2)$. This will imply the lemma. Now given m, let n' be so that $b_{n'-1} \leq m < b_{n'}$. Let us denote

$$n'' = \max\{n \le n' : a_n = a_{n-1} + 1\}$$
$$n''' = \max\{n < n'' : a_n = a_{n-1} + 1\}$$

which are well defined for m large enough. Assume without loss of generality that n' is even (the case of odd n' is similar). Then

$$S(m) \le \frac{1}{2}b_{n'''} + \sum_{x=b_{n'''}+1}^{b_{n''}} \left(F(x) - \frac{1}{2}\right) + \sum_{x=b_{n''}+1}^{b_{n'}} \left(F(x) - \frac{1}{2}\right) =: \frac{1}{2}b_{n'''} + S_1 + S_2$$

Note that we have

$$S_1 = \frac{1}{2} \sum_{n=n'''}^{n''-1} (-1)^n (b_{n+1} - b_n).$$

We claim that

(6.5)
$$|\mathcal{S}_1| \le \frac{1}{2} (b_{n''} - b_{n''-1}).$$

To prove (6.5), observe that for n = n''', n''' + 1, ..., n'' - 1, the sequence a_n is constant and hence the sequence $n \mapsto (b_{n+1} - b_n)$ is monotone increasing for this range of n's. So in case n'' - 1 is even, we have

$$\frac{1}{2} \sum_{n=n'''}^{n''-2} (-1)^n (b_{n+1} - b_n) \le 0 \le \mathcal{S}_1.$$

Hence

$$0 \le \mathcal{S}_1 \le \mathcal{S}_1 - \frac{1}{2} \sum_{n=n'''}^{n''-2} (-1)^n (b_{n+1} - b_n) = \frac{1}{2} (b_{n''} - b_{n''-1})$$

(the case when n'' - 1 is odd is similar). We have verified (6.5). Likewise, we have

$$|\mathcal{S}_2| \le \frac{1}{2}(b_{n'} - b_{n'-1}).$$

Consequently,

$$\frac{\mathcal{S}(m)}{m} \le \frac{b_{n''}}{2b_{n'}} + \frac{b_{n''} - b_{n''-1}}{2b_{n'}} + \frac{b_{n'} - b_{n'-1}}{2b_{n'}}.$$

Since $b_{n'-1} \sim m \sim b_{n'}$, the last two terms on the RHS converge to zero. To estimate the first term, we first use (6.4) to derive

$$\frac{b_{n'''}}{b_{n'}} \le \frac{b_{n'''}}{b_{n''}} < 2^{n''' - c\sqrt{n''}}.$$

Now by the definition of n'' and n''', there is some k_0 so that $n'' = d_{k_0}$ and $n''' = d_{k_0-1}$ and so the RHS of the last displayed inequality also converges to zero as $m \to \infty$ (and consequently $k_0 \to \infty$) by (6.3). The lemma follows.

Let us denote $c_n = (b_n + b_{n+1})/2$, $t_n = \lfloor c_n^{\alpha} \rfloor$ and

$$I_n = [c_n - \lfloor b_n/4a_n \rfloor, c_n + \lfloor b_n/4a_n \rfloor].$$

We will show that almost surely, infinitely many of the events

$$A_{2n} = \{ \forall k \in [t_{2n}, 3t_{2n}] : S_k \in I_{2n} \}$$

occur and, likewise, infinitely many of the events

$$A_{2n+1} = \{ \forall k \in [t_{2n+1}, 3t_{2n+1}] : S_k \in I_{2n+1} \}$$

occur. This proves the theorem as A_{2n} implies $T_{3t_{2n}} \ge 2t_{2n}$ and A_{2n+1} implies $T_{3t_{2n+1}} \le t_{2n+1}$.

To complete the proof, let us fix a sequence $D_n \nearrow \infty$ such that

$$\sum_{n} \mathbb{P}(|S_n| \ge D_n) < \infty.$$

Then by the Borel-Cantelli Lemma,

$$\mathbb{P}(\exists N : \forall n > N : |S_n| < D_n) = 1.$$

Now we choose a subsequence $n_k \in \mathbb{Z}$ inductively so that

$$(6.6) n_{k+1} \equiv n_k \pmod{2}$$

and

(6.7)
$$n_{k+1} > \max \left\{ \exp \left(D_{\lceil 2^{\alpha(n_k+1)} \rceil} \right), \exp \left(\frac{1}{\alpha} 2^{(n_k+1)\alpha} \right) \right\}.$$

These bounds, combined with (6.1), give

(6.8)
$$b_{n_{k+1}} > \exp(D_{t_{n_k}}) \text{ and } t_{n_{k+1}} > \exp(t_{n_k}).$$

We want to show that for every $\varepsilon > 0$ and every K,

(6.9)
$$\mathbb{P}\left(\bigcap_{k=K}^{\infty} A_{n_k}^c\right) < \varepsilon.$$

Since $\varepsilon > 0$ is arbitrary, it follows that infinitely many of the events A_{n_k} happen.

Choosing n_k to be even for all k we see that almost surely infinitely many of the events A_{2n} and happen. Likewise, choosing n_k to be odd for all k we see that infinitely many of the events A_{2n+1} happen. Thus it remains to verify (6.9).

Given ε and K, choose K' > K so that $\mathbb{P}(\mathcal{B}) < \varepsilon$, where

$$\mathcal{B} = \{ \exists n \ge n_{K'}^{\alpha} : |S_n| \ge D_n \}.$$

Then we write

$$(6.10) \mathbb{P}\left(\bigcap_{k=K}^{\infty} A_{n_k}^c\right) \leq \mathbb{P}\left(\bigcap_{k=K'}^{\infty} A_{n_k}^c\right) \leq \varepsilon + \mathbb{P}\left(\bigcap_{k=K'}^{\infty} A_{n_k}^c \cap \mathcal{B}^c\right).$$

By construction, we have

$$\mathbb{P}\left(A_{n_{k+1}}^{c} \cap \mathcal{B}^{c} \middle| \bigcap_{j=K'}^{k} A_{n_{j}}^{c} \cap \mathcal{B}^{c}\right) \leq 1 - \mathbb{P}\left(A_{n_{k+1}} \middle| \bigcap_{j=K'}^{k} A_{n_{j}}^{c} \cap \mathcal{B}^{c}\right) \\
\leq 1 - \min_{x:|x| < D_{3t_{n_{k}}}} \mathbb{P}(A_{n_{k+1}} \middle| S_{3t_{n_{k}}} = x)$$
(6.11)

for k > K'.

Lemma 6.2. There is a constant K_0 and a sequence $a_n \nearrow \infty$ with $a_n - a_{n-1} \in \{0, 1\}$ such that for any $k \ge K_0$,

$$\min_{x:|x| < D_{3t_{n_k}}} \mathbb{P}(A_{n_{k+1}}|S_{3t_{n_k}} = x) \ge \frac{1}{k}.$$

Also, (a_n) is such that the corresponding (d_n) satisfies (6.3).

Clearly, Lemma 6.2 combined with (6.11) and (6.10) implies (6.9). Thus the proof of Theorem 1.9 for the case d = 1 will be completed once we prove Lemma 6.2.

Proof of Lemma 6.2. Recall that the invariance principle gives

$$\frac{S_{\lfloor Nt \rfloor}}{N^{1/\alpha}} \Rightarrow Y_t,$$

where Y_t is a stable Lévy process. In particular, Y_1 is a stable random variable with parameter α and "skewness" $\beta \in [-1, 1]$ (see e.g. [3], Chapter VIII). Now we distinguish two cases.

Case 1 $\alpha > 1$ or $|\beta| \neq 1$. The proof in case 1 consists of 3 steps.

Step 1: We prove that q > 0 where

(6.13)
$$q = \inf_{y \in [-1/16, 1/16]} \mathbb{P}\left(\sup_{t < 1} |Y_t| < \frac{1}{4}, |Y_1| < \frac{1}{16} |Y_0 = y|\right).$$

Indeed, as $\alpha > 1$ or $|\beta| \neq 1$, the stable process Y_t cannot be a subordinator. In particular, the density of Y_t is positive everywhere for every t > 0 and Y_t has the scaling property (see page 216 in [3]). Thus we have

$$\liminf_{\varepsilon \searrow 0} \inf_{y \in [-1/16, 1/16]} \mathbb{P}\left(|Y_{\varepsilon}| < \frac{1}{16} | Y_0 = y \right) = p > 0.$$

By Exercise 2 of Chapter VIII in [3], there is some $\varepsilon > 0$ such that

$$\mathbb{P}\left(\sup_{t\leq\varepsilon}|Y_t|<\frac{3}{16}\Big|Y_0=0\right)>1-p/2.$$

Combining the last two displayed equations, we derive

$$\inf_{y \in [-1/16, 1/16]} \mathbb{P}\left(\sup_{t \le \varepsilon} |Y_t| < \frac{1}{4}, |Y_{\varepsilon}| < \frac{1}{16} |Y_0 = y\right) \ge p/2.$$

Applying this inequality inductively, we obtain that $q \geq (p/2)^{\lceil 1/\varepsilon \rceil}$, which completes Step 1.

Step 2: We prove that there is some \bar{c} so that for every k

(6.14)
$$\min_{x:|x|< D_{3t_{n_k}}} \mathbb{P}\left(\left|S_{t_{n_{k+1}}} - c_{n_{k+1}}\right| < \frac{b_{n_{k+1}}}{16a_{n_{k+1}}} \left|S_{3t_{n_k}} = x\right| \ge \frac{\bar{c}}{a_{n_{k+1}}}.$$

To simplify the formulas, let us write

$$\tilde{a} = a_{n_{k+1}}, \ \tilde{b} = b_{n_{k+1}}, \ \tilde{c} = c_{n_{k+1}}, \ \tilde{t} = t_{n_{k+1}}.$$

Recall that the density of Y_1 is strictly positive. The LLT now implies that for every K > 0 there is some c so that for all z with $|z| \le K$,

$$\mathbb{P}(S_N = \lfloor zN^{1/\alpha} \rfloor) \ge \frac{c}{N^{1/\alpha}}.$$

Next observe that by the definition of $\tilde{a}, \tilde{b}, \tilde{c}, \tilde{t}$ and by (6.8)

$$t_{n_{k+1}} - 3t_{n_k} \sim \tilde{t} \sim \tilde{b}^{\alpha}, \quad \tilde{c} \pm D_{3\tilde{t}} \sim \tilde{t}^{1/\alpha} \text{ as } k \to \infty.$$

Thus for every $x \in \mathbb{Z}$ with $|x| < D_{3\tilde{t}}$,

$$\mathbb{P}\left(\left|S_{\tilde{t}} - \tilde{c}\right| < \frac{\tilde{b}}{4\tilde{a}} \middle| S_{3t_{n_k}} = x\right) = \mathbb{P}\left(\left|S_{\tilde{t} - 3t_{n_k}} - (\tilde{c} - x)\right| < \frac{\tilde{b}}{4\tilde{a}}\right) \ge \frac{c}{\tilde{b}} \frac{\tilde{b}}{4\tilde{a}} = \frac{\bar{c}}{\tilde{a}}.$$

Step 3: We prove that there is some $\bar{p} \in (0,1)$ such that

(6.15)
$$\min_{x:|x|$$

The lemma will follow from (6.15) as we can assume that our sequence $a_n \nearrow \infty$ with $a_n - a_{n-1} \in \{0, 1\}$ satisfies $a_{n_k} \le (-\log k/\log \tilde{p})^{1/\alpha}$ for a fixed $\tilde{p} \in (0, \bar{p})$.

To prove (6.15), we will combine (6.13) with (6.14). Namely, define

$$\tilde{A} = A_{n_{k+1}}, \quad \tilde{N} = \left\lceil \left(\frac{\tilde{b}}{\tilde{a}}\right)^{\alpha} \right\rceil, \quad \tilde{l} = \left\lceil 2\tilde{a}^{\alpha} \right\rceil + 1.$$

Let us define the event $\tilde{B}_0 = \{|S_{\tilde{t}} - \tilde{c}| < \frac{\tilde{N}^{1/\alpha}}{16}\}$ and for $\bar{l} = 1, ..., \tilde{l}$:

$$\tilde{B}_{\bar{l}} = \tilde{B}_0 \cap \left\{ \forall l = 0, ..., \bar{l} - 1, \forall m = 1, ..., \tilde{N} - 1 : \\ |S_{\tilde{t} + l\tilde{N} + m} - \tilde{c}| < \frac{\tilde{N}^{1/\alpha}}{4}, |S_{\tilde{t} + (l+1)\tilde{N}} - \tilde{c}| < \frac{\tilde{N}^{1/\alpha}}{16} \right\}.$$

Note that $\tilde{B}_{\tilde{l}}$ implies \tilde{A} . Thus it is enough to prove (6.15) with \tilde{A} replaced by $\tilde{B}_{\tilde{l}}$. First, (6.14) implies

$$\min_{x:|x|< D_{3t_{n_k}}} \mathbb{P}(\tilde{B}_0|S_{3t_{n_k}} = x) \ge \frac{\bar{c}}{\tilde{a}}.$$

Next, we derive

$$\min_{x:|x|< D_{3t_{n_k}}} \mathbb{P}(\tilde{B}_1|S_{3t_{n_k}} = x) \ge$$

$$\min_{x:|x|< D_{3t_{n_k}}} \mathbb{P}\left(\max_{x\in S_m} |S_m| < \frac{\tilde{N}^{1/\alpha}}{2} |S_{\tilde{N}}| < \frac{\tilde{N}^{1/\alpha}}{2}$$

$$\min_{x:|x|< D_{3t_{n_k}}} \mathbb{P}(\tilde{B}_0|S_{3t_{n_k}} = x) \min_{y:|y|<\frac{\tilde{N}^{1/\alpha}}{16}} \mathbb{P}\left(\max_{m=1,\dots,\tilde{N}-1} |S_m| < \frac{\tilde{N}^{1/\alpha}}{4}, |S_{\tilde{N}}| < \frac{\tilde{N}^{1/\alpha}}{16} |S_0 = y\right) \\
\geq \frac{\bar{c}}{\tilde{a}} \frac{q}{2},$$

where the first inequality follows from the Markov property and the second one follows from (6.12) and (6.13). Now we can apply an induction on \bar{l} , to obtain

$$\min_{x:|x|< D_{3t_{n_k}}} \mathbb{P}(\tilde{B}_{\bar{l}}|S_{3t_{n_k}} = x) \ge \frac{\bar{c}}{\tilde{a}} \left(\frac{q}{2}\right)^{\bar{l}}.$$

In particular,

$$\min_{x:|x| < D_{3t_{n_k}}} \mathbb{P}(\tilde{B}_{\tilde{l}}|S_{3t_{n_k}} = x) \ge \frac{\bar{c}}{\tilde{a}} \left(\frac{q}{2}\right)^{\tilde{l}}$$

which completes the proof of Step 3.

Case 2 $\alpha < 1$ and $|\beta| = 1$. Let us assume $\beta = 1$ (otherwise apply the forthcoming argument to $-X_i$).

(6.14) still holds in case 2, however a new approach is required to estimate $\mathbb{P}(A_{n_{k+1}}|S_{t_{n_k}}=x)$ since now the process Y_t (a.k.a. stable subordinator) is non-decreasing and thus q=0. Note however that in case 2, $\sup_{t\leq 1}|Y_t|=Y_1$ and thus it suffices to estimate one random variable instead of a stochastic process. Recall that the density of Y_1 is strictly positive on \mathbb{R}^+ . Thus applying (6.12) to $|X_i|$ (which is also in the standard domain of attraction of the totally skewed α -stable distribution) we obtain the following: for any $\varepsilon > 0$ there exist $N_0(\varepsilon)$ and $\delta(\varepsilon) > 0$ such that for any $N \geq N_0(\varepsilon)$,

(6.16)
$$\mathbb{P}\left(\sum_{n=1}^{3N} |X_n| \le \frac{\varepsilon}{8} N^{1/\alpha}\right) > \delta(\varepsilon).$$

Without loss of generality, we assume that N_0 and δ are, respectively, non-decreasing and non-increasing functions of ε . Now we define the sequence a_n inductively. First, let $a_1 = 1$. Now assume that a_{n_k} is defined. Let $a_m = a_{n_k}$ for $m = n_k + 1, ..., n_{k+1} - 1$. Next, we define $a_{n_{k+1}} = a_{n_k} + 1$ if both of the following conditions are satisfied:

(A)
$$N_0\left(\frac{1}{a_{n_k}+1}\right) < n_{k+1}^{\alpha}$$
 and (B) $\frac{\bar{c}}{2a_{n_k}+1} \delta\left(\frac{1}{a_{n_k}+1}\right) > \frac{1}{k+1}$.

Here, \bar{c} is the constant from (6.14). If either (A) or (B) fails, we put $a_{n_{k+1}} = a_{n_k}$. Note that by (6.7), the resulting sequence d_k satisfies (6.3).

Observe that by our construction, for all k, we have

$$(6.17) N_0\left(\frac{1}{a_{n_k}}\right) < n_k^{\alpha};$$

(6.18)
$$\frac{\bar{c}}{2a_{n_k}} \delta\left(\frac{1}{a_{n_k}}\right) > \frac{1}{k}.$$

Indeed, if $a_{n_{k+1}} = a_{n_k} + 1$ then (6.17) and (6.18) follow from conditions (A) and (B) above. If $a_{n_{k+1}} = a_{n_k}$ then (6.17) and (6.18) follow by induction since the LHSs of both (6.17) and (6.18) do not change when we replace k by k + 1, while the RHS of (6.17) increases and the RHS of (6.18) decreases.

By construction, $a_n \nearrow \infty$. Let K_0 be the smallest integer k so that $a_{n_k} = 2$. We prove that the lemma holds with this choice of K_0 and a_n . Recall that by (6.1), $b_{n_k} > n_k$ and so by (6.17), $N := b_{n_k}^{\alpha} > N_0(\varepsilon)$ with $\varepsilon = 1/a_{n_k}$. Applying (6.16) with this N and ε and using (6.18), we obtain

$$\mathbb{P}\left(\left|S_m - S_{t_{n_k}}\right| < \frac{b_{n_k}}{8a_{n_k}} \ \forall m \le 3b_{n_k}^{\alpha}\right) > \frac{a_{n_k}}{\bar{c}} \frac{2}{k}.$$

and since $t_{n_k} < b_{n_k}^{\alpha}$ and k-1 > k/2, we arrive at

(6.19)
$$\mathbb{P}\left(\left|S_m - S_{t_{n_k}}\right| < \frac{b_{n_k}}{8a_{n_k}} \ \forall m \le 3t_{n_k}\right) > \frac{a_{n_{k-1}}}{\bar{c}} \frac{1}{k-1}.$$

As discussed above, (6.14) holds. Since $a_{n_{k+1}} - a_{n_k} \in \{0, 1\}$ and $a_{n_k} \nearrow \infty$, we can also assume changing \bar{c} if necessary that (6.14) holds with $a_{n_{k+1}}$ replaced by a_{n_k} on the RHS. Now we use the Markov property, (6.19) with k replaced by k+1 and (6.14) (with the updated RHS) to derive

$$\begin{split} \min_{x:|x| < D_{3t_{n_k}}} \mathbb{P}(A_{n_{k+1}} | S_{3t_{n_k}} = x) \\ & \geq \min_{x:|x| < D_{3t_{n_k}}} \mathbb{P}\left(\left| S_{t_{n_{k+1}}} - c_{n_{k+1}} \right| < \frac{b_{n_{k+1}}}{16a_{n_{k+1}}} \middle| S_{3t_{n_k}} = x \right) \times \\ & \times \mathbb{P}\left(\left| S_m - S_{t_{n_{k+1}}} \right| < \frac{b_{n_{k+1}}}{8a_{n_{k+1}}} \ \forall m \leq 3t_{n_k} \right) \\ & \geq \frac{\bar{c}}{a_{n_k}} \frac{a_{n_k}}{\bar{c}} \frac{1}{k} = \frac{1}{k}. \end{split}$$

The lemma follows.

The above proof, with a few minor adjustments, applies to arbitrary dimension d. Specifically, we need to consider the function $\mathcal{F} \in \mathbf{G}_0$ defined by

$$\mathcal{F}(x_1, ..., x_d) = \begin{cases} F(x_1) & \text{if } |x_i| \le |x_1| \text{ for } i = 2, ..., d\\ \frac{1}{2} & \text{otherwise,} \end{cases}$$

where F is given by (6.2) and we need to replace I_n by

$$[c_n - \lfloor b_n/4a_n \rfloor, c_n + \lfloor b_n/4a_n \rfloor] \times [-\lfloor b_n/4a_n \rfloor, \lfloor b_n/4a_n \rfloor]^{d-1}.$$

Remark 6.3. It is easy to adjust the above proof to derive the following stronger version of Theorem 1.9: There is a function $F \in \mathbf{G}_0$ so that F only takes values $\{0,1\}$, $\bar{F} = 1/2$ and for almost every ω and for any $a \in [0,1]$, there is a subsequence $n_k = n_k(a,\omega)$ such that $T_{n_k}/n_k \to a$.

7. Conclusions.

The results proven in this paper show that for random walks the weak law of large numbers holds in the largest possible space of global observables, namely \mathbf{G}_0 . On the other hand, the strong law of large numbers fails, for general \mathbf{G}_0 observables, except for the walks with drift in dimension 1. In that case the path of the walk is almost deterministic and so the ergodic theory for occupation times could be used. The good news is that the weak law of large numbers seems to be a good setting for homogenization theory (cf Theorem 1.2), so the space \mathbf{G}_0 could be useful for that purpose. If we have some control on fluctuations over the mesoscopic scale as provided, for example, by the space \mathbf{G}_{γ} , then we can ensure the strong law. If we have polynomial control on the mesoscopic scale, as provided by the space $\mathbf{G}_{\gamma}^{\beta}$ then we can estimate the rate of convergence. In particular, our results give optimal rate of convergence for two important special cases: random walks in random scenery and quasi-periodic observables.

We note that the main ingredient in most proofs is local limit theorem and its extensions, such as the Edgeworth expansion used in Section 4. This makes it plausible that similar results hold for other systems where the local limit theorem holds, including the systems described in [2, 9, 11, 12, 13, 14, 15]. Another natural research direction motivated by the present work is limit theorems for global observables. It is likely that just assuming that F belongs to an appropriate \mathbf{G}_* will not be enough to derive limit theorems. For example, the computation of variance for T_N done in Sections 2 and 4 involve the expression of the form $\tilde{F}_z(x) := F(x)F(x+z)$ for a fixed $z \in \mathbb{Z}^d$. Therefore additional restrictions seem to be required to obtain limit theorems.

Extending our results to more general systems as well as limit theorems for global observables will be the subject of future work.

ACKNOWLEDGEMENTS.

The research of DD was partially sponsored by NSF DMS 1665046. The research of ML was partially supported by PRIN 2017S35EHN (MUR, Italy). The research of PN was partially sponsored by NSF DMS 1800811 and NSF DMS 1952876.

REFERENCES

- [1] J. Aaronson An introduction to infinite ergodic theory, Mathematical Surv. & Monographs 50 (1997) AMS, Providence, RI. xii+284 pp.
- [2] N. Berger, M. Cohen, R. Rosenthal Local limit theorem and equivalence of dynamic and static points of view for certain ballistic random walks in i.i.d. environments, Ann. Probab. 44 (2016) 2889–2979.
- [3] J. Bertoin, Levy processes (1996) Cambridge Univ. Press, Cambridge. x+265 pp.
- [4] E. Bolthausen, A central limit theorem for two-dimensional random walks in random sceneries. Ann. Probab. 17 (1) (1989), 108–115.
- [5] E. Bolthausen, A.-S. Sznitman *Ten lectures on random media*, DMV Seminar **32** (2002) Birkhauser, Basel, vi+116 pp.
- [6] C. Bonanno, P. Giulietti, M. Lenci Infinite mixing for one-dimensional maps with an indifferent fixed point, Nonlinearity 31 (2018), 5180–5213.
- [7] C. Bonanno, M. Lenci Pomeau-Manneville maps are global-local mixing, Discrete Contin. Dyn. Syst. 41 (3) (2021) 1051–1069.
- [8] A. N., Borodin A limit theorem for sums of independent random variables defined on a recurrent random walk, (in Russian) Dokl. Akad. Nauk SSSR **246** (1979), 4, 786–787.
- [9] A. Chiarini, J.-D. Deuschel Local central limit theorem for diffusions in a degenerate and unbounded random medium, Electron. J. Probab. 20 (2015), 112, 30 pp.
- [10] D. Dolgopyat, I. Goldsheid Quenched limit theorems for nearest neighbour random walks in 1D random environment, Comm. Math. Phys. 315 (2012) 241–277.
- [11] D. Dolgopyat, I. Goldsheid Local limit theorems for random walk in 1D random environment, Arxiv der Mathematik 101 (2013) 191–200.
- [12] D. Dolgopyat, I. Goldsheid Constructive approach to limit theorems for recurrent diffusive random walks on a strip, to appear in Asymptotic Analysis.
- [13] D. Dolgopyat, P. Nándori On mixing and the local central limit theorem for hyperbolic flows, Erg. Theory Dynam. Systems **20** (2020) 142–174.
- [14] D. Dolgopyat, P. Nándori Infinite measure renewal theorem and related results, Bull. Lond. Math. Soc. 51 (2019) 145–167.
- [15] D. Dolgopyat, P. Nándori Infinite measure mixing for some mechanical systems, preprint (2018).
- [16] W. Feller An introduction to probability theory and its applications, Vol. II. 2d ed. (1971) John Wiley & Sons, New York-London-Sydney. xxiv+669 pp.
- [17] B. V. Gnedenko, A. N. Kolmogorov Limit distributions for sums of independent random variables (1954) Cambridge, Addison-Wesley.
- [18] N. Guillotin-Plantard. Dynamic \mathbb{Z}^d -random walks in a random scenery: a strong law of large numbers, J. Theoret. Probab. **14** (2001) 241–260.
- [19] V. V. Jikov, S. M. Kozlov, O. A. Oleinik Homogenization of differential operators and integral functionals (1994) Springer, Berlin. xii+570 pp.
- [20] A. Katok, B. Hasselblatt Introduction to the modern theory of dynamical systems, Encyclopedia of Math. & Appl., **54** (1995) Cambridge University Press, Cambridge. xviii+802 pp.
- [21] H. Kesten, F. Spitzer A Limit Theorem Related to a New Class of Self Similar Processes, Z. Wahrscheinlichkeitstheorie verw. Gebiete 20 (1979) 5–25.
- [22] A. I. Khinchin Mathematical Foundations of Statistical Mechanics (1949) Dover, New York, N.Y. viii+179 pp.
- [23] D. Khoshnevisan, T. M. Lewis A law of the iterated logarithm for stable processes in random scenery, Stoch. Process. Appl. 74 (1998) 89–121.
- [24] M. Lenci On infinite-volume mixing, Comm. Math. Phys. 298 (2010) 485–514.
- [25] M. Lenci Uniformly expanding Markov maps of the real line: exactness and infinite mixing, Discrete Contin. Dyn. Syst. 37 (2017) 3867–3903.
- [26] P. Lévy Sur les séries dont les termes sont des variables éventuelles indépendantes, Studia Math. 3 (1931) 117–155.

- [27] P. Lévy Sur certains processus stochastiques homogènes, Compositio Math. 7 (1939) 283–339.
- [28] M. Lenci, S. Munday *Pointwise convergence of Birkhoff averages for global observables*, Chaos **28** (2018), 083111, 16 pp.
- [29] P. Lochak, C. Meunier *Multiphase averaging for classical systems*, Appl. Math. Sci. **72** (1988) Springer-Verlag New York. xii+360 pp.
- [30] J. Marcinkiewicz Quelques théorèmes de la théorie des probabilités, Travaux Société des Sciences et des Lettres de Wilno, Classe des Sciences Mathématiques et Naturelles, 13 (1939), 1–13.
- [31] G. C. Papanicolaou, S. R. S. Varadhan Boundary value problems with rapidly oscillating random coefficients, Colloq. Math. Soc. Janos Bolyai, 27 (1981) 835–873.
- [32] B. A. Rogozin An estimate of the remainder term in limit theorems of renewal theory, Theory Probab. Appl. 18 (1973) 662–677.
- [33] D. Ruelle Thermodynamic formalism. The mathematical structures of classical equilibrium statistical mechanics, Encyclopedia of Math. & Appl. 5 (1978) Addison-Wesley, Reading, Mass. xix+183 pp.
- [34] J. A. Sanders, F. Verhulst, J. Murdock Averaging methods in nonlinear dynamical systems, 2nd ed., Appl. Math. Sci. **59** (2007) Springer, New York. xxii+431 pp.
- [35] N. Veraverbeke Asymptotic behaviour of Wiener-Hopf factors of a random walk, Stoch. Process. Appl. 5 (1977) 27–37.

DEPARTMENT OF MATHEMATICS, UNIVERSITY OF MARYLAND, COLLEGE PARK, MD 20741, USA

Email address: dmitry@math.umd.edu

DIPARTIMENTO DI MATEMATICA, UNIVERSITÀ DI BOLOGNA, 40126 BOLOGNA, ITALY; ISTITUTO NAZIONALE DI FISICA NUCLEARE, SEZIONE DI BOLOGNA, 40126 BOLOGNA, ITALY

Email address: marco.lenci@unibo.it

Department of Mathematical Sciences, Yeshiva University, New York, NY 10016, USA

Email address: peter.nandori@yu.edu